

# Lecture 8

Continuous distributions  
uniform (continued)

# Lecture Notes IV – Continuous distributions. Parametric density estimation.

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CDF and PDF. Sampling



Examples of continuous distributions



ML estimation for continuous distributions



ML estimation by gradient ascent



Reading: Ch.5, 6

# CDF and PDF refresher

## Cumulative distribution function (CDF)

$$F(x) = P[X \leq x] \quad (1)$$

1.  $F \geq 0$  positivity.
2.  $\lim_{x \rightarrow -\infty} F = 0$
3.  $\lim_{x \rightarrow \infty} F = 1$
4.  $F$  is an increasing function

## Probability density [function] (PDF)

$$f = \frac{dF}{dx} \quad (2)$$

$$P(a, b) = P[a, b] = F(b) - F(a) = \int_a^b f(x) dx \quad (3)$$

## normalization condition

$$\int_{-\infty}^{\infty} f(x) dx = 1 \quad (4)$$

## Examples of continuous distributions

# Families

$$\underline{\mathcal{F}_1} = \{u_{[a,b]}, a < b\} \quad \text{uniform}$$

(5)

$$f(x; a, b) = \begin{cases} \frac{1}{b-a}, & x \in [a, b] \\ 0, & \text{otherwise} \end{cases}$$

(6)

$$\underline{\mathcal{F}_2} = \{N(\cdot; \mu, \sigma^2)\} \quad \text{normal}$$

(7)

$$f(x; \mu, \sigma^2) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

(8)

$$\underline{F(x; a, b)} = \frac{1}{1 + e^{-ax-b}}, \quad a > 0 \quad \text{logistic}$$

(9)

$$f(x; a, b) = \frac{ae^{-ax-b}}{(1 + e^{-ax-b})^2}$$

(10)

ML estimation for continuous distributions = density estimation

$$S = (-\infty, \infty) \text{ or } S \subset (-\infty, \infty)$$

Sample space

Data  $\mathcal{D} = \{x^1, x^2, \dots, x^n\} \subset S$  iid from true unknown f

Model family  $\mathcal{F} = \{f(x|\theta), \theta = \text{parameters}\}$

Chosen  $\uparrow$  Normal, Uniform, ...

Problem estimate  $\theta$  from  $\mathcal{D}$

Solution : Max likelihood  $\rightarrow \theta^{ML} \rightarrow f(x|\theta^{ML})$   
estimated  $\uparrow$   
density

## ML estimation for continuous distributions

Likelihood of  $\theta$

$$L(\theta) = \prod_{i=1}^n f(x^i | \theta)$$

$P(x^i | \theta)$  for S discrete

$F(x)$

log-likelihood

$$\ell(\theta) = \sum_{i=1}^n \ln f(x^i | \theta)$$

Max Likelihood Principle

$$\text{choose } \theta^{ML} = \operatorname{argmax}_{\theta} \ell(\theta)$$

Data  $\mathcal{D} = \{x^{1:n}\} \subset \underbrace{\mathbb{R}}_S$

Model family

$$\mathcal{F} = \{U[a, b] : a < b, a, b \in \mathbb{R}\}$$

Likelihood

$$f(x) = \begin{cases} \frac{1}{b-a} & x \in [a, b] \\ 0 & \text{otherwise} \end{cases}$$

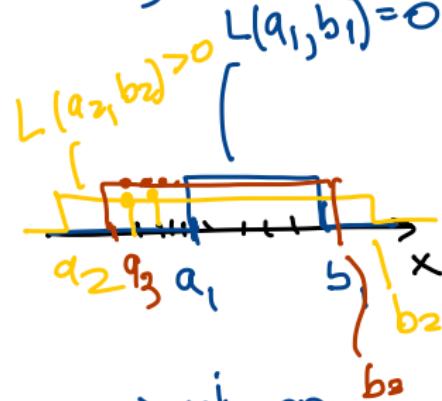
statistics

↑ parameters:  $a, b$

↑ calculate  $L(a, b | \mathcal{D})$

$$L(a, b) = \prod_{i=1}^n f(x_i) = \begin{cases} 0 & \text{if } a > \min_{i=1:n} x_i \text{ or} \\ & b < \max_{i=1:n} x_i \\ \frac{1}{(b-a)^n} & \Rightarrow \min|b-a| \end{cases}$$

