



Lecture 11

Training predictors

all material = Q2 - Thu 11/10 < today HW4 posted (ap to eto) LIV-1 posted

Lecture IV: Training predictors, Part I

Marina Meilă mmp@stat.washington.edu

> Department of Statistics University of Washington

October, 2022

Analytic optimization

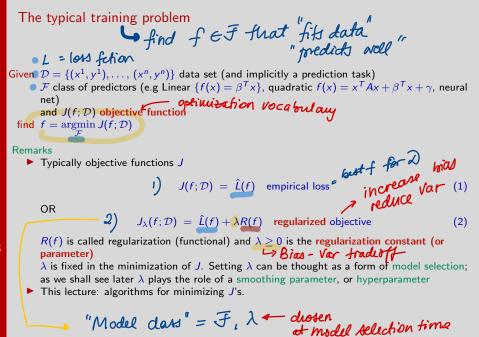
Optimization generalities Optimization glossary Descent methods zoo

Descent methods

The steepest descent method Line minimization algorithms The Newton-Raphson method Examples: Logistic regression and Backpropagation

Reading HTF Ch.: Lasso 3.1,2,4, Logistic regression 4.4, Neural networks 11, Murphy Ch.: Ridge regression (including numerics) 7.5, Descent methods 8.3.2,3,5, Neural networks 16.5.1–4, Autoencoders 28, Bach Ch.: 5.

For more mathematical background, look e.g. at "Numerical recipes" chapter 10 or, for really advanced treatment Nocedal and Wrigth (Ch 3, 6).



farina Meila: Lecture IV: Training predictors, B

Model Selection S۷ Training $f'' \in \mathcal{F}_1$ $\widehat{\mathcal{Q}}$ given F1 (lihear) model class 1 $\in \mathcal{F}_2$ Fz (CART E F2 kernel h \mathcal{F}_3 3 Gauttian $f^{(4)} \in \mathcal{F}_{4}$ h F4 choose e.g. Lasso EF 5 F (-linear, R, 2) predictor with differently f⁽⁶⁾ € F -u-, R. F OR Ridge R, Lasso, OLS Model Selection 1 same Fr different \$

Linear Least Squares regression

- Problem $\mathcal{F} = \{f(x) = \beta^T x\}, L_{LS}(y, f(x)) = (y f(x))^2, y \in \mathbb{R}, J = \hat{L}$ Solution

Finding $f \in \mathcal{F}$ is equivalent to finding $\beta \in \mathbb{R}^d$ (or \mathbb{R}^{d+1})

 $\mathcal{D} = [(x'y')]_{i=r:n}$

Linear Least Squares regression

- Problem $\mathcal{F} = \{f(x) = \beta^T x\}, L_{LS}(y, f(x)) = (y f(x))^2, y \in \mathbb{R}, J = \hat{L}$
- Solution
 - Finding $f \in \mathcal{F}$ is equivalent to finding $\beta \in \mathbb{R}^d$ (or \mathbb{R}^{d+1})
- define data matrix or (transpose) design matrix $\begin{pmatrix} d \text{ exign} \end{pmatrix} \\ X = \begin{bmatrix} \begin{pmatrix} x^1 \end{pmatrix}^T \\ \begin{pmatrix} x^2 \end{pmatrix}^T \\ \vdots \\ \begin{pmatrix} x^1 \end{pmatrix}^T \\ \vdots \\ \begin{pmatrix} x^n \end{pmatrix}^T \\ \vdots \\ \begin{pmatrix} y^n \end{pmatrix}^T \end{bmatrix} \in \mathbb{R}^{\mathbb{N} \times \mathbb{Q}} \text{ and } Y = \begin{bmatrix} y^1 \\ y^2 \\ \vdots \\ y^n \end{bmatrix}, E = \begin{bmatrix} \varepsilon^1 \\ \varepsilon^2 \\ \vdots \\ \varepsilon^n \end{bmatrix} \in \mathbb{R}^d$
 - Then we can write

$$Y = X\beta + E$$

OLS

(attains global min)

• The solution $\hat{\beta}$ is chosen to minimize the sum of the squared errors

minimize the sum of the squared errors (Ordinary Linear

$$J(\beta) = \hat{L}_{LS} = \sum_{i=1}^{n} (y^{i} - \beta^{T} x_{i})^{2} = ||E||^{2}$$
Repression) (3)
it

$$\beta = (X^{T}X)^{-1}X^{T}y$$
- closed form
. exact (4)
minimization

which gives Exercise Derive it

Linear Least Squares regression

- ▶ Problem $\mathcal{F} = \{f(x) = \beta^T x\}, L_{LS}(y, f(x)) = (y f(x))^2, y \in \mathbb{R}, J = \hat{L}$
- Solution
 - Finding $f \in \mathcal{F}$ is equivalent to finding $\beta \in \mathbb{R}^d$ (or \mathbb{R}^{d+1})
 - define data matrix or (transpose) design matrix F (1)T 7

