

Lecture 23 (ch. 8-end)

Back to proportions

When we compare 2 props (π_1, π_2) (e.g. $H_1: \pi_1 - \pi_2 < 0.1$) we have Two populations, each with 2 groups/catg. (e.g. Boy/Girl). In that case, ONE proportion (e.g. prop of boys, π_{Boy}) is enough to describe each pop., because the other prop. (e.g. π_{Girls}) is fixed by $1 - \pi_{\text{Boys}}$. In other words, our 2-sample CIs/tests involve Two proportions, one from each of Two populations.

So, an example would be $\begin{cases} \pi_1 = \pi_{\text{Boys}} & \text{in Northern hemisphere} \\ \pi_2 = \pi_{\text{Boys}} & \text{in Southern hemisphere} \end{cases}$

Note that both π_1 and π_2 refer to Boys, but in 2 different populations (e.g. Northern and Southern hemispheres).

Then, we can test, e.g. $H_0: \pi_1 - \pi_2 = 0$ vs. $H_1: \pi_1 - \pi_2 < 0.1$

But there are situations where each population has more than 2 categories, and we want to test some claim about the proportions of each category.

In fact, we can even have ONE population with many catgs.

If we have ONE pop, with k categories, we can test $H_0: \pi_1 = \pi_{01}, \pi_2 = \pi_{02}, \dots, \pi_k = \pi_{0k}$, prop. of k^{th} catg. in pop.

H_1 : At least one of π is wrong
I'll explain this later.

$$\sum_{i=1}^k \pi_{0i} = 1$$

Of course, given that there is only ONE pop., we must have

Below, we will see how to do this test.

There will be a new distribution: Chi-squared.

Also note that a pop. with 2 groups can be thought of as being described by one random variable with 2 levels. Similarly, a pop. with k groups can be described with one r.v. with k levels.

Monthly Weather Review, 2008:
Vol. 136, p. 3121. Cook & Schaefer.

E.g.

Does data provide sufficient evidence to support an association between climate and tornadic activity?

of Days with violent tornadoes in each climate category

Climate Category	El Niño	La Niña	Normal
# of Days with violent tornadoes	$n_1 = 14$	$n_2 = 28$	$n_3 = 44$

(86)

generalization of Binomial to more than 2 categories, FYI: Multinomial dist.

proportion: $\frac{14}{86} = 0.16$ 0.33 0.51 (1)

Data.

of years classified as

Climate Category	El Niño	La Niña	Normal
# of years classified as	12	17	25

(54)

proportion: $\frac{12}{54} = 0.22$ 0.32 0.46 (1)

H_0 : true prop. of tornadic days in El Niño years. Etc.
There is no association, i.e.

H_0 : $\pi_1 = 0.22$ $\pi_2 = 0.32$ $\pi_3 = 0.46$

H_1 : At least one of these assignments is wrong.

Q: If $H_0 = \text{True}$, how many tornadoes do you expect in each of the $k=3$ categories?

A: Expected counts:

Climate Category	El Niño	La Niña	Normal
Expected counts	$0.22(86) \approx 18.9$	$0.32(86) \approx 27.5$	$0.46(86) \approx 39.6$

(86)

observed counts:

Climate Category	El Niño	La Niña	Normal
observed counts	14	28	44

$(\text{Exp.} - \text{obs})^2$:

Climate Category	El Niño	La Niña	Normal
$(\text{Exp.} - \text{obs})^2$	$(4.9)^2$	$(-0.5)^2$	$(-4.4)^2$

$\frac{(\text{Exp.} - \text{obs})^2}{\text{Exp.}}$:

Climate Category	El Niño	La Niña	Normal
$\frac{(\text{Exp.} - \text{obs})^2}{\text{Exp.}}$	1.27	0.009	0.49

Like $z_{\text{obs}}, t_{\text{obs}}$, $\chi^2_{\text{obs.}} = \sum_{i=1}^3 \frac{(\text{exp.} - \text{obs})^2}{\text{exp.}} = 1.77$

If there were really no difference at all in the # of tornadoes between the 3 categories, then this would be near zero.

Q So, is this χ^2_{obs} far away from 0 to reject H_0 (in favor of H_1)?

Note: χ^2 is non-negative, unlike z, t

A We need to know the sample distr. of χ^2 , when $H_0 = T$.