Maximize

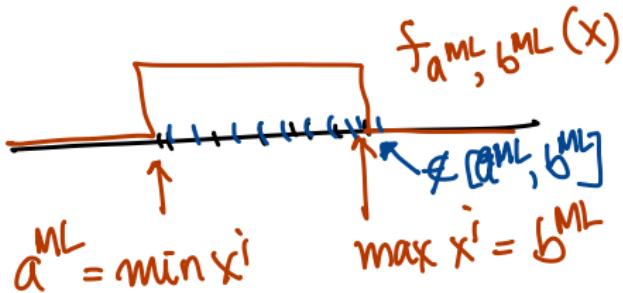


## ML estimation for continuous distributions

⇒

$$a^{ML} = \min_{i=1:n} x^i$$

$$b^{ML} = \max_{i=1:n} x^i$$



Pls even if  $x \sim \text{unif}[a^*, b^*]$  true distribution

$$\begin{matrix} a^{ML} > a^* \\ b^{ML} < b^* \end{matrix}$$

$$\Rightarrow P[x \in [a^*, a]] > 0$$

$$P[x \in (b, b^*)] > 0$$

but  $f_{a^{ML}, b^{ML}}(x) = 0 !!$  BAD!

# ML estimation for continuous distributions

$\hat{a}_{[a,b]}$

Solution choose  $\hat{a} < \hat{a}^{\text{ML}} = \min x_i$   
 $\hat{b} > \hat{b}^{\text{ML}} = \max x_i$

## Statistical solution 1

①  $1 - \delta = \text{confidence level}$

$\delta = \Pr[\text{fail}]$

$$[\hat{a}, \hat{b}] \neq [a^{\text{true}}, b^{\text{true}}]$$

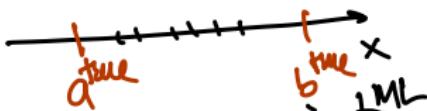
② Assume  $a^{\text{true}}, b^{\text{true}}$  known

$$\Pr [\hat{a} \in [\hat{a}^{\text{ML}}, \hat{b}^{\text{ML}}] \mid a^{\text{true}}, b^{\text{true}}] = \Pr \left[ \max_{i=1:n} x_i = \hat{a}^{\text{ML}} \right]$$

$$\min x_i = \hat{a}^{\text{ML}}$$

$$l^* = b^{\text{true}} - a^{\text{true}}$$

$$l = \hat{b}^{\text{ML}} - \hat{a}^{\text{ML}} = \max \hat{a} - \min \hat{b}$$



## ML estimation for continuous distributions

Side problem  $X^{1:n} \sim P$

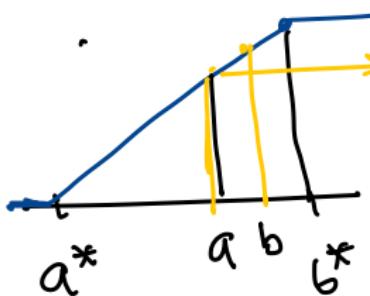
What is  $\max X^{1:n} \sim P^{\text{max}}$ ?

$$P[X < a] = F(a)$$

$$P[X^{1:n} < a] = \prod_{i=1}^n F(a) = F^n(a)$$

iid ↑

for  $u[a^*, b^*]$



$$F(a) = \frac{a - a^*}{b^* - a^*}$$

$$\begin{aligned} F(b) - F(a) &= \\ &= \frac{b - a}{b^* - a^*} \end{aligned}$$

## ML estimation for continuous distributions

Side pb part 2:

