## Statistics 581, Practice Final Exam Wellner; 11/28/2018

## This exam is to be taken without the use of any books or notes.

- 1. (48) points) **Define** each of the following terms. In each case, provide an appropriate context for your definition.
  - (a) The information matrix for  $\theta$  in a regular parametric model  $\mathcal{P} = \{P_{\theta}: \ \theta \in \Theta \subset R^k\}.$
  - (b) The efficient score function for a parameter  $\theta_1$  when  $\theta = (\theta_1, \theta_2)$ .
  - (c) The efficient influence function  $l_1$  for a parameter  $\theta_1$  when  $\theta = (\theta_1, \theta_2)$ .
  - (d) The efficient influence function for  $\tilde{l}_{\nu}$  for a differentiable parameter  $q(\theta) = \nu(P_{\theta})$  in a regular parametric model  $\mathcal{P}$ .
  - (e) An asymptotically linear estimator  $T_n$  of a parameter  $\nu(P)$  with influence function  $\psi$ .
  - (f) A locally regular estimator  $T_n$  of a parameter  $\nu(P_{\theta})$ .
- 2. (32 points) **State** four of the following six results, providing the appropriate (brief) context for your statement:
  - (a) The (elementary) Skorokhod theorem.
  - (b) The multiparameter Cramér Rao inequality (for an unbiased estimator)  $T = T(\underline{X})$  of a real-valued parameter  $q(\theta) = \nu(P_{\theta})$ .
  - (c) A result about the finite-dimensional limiting distributions of the sample quantile process  $\{\sqrt{n}(\mathbb{F}_n^{-1}(t) F^{-1}(t)): 0 < t < 1\}$  specifying the assumption(s) carefully.
  - (d) The Glivenko-Cantelli theorem.
  - (e) LAN (Local Asymptotic Normality) of the local log likelihood ratios for a regular parametric model satisfying the Cramér hypotheses).
  - (f) The asymptotic behavior of the likelihood ratio statistic  $2 \log \lambda_n$  for testing a simple null hypothesis  $\theta = \theta_0$  versus  $\theta \neq \theta_0$  under a fixed alternative  $P_{\theta}$ with  $\theta \neq \theta_0$ .
- 3. (40 points) Suppose that  $\mathcal{P} = \{P_{\theta} : \theta \in (0, 1)\}$  where  $P_{\theta}$  has density with respect to Lebesgue measure  $\lambda$  on [0, 1] given by

$$p_{\theta}(x) = 2\left\{\frac{x}{\theta}\mathbf{1}_{[0,\theta]}(x) + \frac{1-x}{1-\theta}\mathbf{1}_{(\theta,1]}(x)\right\}.$$

We showed in problem 5 of problem set # 9 that

$$\dot{\mathbf{l}}_{\theta}(x) = -\frac{1}{\theta} \mathbf{1}_{[0,\theta]}(x) + \frac{1}{1-\theta} \mathbf{1}_{(\theta,1]}(x),$$
$$I(\theta) = \frac{1}{\theta} + \frac{1}{1-\theta} = \frac{1}{\theta(1-\theta)},$$

and that if  $X_1, \ldots, X_n$  are i.i.d.  $P_{\theta}$  with  $\theta \in (0, 1)$ , then a  $\sqrt{n}$ -consistent preliminary estimator of  $\theta$  is given by  $\overline{\theta}_n = 3\overline{X}_n - 1$  which satisfies

$$\sqrt{n}(\overline{\theta}_n - \theta) \rightarrow_d N(0, (1 - \theta + \theta^2)/2).$$

- (a) Use the above facts to suggest a one-step estimator  $\dot{\theta}_n$  of  $\theta$ .
- (b) Show that the estimator  $\check{\theta}_n$  you proposed in (a) can be written as

$$\check{\theta}_n = \overline{\theta}_n - (\mathbb{F}_n(\overline{\theta}_n) - \overline{\theta}_n)$$

where  $\mathbb{F}_n(x) = n^{-1} \sum_{i=1}^n \mathbb{1}_{[X_i \leq x]}$  is the empirical distribution of the  $X_i$ 's.

(c) Show by a direct argument using the result of (b) that  $\check{\theta}_n \to_p \theta$ . Hints: Calculate the distribution function  $F_{\theta}$  corresponding to the density  $p_{\theta}$ , show that  $F_{\theta}(\theta) = \theta$ , and use the Glivenko-Cantelli theorem.

