

Lecture 18

Gibbs Sampling

- Project + data POSTED
- Some lecture + Lab time
→ about Project
- HW 6 posted
- HW 7 will be last HW

& Congrats to
the many
who have
already
Participation —
don't miss the last
chances participated
A lot !!

Lecture Notes VIII – Markov Chain Monte Carlo

Marina Meila
`mmp@stat.washington.edu`

Department of Statistics
University of Washington

April 2025

Notation ✓

Gibbs sampling ←

The detailed balance ←

Metropolis-Hastings sampling

Notation

- ▶ $V = \{X_1, \dots, X_n\}$ set of random variables (nodes of a graphical model)
also known as Markov network
- ▶ $X_{1:n} \in \{\pm 1\}$ (for simplicity)
- ▶ $E = \{(i, j), 1 \leq i < j \leq n\}$ edges of graph
- ▶ graph is not complete, some edges are missing
- ▶ We write $i \sim j$ for $(i, j) \in E$ or $(j, i) \in E$
- ▶ $\text{neigh}_i = \text{neighbors of } X_i$
- ▶ **Markov property** $X_i \perp \text{all other variables} \mid \text{neigh}_i$
- ▶ $x = (x_1, \dots, x_n) \in \{\pm 1\}^n = S$ an assignment to all variables in V

Notation

- ▶ $V = \{X_1, \dots, X_n\}$ set of random variables (nodes of a **graphical model**)
also known as **Markov network**
- ▶ $X_{1:n} \in \{\pm 1\}$ (for simplicity)
- ▶ $E = \{(i, j), 1 \leq i < j \leq n\}$ **edges** of graph
- ▶ graph is not complete, some edges are missing
- ▶ We write $i \sim j$ for $(i, j) \in E$ or $(j, i) \in E$
- ▶ $\text{neigh}_i = \text{neighbors of } X_i$
- ▶ **Markov property** $X_i \perp \text{all other variables} \mid \text{neigh}_i$
- ▶ $\mathbf{x} = (x_1, \dots, x_n) \in \{\pm 1\}^n = S$ an assignment to all variables in V
- ▶ Distribution over S

$$P(\mathbf{x}) = \frac{1}{Z} e^{-\phi(\mathbf{x})} \quad \text{with } \phi(\mathbf{x}) = \sum_{i=1}^n h_i x_i + \sum_{(i,j) \in E} h_{ij} x_i x_j \quad (1)$$

(and $h_{ij} > 0$ for all $(i, j) \in E$)

- ▶ $Z = \sum_{\mathbf{x} \in S} e^{-\phi(\mathbf{x})}$
- ▶ Usually, intractable to compute Z