Theorem: Under the null hypothesis, χ^2 has a chi-squared distr. with $df = k - 1$ ($= 3 - 1 = 2$)

What's a chi-squared dist? It's just another Table (VII).
But FYI.

$$p\text{-value} = \text{prob}(\chi^2 > \chi^2_{obs}) = \text{prob}(\chi^2 > 1.77) > 0.1$$

see two pages down
 $\uparrow df = 3 - 1 = 2$

$$R: \text{pchisq}(1.77, df = 3 - 1, \text{lower.tail} = F) = 0.41$$

For the chi-squared test, p-value is always right area.

Conclusion (at $\alpha = .01$): $p\text{-value} > \alpha$

In words: Cannot reject H_0 in favor of H_1 . at least 1 is wrong.

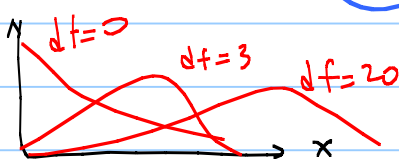
$$(\pi_1 = .22, \pi_2 = .32, \pi_3 = .46)$$

In English: There is no evidence from data to suggest that the 3 props are NOT .22, .32, .46, i.e.

I.e. There is no evidence from data that there is an association between tornadic activity and climate.

The chi-squared density function is (FYI)

$$f(x) = \frac{x^{\frac{df}{2}-1} e^{-x/2}}{\Gamma(\frac{df}{2}) 2^{df/2}}$$



*

Here is a generalization of the above to any k (above $k=3$).
 I propose that you do NOT use the formula at the bottom,
 but instead do things like I did above. After, you are
 completely comfortable with the steps, then you can use this page.
But Do read The IMPORTANT Note at the bottom.

Let π_i = proportion of cases in category i (where $i=1, 2, \dots, k$).

	Null params	Tornado Example
π_1 = proportion of categ. 1's	π_{01}	0.22
π_2 = - - - - 2's	π_{02}	0.32
π_3 = - - - - 3's	π_{03}	0.46

If $H_0 = \text{True}$, $H_0: \pi_1 = \pi_{01}, \pi_2 = \pi_{02}, \dots$

Then in a sample of size n , how many would

we <u>expect</u> in category 1 :	$n\pi_{01}$	18.9
" " " " 2 :	$n\pi_{02}$	27.5
" " " " 3 :	$n\pi_{03}$	39.6
⋮		

But according to data,
 we observe this many :

$\sum_{i=1}^k n_i = n$	n_1	14
	n_2	28
	n_3	44

Punch line:

Then the theorem tells us that

$$\chi^2_{obs} = \sum_i \frac{(\text{exp.} - \text{obs})^2}{\text{exp.}} = \sum_{i=1}^k \frac{(n\pi_{0i} - n_i)^2}{n\pi_{0i}}$$

has a chi-sqd. distr with $df = k-1$

$$p\text{-value} = \text{prob}(\chi^2 > \chi^2_{obs})$$

IMPORTANT

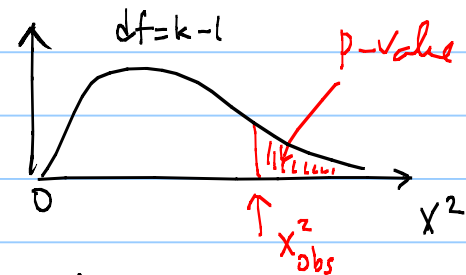
First, note that we are dealing with props, not means,
 (i.e. our r.v. is discrete)
 But, even though we wrote a lot of things in terms of props, the data are all in counts. This chi-sqd test deals with count data

How to use Table VII:

Table VII gives the area to the right of some value of χ^2_{obs} , i.e. it gives a p-value. However, it does not give all p-values; the only ones it provides are listed in the left-most column. E.g.

$$\chi^2_{obs} = 8.49, df=4 \Rightarrow p\text{-value} = 0.075$$

$$\chi^2_{obs} = 8.66, df=4 \Rightarrow p\text{-value} = 0.070$$



One might think that putting bounds on p-value is not enough for hypothesis testing, but it often is.

For example, suppose we get $\chi^2_{obs} = 8.55$ with $df=4$.

Then we can say $0.070 < p\text{-value} < 0.075$. That is good enough if $\alpha = .05$, because $p\text{-value} > \alpha$, and so we cannot reject H_0 in favor of H_1 .

Additional comments:

Note that the first H_0, H_1 above are just a generalization of

$$H_0: \pi = \pi_0 \quad (z\text{-test}).$$

$$H_1: \pi \neq \pi_0$$

to more than 2 categories in the population.

