

Module 2: Splines and Kernel Methods

B-Splines Recap

STAT/BIOSTAT 527, University of Washington

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April 15th, 2014

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Cubic Spline Basis and Fit

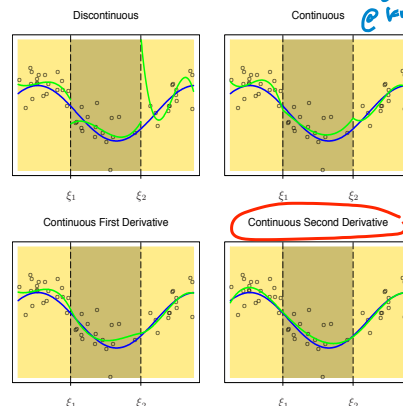
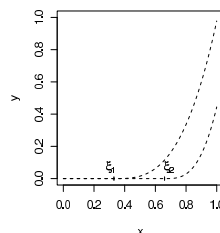
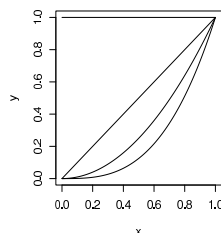
- Cubic spline function with K knots:

$$f(x) = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \sum_{k=1}^K b_k (x - \xi_k)_+^3$$

using truncated power basis

$M=4$
 $M-1$ deg poly
 $M-2$ cont. der.
@ knots

basis on $(0,1)$



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B-Splines

- Alternative basis for representing polynomial splines
- Computationally attractive... Non-zero over limited range
- As before:

- Knots $\xi_1 < \dots < \xi_K$ ✓
- Domain (a, b) ✓
- Number of basis functions = $M + K$ ✓

- Step 1: Add knots $\xi_0 = a$ $\xi_{K+1} = b$
- Step 2: Define auxiliary knots τ_j *needed to construct basis*

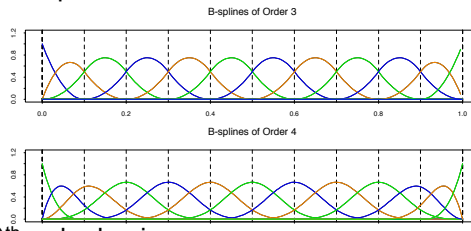
choice is arb. $\tau_1 \leq \tau_2 \leq \dots \leq \tau_M \leq \xi_0$
 $\tau_{j+M} = \xi_j$
 $\xi_{K+1} \leq \tau_{K+M+1} \leq \dots \leq \tau_{K+2M}$

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B-Splines

- For m^{th} order B-spline, $m=1, \dots, M$



From Hastie, Tibshirani, Friedman book

- Modify $(m-1)^{\text{th}}$ order basis:

$$B_j^m(x) = \frac{x - \tau_j}{\tau_{j+m-1} - \tau_j} B_j^{m-1} + \frac{\tau_{j+m} - x}{\tau_{j+m} - \tau_{j+1}} B_{j+1}^{m-1}$$

- B-spline bases are non-zero over domain spanned by at most $M+1$ knots
- Only subsets $\{B_i^m \mid i = M - m + 1, \dots, M + K\}$ are needed for basis of order m with knots ξ_j

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Cubic Splines as Linear Smoothers

- Cubic spline function with K knots:

$$f(x) = \beta_0 + \beta_1 x + \beta_2 x^2 + \beta_3 x^3 + \sum_{k=1}^K b_k (x - \xi_k)_+^3$$

truncated power basis

- Simply a linear model

$$f(x) = E[Y|c] = c^T \gamma$$

$$C = \begin{bmatrix} 1 & x_1 & x_1^2 & x_1^3 & (x_1 - \xi_1)_+^3 & \dots & (x_1 - \xi_K)_+^3 \\ \vdots & \vdots & \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & x_n & x_n^2 & x_n^3 & (x_n - \xi_1)_+^3 & \dots & (x_n - \xi_K)_+^3 \end{bmatrix}$$

n x (4K)

$$\gamma = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ b_1 \\ \vdots \\ b_K \end{bmatrix}$$

big matrix inversion

- Estimator:

$$\hat{\gamma} = (C^T C)^{-1} C^T Y$$

not (very) sparse

- Linear smoother:

$$\hat{f} = C(C^T C)^{-1} C^T Y$$

L

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Cubic B-Splines as Linear Smoother