Wanted samples $\mathbf{x}^1, \mathbf{x}^2, \dots$ from P

n variables $\{x_{1:n}\} = V$
 $\{(i,j)\} = E$ edges

$$S = \{\pm 1\}^n \quad x_i \in \{\pm 1\}$$

$$|S| = 2^n$$

\Rightarrow can't calculate
 $Z = \sum_{x \in S} e^{-\varphi(x)}$

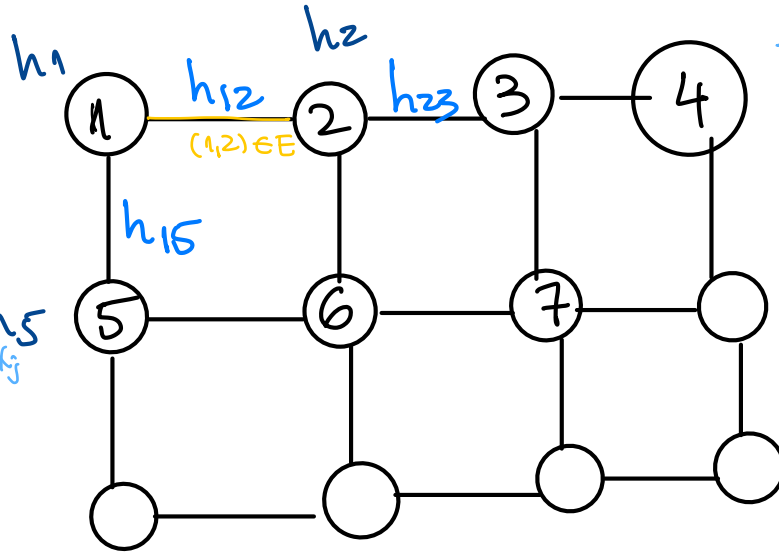
Ising Model
 our running example

$$\rightarrow P(x) = ?$$

know $e^{-\varphi(x)}$ unnormalized

$$\frac{P(x)}{P(x)} = \frac{e^{-\varphi(x)}}{e^{-\varphi(x)}} = \dots$$

Want: samples
 from $P(x)$ iid



$$P(x) = \frac{1}{Z} e^{-\varphi(x)}$$

$$x = (x_1 \dots x_n)$$

$$\varphi = \sum_{i \in V} h_i x_i + \sum_{(i,j) \in E} h_{ij} x_i x_j$$

Simplify:

$$h_1 = h_2 = \dots = 0$$

$$h_{12} = h_{15} = \dots = 1$$

$$\Rightarrow \varphi(x) = \sum_{(i,j) \in E} x_i x_j$$

Gibbs sampling idea

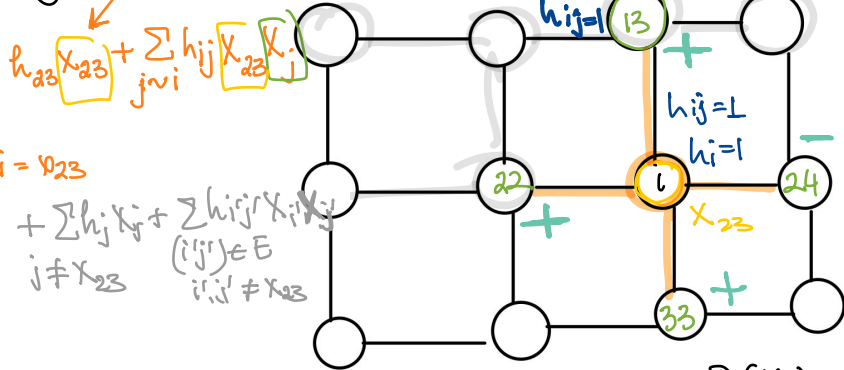
$$P(x) = \frac{1}{Z} e^{-\phi(x)} \quad \text{with } \phi(x) = \sum_{i=1}^n h_i x_i + \sum_{(i,j) \in E} h_{ij} x_i x_j$$

- ▶ We cannot sample directly from P
- ▶ But we can sample each $X_i \sim P_{X_i | X_{-i}} = P_{X_i | \text{neigh}_i}$ for any $i = 1 : n$
Pr X_i | all other variables

$$P_{X_i | \text{neigh}(X_i)}$$

$$\varphi(x) = \sum h_i x_i + \sum h_{ij} x_i x_j$$

$$= \frac{e^{-\varphi(x)}}{e^{-\varphi(x_{23}=1, x_{-23})} + e^{-\varphi(x_{23}=-1, x_{-23})}} = \frac{e^{-(h_{23}x_{23} + \sum_{j \neq i} h_{ij} x_{23} x_j)}}{e^{-(+1)[h_{23} + \sum_{j \neq i} h_{ij} x_j]} + e^{-(-1)[\quad]}}$$



$$= \frac{e^{-x_{23} \cdot A}}{e^{-A} + e^{+A}}$$

$$x_{23} = \pm 1$$

$$A = 1 + (1 + 1 + 1) = 3$$

$$P_{x_{23} | x_{22}, x_{24}, x_{33}, x_{13}, x_{11}, x_{12}, \dots} = \frac{P(x)}{P(x_{-23})}$$