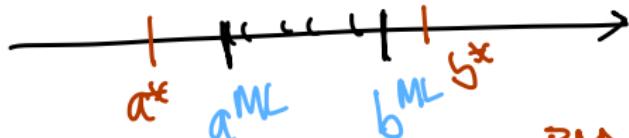
$$\Pr [ X^{1:n} \in [a,b] ] = \underset{n}{\underbrace{[F(b) - F(a)]^n}}$$

$$\Pr [ X \in [a,b] ] = F(b) - F(a)$$

Back to uniform  $[a^*, b^*]$ :

$$\Pr [ X^{1:n} \in [a,b] \mid a^*, b^* ] = \left( \frac{b-a}{b^*-a^*} \right)^n \quad \text{with } n$$

**STATISTICS**



BAD

$\frac{e}{e^*}$  small

$$= \left( \frac{e}{e^*} \right)^n = \delta$$

$\Pr [ \text{BAD} ]$

calculus

## ML estimation for continuous distributions

$$\left(\frac{\ell}{\ell^*}\right)^n = \delta \Rightarrow \frac{\ell}{\ell^*} = \sqrt[n]{\delta}$$

$b^* - a^* \leftarrow \text{from data}$

$\ell^* = \frac{\ell}{\sqrt[n]{\delta}}$

confidence

$$\Rightarrow b^* - a^* = \ell^* = (b - a) \frac{1}{\sqrt[n]{\delta}} = \Delta \text{ correction}$$

$$\ell^* - \ell = \ell (\delta^{-\frac{1}{n}} - 1)$$

$$a^* = a^{ML} - \frac{\ell^* - \ell}{2} = a^{ML} - \frac{1}{2} (b^{ML} - a^{ML}) (\delta^{-\frac{1}{n}} - 1)$$

$$b^* = b^{ML} + \frac{\ell^* - \ell}{2} = b^{ML} - \frac{1}{2} (b^{ML} - a^{ML}) (\delta^{-\frac{1}{n}} - 1)$$

→ corrected estimates

## ML estimation for continuous distributions

$$\boxed{(b-a) \frac{1}{\sqrt{\frac{n}{\delta}}}} = \Delta \text{ correction}$$

$\delta \downarrow$  want to be safer  $\Rightarrow \frac{1}{\delta} \uparrow \Rightarrow \Delta \uparrow$

$n \uparrow$  more data safer

$\sqrt[n]{\delta} \rightarrow 1$  for  $n \rightarrow \infty \Rightarrow \Delta \rightarrow b^{\text{ML}} - a^{\text{ML}}$

$$\delta \in (0, 1)$$

## ML estimation for continuous distributions

$$S = [0, \infty)$$

Exponential

$$f_\lambda(x) = \lambda e^{-\lambda x} \quad \begin{matrix} x > 0 \\ \lambda > 0 \end{matrix}$$

Model family  $\mathcal{F} = \{f_\lambda, \lambda > 0\}$

Data  $x^{1:n} > 0$

Likelihood  $L(\lambda) = \prod_{i=1}^n (\lambda e^{-\lambda x_i}) = \lambda^n e^{-\lambda \sum_{i=1}^n x_i}$

log-l  $\ell(\lambda) = n \ln \lambda - \lambda \sum_{i=1}^n x_i$

STAT  
CALC

ML  $\lambda^{ML} = \operatorname{argmax} \ell(\lambda)$

$$\lambda > 0$$

## ML estimation for continuous distributions

$$\ell(\lambda) = \underline{n} \cdot \ln \underline{\lambda} - \underline{\lambda} \sum_{i=1}^n x_i$$

$$\lambda^{ML} = \operatorname{argmax}_{\lambda > 0} \ell(\lambda)$$

$$\ell'(\lambda) = \frac{n}{\lambda} - \sum_{i=1}^n x_i = 0$$

$$\frac{1}{\lambda^{ML}} = \frac{1}{n} \sum_{i=1}^n x_i$$

$$> 0$$

mean  
of data!  
Sufficient  
statistic

Uniform  
Exponential  
Family

$$\lambda^{ML} = \frac{1}{\frac{1}{n} \sum_{i=1}^n x_i}$$