- Cubic B-spline with K knots has basis expansion:

$$f(x) = \sum_{j=1}^{K+4} B_j^4(x) \beta_j$$

- Simply a linear model

$$B = \begin{bmatrix} B_1^4(x_1) & \dots & B_{K+4}^4(x_1) \\ \vdots & \ddots & \vdots \\ B_1^4(x_n) & \dots & B_{K+4}^4(x_n) \end{bmatrix}$$

$$\gamma = \begin{bmatrix} \beta_1 \\ \beta_2 \\ \vdots \\ \beta_{K+4} \end{bmatrix}$$

potentially very sparse bc of compact support of basis function (lower 4-banded)

$$\hat{\gamma} = (B^T B)^{-1} B^T Y$$

- Computational gain:

$n \times (K+4)$ matrix B with many 0's
 → fewer multiplies (sparse inv.)

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Return to Smoothing Splines

- Objective:

$$\min_f \sum_{i=1}^n (y_i - f(x_i))^2 + \lambda \int f''(x)^2 dx$$

- Solution:

- Natural cubic spline

- Place knots at every observation location x_i

roughness penalty

- Proof: See Green and Silverman (1994, Chapter 2) or Wakefield textbook

- Notes:

- Would seem to overfit, but penalty term shrinks spline coefficients toward linear fit
 - Will not typically interpolate data, and smoothness is determined by λ

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Smoothing Splines

- Model is of the form: $f(x) = \sum_{j=1}^n N_j(x) \beta_j$

of obs.

What we had before

- Rewrite objective:

$$(y - N\beta)^T (y - N\beta) + \lambda \beta^T \Omega_N \beta$$

natural cubic spline basis

- Solution:

$$\beta = (N^T N + \lambda \Omega_N)^{-1} N^T y$$

as in ridge

$$[N]_{ij} = N_j(x_i)$$

$$[\Omega_N]_{jk} = \int N_j''(t) N_k''(t) dt$$

- Linear smoother:

$$\hat{f} = \underbrace{N(N^T N + \lambda \Omega_N)^{-1} N^T}_{L_\lambda} y$$

"smoothing matrix"

$\nu_\lambda = \text{tr}(L_\lambda)$

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Smoothing Splines

- Previously,*

Model is of the form: $f(x) = \sum_{j=1}^n N_j(x) \beta_j$

Now,

Using B-spline basis instead: $f(x) = \sum_{j=1}^{n+4} B_j^4(x) \beta_j$

*K = n knots
order M = 4
Spline (cubic)*
- Solution: $\hat{\beta} = (B^T B + \lambda \Omega_B)^{-1} B^T y$

*n x (n+4)
lower 4 banded → computational eff.*
- Penalty implicitly leads to natural splines

 - Objective gives infinite weight to non-zero derivatives beyond boundary

*forces soln to be linear beyond boundary pts
→ natural splines*

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Spline Overview (so far)

Smoothing Splines

- Knots at data points x_i
- Natural cubic spline
- $O(n)$ parameters
 - Shrunk towards subspace of smoother functions

due to roughness penalty
- Linear smoothers, for example using natural cubic spline basis:

$$L = N(N^T N + \lambda \Omega_N)^{-1} N^T \quad \text{vs.} \quad L = N(N^T N)^{-1} N^T$$

n x n *penalty* *n x K* *# params = 4 + K - 4* *add'l const.*

Regression Splines

- $K \leq n$ knots chosen
- M^{th} order spline = piecewise $M-1$ degree polynomial with $M-2$ continuous derivatives at knots
- no regularization term, but many fewer params*

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Module 2: Splines and Kernel Methods

Penalized Regression Splines

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Penalized Regression Splines

- Alternative approach:
 - Use $K < n$ knots *few params relative to # of obs.*
 - How to choose K and knot locations?
- Option #1:
 - Place knots at n unique observation locations x_i and do stepwise
 - Issue?? *2^n models!*
- Option #2:
 - Place many knots for flexibility
 - Penalize parameters associated with knots *just like ridge/lasso*
- Note: Smoothing splines penalize complexity in terms of roughness. Penalized reg. splines shrink coefficients of knots.