(d) Show by a direct argument using the result of (b) that  $\sqrt{n}(\check{\theta}_n - \theta) \rightarrow_d N(0, I(\theta)^{-1})) = N(0, \theta(1 - \theta))$ . Hint: Use Donsker's theorem and the deltamethod applied to  $g(x) = F_{\theta}(x) - x$ .

(e) Compare the asymptotic variance of the preliminary estimator  $\overline{\theta}_n$  to the asymptotic variance of the one- step estimator  $\check{\theta}_n$ .

4. (40 points) Suppose that  $X_1, \ldots, X_n$  are i.i.d.  $N(\mu, \sigma^2)$ , and let  $\theta = (\mu, \sigma^2) \in \Theta \equiv \mathbb{R} \times (0, \infty)$ . Consider testing  $H : \sigma^2 = \sigma_0^2$  where  $\sigma_0^2$  is a known constant, versus  $K : \sigma^2 \neq \sigma_0^2$ .

(a) What are the score functions for  $\mu$  and  $\sigma^2$ , and information matrix for  $\theta$  in this model?

(b) What is the maximum likelihood estimator  $\widehat{\theta}_n$  of  $\theta = (\mu, \sigma^2)$  under  $\theta \in \Theta$ ? What is the maximum likelihood estimator  $\widehat{\theta}_n^0$  of  $\theta = (\mu, \sigma^2)$  under  $\theta \in \Theta_0$ ?

(c) Identify the null hypothesis H in terms of a subset  $\Theta_0$  of  $\Theta$ .

(d) What are the likelihood ratio, Wald, and score (or Rao) statistics for testing H versus K? What is the limiting distribution of these three test statistics under the null hypothesis H?

(e) Suppose that  $X_1, \ldots, X_n$  are i.i.d.  $P_{\theta}$  with  $\theta \notin \Theta_0$ . Find the limits of  $n^{-1}2 \log \lambda_n, n^{-1}W_n, n^{-1}R_n$  (almost surely or in probability).

- 5. (48 points). (Poisson regression). Suppose that  $(Y|Z) \sim \text{Poisson}(\lambda e^{\gamma Z})$ , and  $Z \sim \text{Bernoulli}(\eta)$  on  $\{0, 1\}$  You may assume that  $\eta$  is known and that  $0 < \eta < 1$ . Thus Z is a "covariate" or "predictor variable",  $\gamma$  is a "regression parameter" which affects the intensity of the (conditionally) Poisson variable Y, and  $\theta = (\lambda, \gamma)$ .
  - (a) Find the information matrix for  $\theta$ .

(b) Find the information and information bound for estimation  $\gamma$  if the parameter  $\lambda$  is unknown.

(c) Find the efficient score function and the efficient influence function for estimation of  $\gamma$  when  $\lambda$  is unknown. Interpret these geometrically in terms of the scores for  $\gamma$  and  $\lambda$ .

(d) If we observe  $X_i = (Y_i, Z_i)$ , i = 1, ..., n, i.i.d.  $P_{\theta}$ , write down the likelihood equations for the maximum likelihood estimator  $\hat{\theta}_n = (\hat{\lambda}_n, \hat{\gamma}_n)$ . What do our theorems tell us about the asymptotic normality of  $\hat{\theta}_n$ ?

6. (48 points). (Poisson regression, continued).

(a) Suggest three tests of the (composite!) null hypothesis  $H : \gamma = 0$  versus  $K : \gamma \neq 0$ . What is the distribution of each of these three statistics under the null hypothesis and under local alternatives of the form  $\gamma_n = tn^{-1/2}$ ?

(b) Consider estimation of the function

$$q(\theta) = \nu(P_{\theta}) = E_{\theta}(Y|Z=1).$$

Compute  $q(\theta)$  explicitly as a function of  $\theta = (\lambda, \gamma)$ .

(c) Suggest a natural empirical estimator of this conditional expectation which does not rely on the Poisson model. If this estimator is called  $\tilde{\nu}_n$ , show that  $\tilde{\nu}_n$  is asymptotically linear and find its influence function  $\psi$  explicitly.

(d) Find the efficient influence function  $\mathbf{l}_{\nu}$  for estimation of  $\nu(P_{\theta})$  assuming the Poisson model.

(e) Describe the relationship between  $\psi$  and  $\mathbf{l}_{\nu}$  geometrically.