However, there are no 1-sided / 2-sided varieties of chi-sqd. ↩

When X_{obs}^2 is small (say ≈ 0), then the observed counts are consistent with the expected counts if $H_0 = T$

(i.e. $\pi_1 = \pi_{01}, \pi_2 = \pi_{02}, \dots, \pi_k = \pi_{0k}$). So, if X_{obs}^2 is large,

then at least one of these[↑] must be wrong.

In other words the appropriate hypotheses are

$$H_0: \pi_1 = \pi_{01}, \pi_2 = \pi_{02}, \dots, \pi_k = \pi_{0k}$$

H_1 : At least one of these[↑] specifications is wrong.

And it is the "At least" which gives us

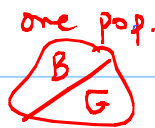
$$p\text{-value} = \text{prob}(X^2 > X_{obs}^2) \quad (\text{Table VII})$$

I.e. We are always interested in the upper tail area only

Said differently for the chi-sqd test of the above H_0/H_1 , the p-value is only the right area, because violation of each part of H_0 , increases X^2 .

Summary

⇒ 1 pop., 1 prop.

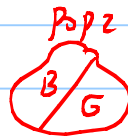


$$\pi_B = ? \quad (\pi_B + \pi_G = 1)$$

Data: $P_{obs} = 13/15$

$$Z_{obs} = (P_{obs} - \pi_0) / \sqrt{\frac{\pi_0(1-\pi_0)}{n}} \Rightarrow \text{p-value.}$$

⇒ 2 pop.'s, 2 props.



$$\pi_{1B} - \pi_{2B} = ? \quad (\pi_{1B} + \pi_{2B} \neq 1)$$

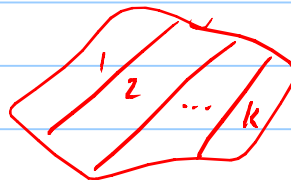
Data: $P_1 = 13/15$

$P_2 = 20/40$

$$Z_{obs} = \frac{(P_1 - P_2) - (\pi_{1B} - \pi_{2B})}{\sqrt{\dots}} \Rightarrow \text{p-value}$$

Note That when we write $H_0: \pi_1 - \pi_2 = \dots$, The π_i are The proportion of 1 of the 2 categories in either pop. e.g. π_1 = prop of something (e.g. boys) in population 1, and π_2 = " " same thing " " 2.

⇒ 1 pop., k categories
(The tornado example)



$$\pi_1 = ? , \pi_2 = ? , \dots , \pi_k = ?$$

$$(\pi_1 + \pi_2 + \dots + \pi_k = 1 \cdot)$$

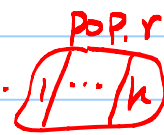
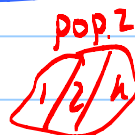
Data: $14 / 28 / 44$
observed counts

$$\chi^2_{obs} = \sum \left(\frac{\text{obs} - \text{exp}}{\sqrt{\text{exp}}} \right)^2 \sim \text{chisqd}$$

$$df = k - 1$$

Note That when we talk about $\pi_1, \pi_2, \dots, \pi_k$, The indices refer to k levels of some categorical v.v.

⇒ r pops, k categories



$$i = 1, \dots, k$$

$$j = 1, \dots, r$$

Data: Contingency Table
(Confusion Matrix)

Skipped

pop.

	Category			
	1	2	...	k
1				
2				
...				
r				

counts

Skipped

$$\chi^2 = \sum_{\text{all cells}} \left(\frac{\text{obs} - \text{exp}}{\sqrt{\text{exp}}} \right)^2 \sim \text{chisqd}$$

$$df = (r-1)(k-1) \Rightarrow \text{p-value.}$$

hw_lect23_1 (By hand)

Consider the data from an example in a past lecture where a survey of students in 390 yielded the following data:

17 students like lab

48 do not like lab

15 have no opinion

Suppose I believed that the proportion of students in each of the 3 categories (like, no-like, no-opinion) was equal. Does this data contradict that belief? Use $\alpha = 0.05$

hw_lect23_2 (By hand)

A sample of 210 Bell computers has 56 defectives. Theory suggests that a third of all Bell computers should be defective. Does this data contradict the theory (at $\alpha=0.05$)? Specifically,

- Do a z-test ,
- Do a chi-squared test with $k=2$ categories. Hint: The π 's (and π_0 's) of the k categories must sum to 1.
- Are the conclusions in a and b consistent?