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Penalized Regression Splines

- General spline model

$$f(x) = \sum_{j=1}^J h_j(x) \beta_j$$

some spline basis

- Definition: A **penalized regression spline** is $\hat{\beta}^T h(x)$ with

$$\hat{\beta} = \min_{\beta} \sum_{i=1}^n (y_i - \beta^T h(x_i))^2 + \lambda \beta^T D \beta$$

penalty matrix

- Form of resulting spline depends on choice of

- ☐ Basis $\{h_j(x)\}$
- ☐ Penalty matrix D
- ☐ Penalty strength λ

- Still need to choose K and associated locations. RoT (Ruppert et al 2003):

$$K = \min\left(\frac{1}{4} \times \# \text{ unique } x_i, 35\right) \quad \xi_k \text{ at } \frac{k+1}{K+2} \text{th points of } x_i$$

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PRS Example #1

$$\sum_{i=1}^n (y_i - \beta^T h(x_i))^2 + \lambda \beta^T D \beta$$

- Cubic B-spline basis + penalty

$$h_j = B_j^4 \quad \lambda \int \left(\sum_{j=1}^{K+4} B_j^4(x) \beta_j \right)^2 dx \quad \beta = \begin{bmatrix} \beta_1 \\ \vdots \\ \beta_{K+4} \end{bmatrix}$$

- For this penalty, the matrix D is given by

$$D_{jk} = \int B_j^4(x) B_k^4(x) dx$$

- Leads to "O'Sullivan splines"

when $K=n$, exactly equivalent to a smoothing spline
 knots @ unique x_i

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PRS Example #2

$$\sum_{i=1}^n (y_i - \beta^T h(x_i))^2 + \lambda \beta^T D \beta$$

- B-spline basis + penalty

$$\lambda \sum_{j=1}^{J-1} (\beta_{j+1} - \beta_j)^2$$

- For this penalty, the matrix D is given by

$$D = \begin{bmatrix} 1 & -1 & 0 & \cdots & 0 \\ -1 & 2 & -1 & 0 & \cdots \\ 0 & -1 & 2 & -1 & 0 \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

- Leads to

"p-splines"

penalizes large changes in coeff.
of adj. basis fns.

→ smoothing

≈ integrated squared derivative
penalty of O'Sullivan Splines

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PRS Example #3

$$\sum_{i=1}^n (y_i - \beta^T h(x_i))^2 + \lambda \beta^T D \beta$$

- Cubic spline using truncated power basis h_j

$$f(x) = \beta_0 + \beta_1 x + \cdots + \beta_3 x^3 + \sum_{k=1}^K b_k (x - f_k)_+^3$$

+ penalty on truncated power coefficients

$$\lambda \sum_k b_k^2 \Leftrightarrow \lambda \| \underline{b} \|_2^2$$

- For this penalty, the matrix D is given by

$$\beta = \begin{bmatrix} \beta_0 \\ \beta_1 \\ \beta_2 \\ \beta_3 \\ b_1 \\ \vdots \\ b_K \end{bmatrix} \quad D = \begin{bmatrix} 0 & \cdots & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & \cdots & 0 \\ \vdots & \ddots & \vdots & \ddots & \vdots \\ 0 & \cdots & 0 & \cdots & 0 \end{bmatrix}$$

β_j 's b_k 's

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A Brief Spline Summary

- **Smoothing spline** – contains n knots
- **Cubic smoothing spline** – piecewise cubic
- **Natural spline** – linear beyond boundary knots
- **Regression spline** – spline with $K < n$ knots chosen
- **Penalized regression spline** – imposes penalty (various choices) on coefficients associated with piecewise polynomial
- The # of basis functions depends on
 - # of knots
 - Degree of polynomial
 - A reduced number if a natural spline is considered (add constraints)

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Reading

- Hastie, Tibshirani, Friedman: 5.1-5.5 (skipping 5.3), Ch. 5 appendix
- Wakefield: 11.1.1-11.2.6

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What you should know...