$$P_{V \setminus x_{23}} = P(x_{23}=+1, x_{-23}) + P(x_{23}=-1, x_{-23})$$

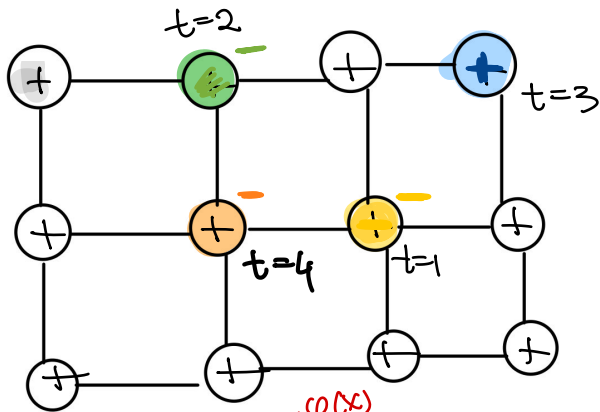
← marginal all $\setminus x_{23}$

$$P(x_{23}=+1 | x_{-23}) = \frac{e^3}{e^3 + e^{-3}} = \pi_i^+$$

Theorem

Ergodic Markov chain
 x^0 initial state

$$P_x^{(t)} \rightarrow P_x^\infty \text{ for any } x_0$$



To show:

$$x^t \sim P^* \neq e^{-\varphi(x)}$$

$x^{t'}$ independent

to show $P_x^\infty = P^*$

Gibbs sampling idea

$$P(x) = \frac{1}{Z} e^{-\phi(x)} \quad \text{with } \phi(x) = \sum_{i=1}^n h_i x_i + \sum_{(i,j) \in E} h_{ij} x_i x_j \quad \leftarrow \text{Ising}$$

- ▶ We cannot sample directly from P
- ▶ But we can sample each $X_i \sim P_{X_i|X_{-i}} = P_{X_i|\text{neigh}_i}$ for any $i = 1 : n$
- ▶ Why? For any x_{-i} let

$$\pi_i^+ = Pr[X_i = +1 | x_{-i}] \quad \pi_i^- = Pr[X_i = -1 | x_{-i}] \quad \pi_i^+ + \pi_i^- = 1 \quad (2)$$

$$\pi_i^+ = \frac{P(X_i = +1, x_{-i})}{P(x_{-i})} = \frac{P(X_i = +1, x_{-i})}{P(X_i = +1, x_{-i}) + P(X_i = -1, x_{-i})} \quad (3)$$

$$\pi_i^- = \frac{P(X_i = -1, x_{-i})}{P(X_i = +1, x_{-i}) + P(X_i = -1, x_{-i})} \quad \text{hence} \quad (4)$$

$$\frac{\pi_i^+}{\pi_i^-} = \frac{P(X_i = +1, x_{-i})}{P(X_i = -1, x_{-i})} = \frac{e^{-\phi(X_i = +1, x_{-i})}}{e^{-\phi(X_i = -1, x_{-i})}} = e^{-\phi(X_i = +1, x_{-i}) + \phi(X_i = -1, x_{-i})} \quad (5)$$

$$= e^{-2h_i - 2 \sum_{j \sim i} h_{ij} x_j} \quad (6)$$

$$1 = \pi_i^+ + \pi_i^- \quad \text{hence } \pi_i^+ = \frac{e^{-2h_i - 2 \sum_{j \sim i} h_{ij} x_j}}{1 + e^{-2h_i - 2 \sum_{j \sim i} h_{ij} x_j}} = \frac{e^{-h_i - \sum_{j \sim i} h_{ij} x_j}}{e^{h_i + \sum_{j \sim i} h_{ij} x_j} + e^{-h_i - \sum_{j \sim i} h_{ij} x_j}} \quad (7)$$

- ▶ $X_i | \text{neigh}_i \sim \text{Bernoulli}(\pi_i^+)$

Gibbs sampling algorithm

1. Initialize x^0 with some arbitrary values
2. For $t = 1, 2, \dots$ we will sample sequentially $x^t | x^{t-1}$ as follows
 - 2.1 Pick $i \in 1 : n$ uniformly at random
 - 2.2 Sample $X_i^t | \text{neigh}_i \sim \text{Bernoulli}(\pi_i^+)$
 - 2.3 Every T steps (where T is a LARGE number), output x^t

Q Why does this work?

- A1 We are sampling from a **Markov chain** on S (transition probability matrix P on next page)
- A2 If we take enough steps T , the distribution converges to the stationary distribution of this chain, let's call it P . We take a sample from it x^T
- A3 If we continue for another T steps, the chain has "forgotten" about x^T ; the new sample x^{2T} is independent of x^T . Etc, samples $x^T, x^{2T}, \dots, x^{NT}$ are i.i.d. from P^∞
- TODO To show that $P^\infty = P$ the distribution we wanted to sample from.