

# Lecture Notes VI.2 – Double descent on a simple linear regression example

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**Reading:** Ch.10.2.3

## Linear regression when $n < d$

- ▶ We describe a very simple linear regression situation (following Bach section 10.2.3)
- ▶ For it, we are able to explicitly obtain the expected estimation error  $E[\|\theta_{true} - \hat{\theta}\|^2]$
- ▶ Surprisingly, the variance of this error **decreases with  $d$** , and the error itself has a limit proportional to  $\|\theta_{true}\|^2$ .
- ▶ Input distribution  $x^{1:n} \sim N(0, I_d)$ , noise  $\epsilon^{1:n} \sim N(0, \sigma^2)$
- ▶ Model  $y^i = (x^i)^T \theta_{true} + \epsilon^i$ .
- ▶ Denote  $X \in \mathbb{R}^{n \times d}$ ,  $y, \epsilon \in \mathbb{R}^n$  the usual input matrix, output, and noise vectors respectively
- ▶ Denote  $K = XX^T \in \mathbb{R}^{n \times n}$  the Gram matrix (or kernel matrix). We assume  $K$  is non-singular
- ▶ From Lecture IV, **The Implicit Bias of Gradient Descent** we know that
  - ▶ When  $X$  is full rank  $n$ , the equation  $y = X\theta$  has multiple solutions  $\theta$
  - ▶ Gradient Descent converges to the min norm solution  $\hat{\theta} = X^T K^{-1} y$

The expected estimation error  $MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2]$

- Decompose  $\hat{\theta}$

$$\hat{\theta} = X^T K^{-1} y = X^T K^{-1} (X \theta_{true} + \epsilon) = X^T K^{-1} X \theta_{true} + X^T K^{-1} \epsilon \quad (1)$$

- Then,

$$MSE(\theta_{true}) = E_{X, \epsilon} [\|\theta_{true} - \hat{\theta}\|^2] \quad (2)$$

$$= \underbrace{E_X [\theta_{true}^T (I_d - X^T K^{-1} X) \theta_{true}]}_{\text{bias}^2} + \underbrace{E_{X, \epsilon} [\epsilon^T K^{-1} X X^T K^{-1} \epsilon]}_{\text{variance}} \quad (3)$$

- The Variance term becomes

$$Var = E_{X, \epsilon} [\epsilon^T K^{-1} \epsilon] \quad (4)$$

$$= E_X [\text{trace } K^{-1}] \sigma^2 \quad \text{Wishart!} \quad (5)$$

$$= \sigma^2 \frac{n}{d - n - 1} \quad (6)$$

The expected estimation error  $MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2]$

- ▶ The **Bias<sup>2</sup>** term:
- ▶ Note that  $\theta_P = X^T K^{-1} X \theta_{true}$  is the orthogonal projection of  $\theta_{true}$  on the row space of  $X$ , and  $\theta_{true}^T X^T K^{-1} X \theta_{true} = \|\theta_P\|^2$ .
- ▶ The subspace is a random subspace of dimension  $n$  in  $\mathbb{R}^d$ . By spherical symmetry, the length of the projection of a fixed vector on a random subspace is the same with that of a projecting a random vector of length (squared)  $\|\theta_{true}\|^2$  on a fixed subspace, e.g. the first  $d$  unit vectors in  $\mathbb{R}^d$ . The latter expected value is easy to compute

$$E[\|\theta_P\|^2] = \frac{n}{d} \|\theta_{true}\|^2 \quad (7)$$

**Exercise** Proving this is a moderately easy exercise

- ▶ Hence,

$$\text{bias}^2 = E_X[\theta_{true}^T (I_d - X^T K^{-1} X) \theta_{true}] = \frac{d - n}{d} \|\theta_{true}\|^2 \quad (8)$$

The expected estimation error  $MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2]$

► Finally

$$MSE(\theta_{true}) = E[\|\theta_{true} - \hat{\theta}\|^2] = \frac{d-n}{d} \|\theta_{true}\|^2 + \sigma^2 \frac{n}{d-n-1} \quad (9)$$

for  $d > n + 1$

► When  $d \rightarrow \infty$ , the variance  $\rightarrow 0$  and the  $\text{bias}^2 \rightarrow \|\theta_{true}\|^2$