- Regression splines
 - Cubic splines, natural cubic splines, ...
 - Interpretation as a linear smoother
 - Degrees of freedom
- Smoothing splines
 - Arising from penalized regression setting with smoothness penalty
 - Cubic spline basis with knots at every data point
- Natural splines
 - Linear beyond boundary points
- B-splines
 - Basis functions with compact support
- Penalized regression splines
 - Choose knots as in regression splines, but penalize associated coefficients

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Module 2: Splines and Kernel Methods

Local Polynomial Reg.,
Kernel Density Estimation

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Motivating Kernel Methods

- Recall original goal from Lecture 1:
 - We don't actually know the data-generating mechanism
 - Need an estimator $\hat{f}_n(\cdot)$ based on a random sample Y_1, \dots, Y_n , also known as **training data**

- Proposed a simple model as estimator of $E[Y|X]$

$$\hat{f}(x) = \text{Avg}(y_i \mid x_i \in \text{Nbhd}(x))$$

↑
use all obs. y_i in
a neighborhood of
target x

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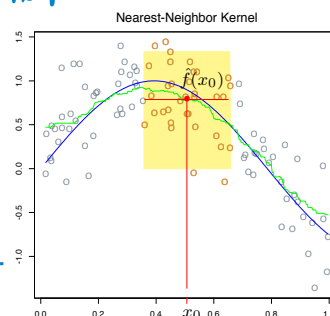
Choice #1: k Nearest Neighbors

- Define nbhd of each data point x_i by the k nearest neighbors
 - Search for k closest observations and average these

$$\hat{f}(x) = \text{Avg}(y_i \mid x_i \in N_k(x))$$

↑
k-nearest neighbors

- Discontinuity is unappealing
neighbors are either in or out
→ disc.



From Hastie, Tibshirani, Friedman book

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Choice #2: Local Averages

- A simpler choice examines a fixed distance h around each x_i

- Define set: $B_x = \{i : |x_i - x| \leq h\}$

- # of x_i in set: n_x

$$\hat{f}(x) = \frac{1}{n_x} \sum_{i \in B_x} y_i \quad \text{avg. obs. within distance } h$$

- Results in a linear smoother

$$\hat{f}(x) = \sum_{i=1}^n l_i(x) y_i \quad l_i(x) = \begin{cases} \frac{1}{n_x} & \text{if } |x_i - x| \leq h \\ 0 & \text{otherwise} \end{cases}$$

- For example, with $x_i = \frac{i}{9}$ and $h = \frac{1}{9}$

$$L = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} & 0 & 0 & \dots \\ \frac{1}{3} & \frac{1}{3} & \frac{1}{3} & 0 & \dots \\ 0 & \frac{1}{3} & \frac{1}{3} & \frac{1}{3} & 0 \dots \end{bmatrix}$$

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More General Forms

- Instead of weighting all points equally, slowly add some in and let others gradually die off

- **Nadaraya-Watson kernel weighted average**

$$\hat{f}(x_0) = \frac{\sum_{i=1}^n K_\lambda(x_0, x_i) y_i}{\sum_{i=1}^n K_\lambda(x_0, x_i)}$$

$$K_\lambda(x_0, x) = K\left(\frac{|x_0 - x|}{\lambda}\right)$$

↑ kernel
 ↑
 bandwidth

- But what is a **kernel** ???

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Kernels

- Could spend an entire quarter (or more!) just on kernels
- Will see them again in the Bayesian nonparametrics portion
- For now, the following definition suffices

$K(\cdot)$ is a kernel if

$$K(x) \geq 0 \quad \forall x$$

$$\int K(u) du = 1$$

$$\int u K(u) du = 0$$

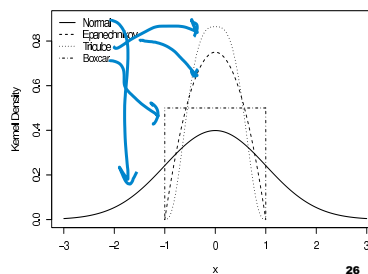
$$\sigma_K^2 = \int u^2 K(u) du < \infty$$

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Example Kernels

- *Gaussian* $K(x) = \frac{1}{2\pi} e^{-\frac{x^2}{2}}$
- *Epanechnikov* $K(x) = \frac{3}{4} (1 - x^2) I(x)$
- *Tricube* $K(x) = \frac{70}{81} (1 - |x|^3)^3 I(x)$
- *Boxcar* $K(x) = \frac{1}{2} I(x)$