$$= \begin{pmatrix} \begin{pmatrix} x \\ y^{2} \end{pmatrix}^{T} \\ \begin{pmatrix} x^{2} \end{pmatrix}^{T} \\ \begin{pmatrix} x^{i} \end{pmatrix}^{T} \\ \begin{pmatrix} x^{i} \end{pmatrix}^{T} \\ \begin{pmatrix} x^{i} \end{pmatrix}^{T} \end{pmatrix} \in \mathbb{R}^{N \times n} \text{ and } Y = \begin{bmatrix} y^{1} \\ y^{2} \\ \vdots \\ y^{n} \end{bmatrix}, E = \begin{bmatrix} \varepsilon^{1} \\ \varepsilon^{2} \\ \vdots \\ \varepsilon^{n} \end{bmatrix} \in \mathbb{R}^{d}$$
write
$$Y = X\beta + E$$

$$\hat{\beta} \text{ is chosen to minimize the sum of the squared errors}$$

$$max \text{ Likelihood (b)}$$

$$min - lw$$

$$(Likelihood (b))$$

Then we can write

Х

$$Y = X\beta + E$$

b The solution $\hat{\beta}$ is chosen to minimize the sum of the squared errors

$$J(\beta) = \hat{L}_{LS} = \sum_{i=1}^{n} (y^{i} - \beta^{T} x_{i})^{2} = ||E||^{2}$$

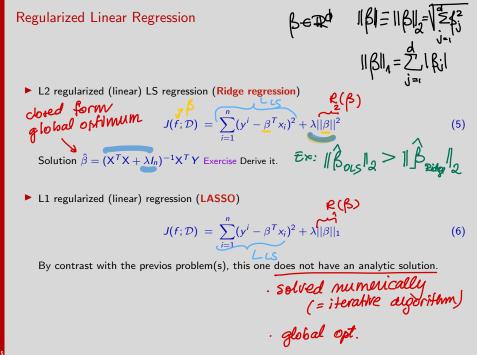
which gives Exercise Derive it

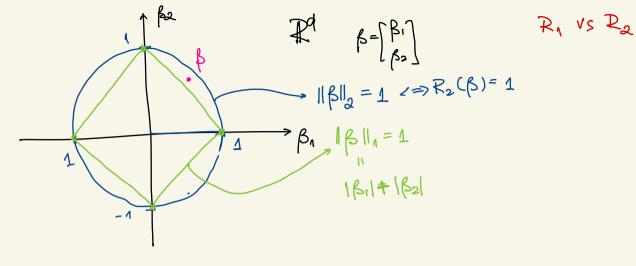
$$\hat{\beta} = (\mathbf{X}^{\mathsf{T}}\mathbf{X})^{-1}\mathbf{X}^{\mathsf{T}}\mathbf{y}$$
(4)

(3)

Underlying statistical model $y = \beta^T x + \varepsilon$, $\varepsilon \sim N(0, \sigma^2)$ (and $(x^{1/2}, y^{1/2})$ sampled I.d., of course) Define the negative log-likelihood $L_{logi} = -\ln P(y|x,\beta)$ \longrightarrow L_{logi} \iff Max Likelihood i.i.d., of course)

- Then, $(y \beta^T x)^2 = \sum_{i=1}^n (\varepsilon^i)^2 = -2\sigma^2 \ln P(y^i | x^i, \beta) = -2\sigma^2 \hat{L}_{logl}$
- Hence, $\hat{\beta}$ from (4) is the Maximum Likelihood (ML) estimator of the parameter β and the minimizer of \hat{L}_{logl}
- **\triangleright** This is an example where the minimizer of J(f) has an analytical formula. A few more examples of this kind follow.





$$\begin{split} \|\beta\|_{2} < 1 \\ \|\beta\|_{1} > 1 > \|\beta\|_{2} \\ X^{T}y = b \\ A = X^{T}X \\ \beta_{OLS} = A^{T}b \\ \beta_{R} = (A + \lambda I)^{T}b \\ \|\beta\|_{2}^{2} = \beta^{T}\beta \end{split}$$

Generative models for classification

- If each class conditional distribution P_{X|Y=y} can be estimated by an analytic formula, then the generative classifier can be estimated analytically
- Examples LDA, QDA

Example (Naive Bayes for text classification)

The bag of words model of text. Let $D = \{words\}$ be a dictonary. For simplicity, we shall assume $D = \{$ and, average, batting, score, variance $\}$. We are given a corpus \mathcal{D} containing N documents of two classes { sports= -1, statistics= 1}. For each document in \mathcal{D} , we form a vector $x \in \{0,1\}^{|D|}$ by setting $x_w = 1$ if the document contains word w and 0 otherwise. For example, the document "Reddick's batting average is 0.50" has $x = [0\,11\,0\,xs0]$. The x vectors are the data to which we fit a Naive Bayes model. Assume our toy corpus is now the table on the left, and the estimated class conditional distributions are on the right.