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Nadaraya-Watson Estimator

- Return to Nadaraya-Watson kernel weighted average

$$\hat{f}(x_0) = \frac{\sum_{i=1}^n K_\lambda(x_0, x_i) y_i}{\sum_{i=1}^n K_\lambda(x_0, x_i)}$$

- Linear smoother:

$$\hat{f}(x_0) = \sum_{i=1}^n \underbrace{\frac{K_\lambda(x_0, x_i)}{\sum_{i=1}^n K_\lambda(x_0, x_i)}}_{l_i(x_0)} y_i = \sum_{i=1}^n l_i(x_0) y_i$$

$$\hat{f} = L_\lambda Y$$

$$V_\lambda = \text{tr}(L_\lambda)$$

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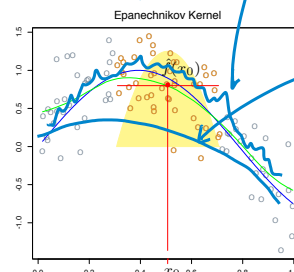
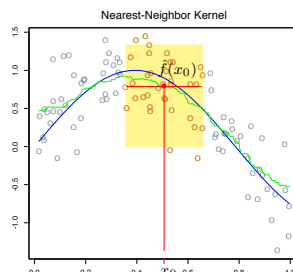
Nadaraya-Watson Estimator

- Example:

- ☐ Boxcar kernel → local avg
- ☐ Epanechnikov
- ☐ Gaussian typical

$$\hat{f}(x_0) = \frac{\sum_{i=1}^n K_\lambda(x_0, x_i) y_i}{\sum_{i=1}^n K_\lambda(x_0, x_i)}$$

- Often, choice of kernel matters much less than choice of λ



From Hastie, Tibshirani, Friedman book

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Local Linear Regression

- Locally weighted averages can be badly biased at the boundaries because of asymmetries in the kernel

- Reinterpretation:

$$\hat{f} = \arg \min_a \sum (y_i - a)^2$$

$$\rightarrow \hat{f} = \bar{y}$$

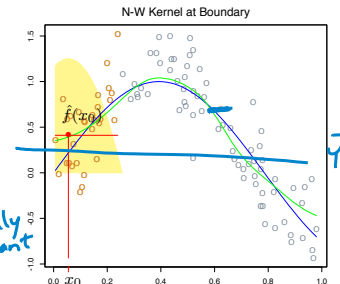
$$\hat{f}(x_0) = \arg \min_a \sum w_i(x_0) (y_i - a)^2$$

local to x_0

$$\hat{f}(x_0) = \frac{\sum w_i(x_0) y_i}{\sum w_i(x_0)}$$

restrict to locally constant

$k\left(\frac{|x_0 - x_i|}{h}\right)$



From Hastie, Tibshirani, Friedman book

- Equivalent to the Nadaraya-Watson estimator
- Locally constant estimator obtained from weighted least squares

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Local Linear Regression

- Consider locally weighted linear regression instead
- Local linear model around fixed target x_0 :

$$\beta_{0x_0} + \beta_{1x_0}(x - x_0)$$

- Minimize:

$$\min_{\beta_{x_0}} \sum_i K_h(x_0, x_i) (y_i - (\beta_{0x_0} + \beta_{1x_0}(x_i - x_0)))^2$$

- Return:

$$\hat{f}(x_0) = \hat{\beta}_{0x_0} \leftarrow \text{fit at } x_0$$

Note: not equivalent to fitting a local constant

- Fit a new local polynomial for every target x_0

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Local Linear Regression

$$\min_{\beta_{x_0}} \sum_{i=1}^n K_{\lambda}(x_0, x_i) (y_i - \beta_{0x_0} - \beta_{1x_0}(x_i - x_0))^2$$

- Equivalently, minimize

$$(y - X_{x_0} \beta_{x_0})^T W_{x_0} (y - X_{x_0} \beta_{x_0})$$

$\begin{bmatrix} K_{\lambda}(x_0, x_1) \\ \vdots \\ K_{\lambda}(x_0, x_n) \end{bmatrix}$

- Solution:

$$\hat{\beta}_{x_0} = (X_{x_0}^T W_{x_0} X_{x_0})^{-1} X_{x_0}^T W_{x_0} y$$

$\begin{bmatrix} x_1 - x_0 \\ x_2 - x_0 \\ \vdots \\ x_n - x_0 \end{bmatrix}$

$$\hat{f}(x_0) = e_1^T \hat{\beta}_{x_0} \quad (1, 0, \dots, 0) \text{ grabs out 1st element}$$

$$= \sum l_i(x) y_i$$

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Local Linear Regression

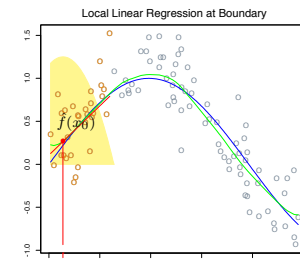
$$y_i = f(x_i) + \epsilon_i$$

- Bias calculation:

$$\begin{aligned} E[\hat{f}(x_0)] &= \sum_i l_i(x_0) f(x_i) \quad E[y_i] = 0 \text{ (can show)} \\ &= f(x_0) \sum l_i(x_0) + f'(x_0) \sum (x_i - x_0) l_i(x_0) + \\ &\quad \frac{f''(x_0)}{2} \sum (x_i - x_0)^2 l_i(x_0) + R \leftarrow \text{higher order terms} \\ &= f(x_0) + f''(x_0) \dots \end{aligned}$$

- Bias $E[\hat{f}(x_0)] - f(x_0)$ only depends on quadratic and higher order terms

- Local linear regression corrects bias exactly to 1st order



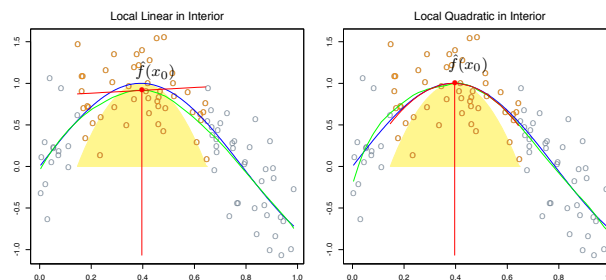
From Hastie, Tibshirani, Friedman book

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Local Polynomial Regression

- Local linear regression is biased in regions of curvature
 - “Trimming the hills” and “filling the valleys”
- Local quadratics tend to eliminate this bias, but at the cost of increased variance



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Local Polynomial Regression

- Consider local polynomial of degree d centered about x_0

$$P_{x_0}(x; \beta_{x_0}) = \beta_0 x_0 + \beta_1 x_0 (x - x_0) + \frac{\beta_2 x_0}{2!} (x - x_0)^2 + \dots + \frac{\beta_d x_0}{d!} (x - x_0)^d$$

- Minimize: $\min_{\beta_{x_0}} \sum_{i=1}^n K_\lambda(x_0, x_i) (y_i - P_{x_0}(x; \beta_{x_0}))^2$

- Equivalently:

$$\min (\mathbf{y} - \mathbf{X}_{x_0} \underline{\beta}_{x_0})^T \mathbf{W}_{x_0} (\mathbf{y} - \mathbf{X}_{x_0} \underline{\beta}_{x_0})$$

$$\mathbf{X}_{x_0} = \begin{bmatrix} 1 & x_1 - x_0 & \dots & \frac{(x_1 - x_0)^d}{d!} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & x_n - x_0 & \dots & \frac{(x_n - x_0)^d}{d!} \end{bmatrix}$$

- Return: $\hat{f}(x_0) = \hat{\beta}_0 x_0$
- Bias only has components of degree $d+1$ and higher

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Local Polynomial Regression

- Rules of thumb:

- ☐ Local linear fit helps at boundaries with minimum increase in variance
- ☐ Local quadratic fit doesn't help at boundaries and increases variance
- ☐ Local quadratic fit helps most for capturing curvature in the interior
- ☐ Asymptotic analysis →
local polynomials of odd degree dominate those of even degree
(MSE dominated by boundary effects)
- ☐ Recommended default choice: local linear regression

Reading

- Hastie, Tibshirani, Friedman: 6.1-6.2, 6.6
- Wakefield: 11.3

What you should know...

- Definition of a kernel and examples
- Nearest neighbors vs. local averages
- Nadarya-Watson estimation
 - Interpretation as local linear regression
- Local polynomial regression
 - Definition
 - Properties/ rules of thumb