		X			У							
0	1	1	0	0	-1							
1	1	0	1	0	-1							
1	0	1	1	0	-1							
1	0	1	0	0	-1		$P_{X_j=1 Y=-1}$ y					
1	0	1	0	0	-1	.8	.4	.8	.4	0	-1	$P_Y(1)=0.5=p$
1	1	0	0	1	1	.8	.6	0	.2	.6	1	
0	0	0	1	0	1						·	
1	1	0	0	1	1							
1	1	0	0	0	1							
1	0	0	0	1	1							

Example (Naive Bayes for text (continued))

Before we go further, we will "adjust" the 0 and 1 estimated probabilities for "variance" and "batting", setting them to 1/M, respectively 1 - 1/M for some "large" M = 10 Exercise Do you think this problem occurs often in real applications? Find a statistical framework to justify the adjustment. The probability of a new document, e.g x = [10010] in class 1, respectively -1 is

$$P(x|y = -1) = 0.8 \times (1 - 0.4) \times (1 - 0.8) \times 0.4 \times 0.9$$
(7)

$$= 0.8^{x_1}(1-0.8)^{1-x_1} \times 0.4^{x_2}(1-0.4)^{1-x_2} \dots$$
(8)

$$P(x|y=1) = 0.8 \times (1-0.6) \times (1-0.1) \times 0.2 \times (1-0.6)$$
(9)

The NB classifier we obtain is

$$f(x) = \ln \frac{p}{1-p} + x_1 \ln \frac{P_{and|1}}{P_{and|-1}} + x_2 \ln \frac{P_{average|1}}{P_{average|-1}} + x_3 \ln \frac{P_{batting|1}}{P_{batting|-1}} + x_4 \ln \frac{P_{score|1}}{P_{score|-1}} + x_5 \ln \frac{P_{variance|1}}{P_{variance|-1}} + (1-x_1) \ln \frac{1-P_{and|1}}{1-P_{and|-1}} + (1-x_2) \ln \frac{1-P_{average|-1}}{1-P_{average|-1}} + (1-x_3) \ln \frac{1-P_{batting|-1}}{1-P_{batting|-1}} + (1-x_4) \ln \frac{1-P_{batting|-1}}{1-P_{ba$$

which evaluates to $f([10010]) = \ln 1 + \ln \frac{3}{2} + \ln \frac{2}{9} + \ln 2 + \ln \frac{9}{4} = 0.405 > 0$ Notes As you saw above, we are not required to include all possible words in the dictionary Exercise Find some reasons why. A common preprocessing step stemming which aims to map words in the same word family to a single w; e.g "batting" \rightarrow "bat". Sometimes stop words like "the", "and" are removed.

Most predictors can't be estimated analytically

- Unfortunately, minimizing J analytically is possible only in a handful of examples.
- In all other cases, we find f = argmin J by numerical/iterative methods also called search (or training/learning, of course). For example
 - CART algorithm Exercise What J is the CART algorithm minimizing?
 - Perceptron algorithm (will be revisited this lecture)
- Therefore now we study generic algorithms for finding minima of functions of *n* variables. This is called (numerical) optimization.

Optimization glossary

Change in notation!

Here, f is a function to be minimized, and x the variable in the domain of the function. In a learning task, f will be replaced by an objective J like \hat{L}_{logl} and x by the parameters of the predictor, e.g. w, β, θ, \ldots

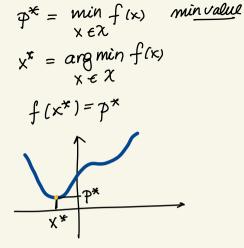
The methods in this lecture belong to the class of unconstrained optimization methods. **Problem** Find $\min_x f(x)$ for $x \in \mathbb{R}^d$ or $x \in D$ the domain of f. We assume that f is a twice differentiable function with continuous second derivatives. **Notation** The gradient of f is the column vector

$$\nabla f(x) = \left[\frac{\partial f}{\partial x_i}(x)\right]_{i=1}^n \equiv \frac{\partial f}{\partial x} \longrightarrow \mathbb{P}^d$$
(11)

Change notation min f(x) 76: x e X

and the Hessian of f is the square symmetric matrix of second partial derivatives of f

$$\nabla^2 f(\mathbf{x}) = \left[\frac{\partial^2 f}{\partial x_i x_j}(\mathbf{x})\right]_{i,j=1}^n \in \mathbb{R}^{d \times d}$$
(12)



Optimization basics

 $A \in \mathbb{R}^{d \times d}$ p<u>osifive definite</u> (⇒ X^TAX ≥0 for all X ETR • A %0 $\langle \Rightarrow \lambda(A) \ge 0$ · A>o strictly p.d. for all e-values for X=0 <=> XTAX >0 for all e-values \mathcal{E}_{X} : $\nabla^{2}_{f}(\tilde{X}) > 0$ $\langle \neg \rangle$ X* isolated local & strict - (X*) =0

Optimization basics

• $\nabla f(\mathbf{x}^{\star}) = 0$, $\nabla^2 f(\mathbf{x}^{\star}) \ge 0 \Rightarrow \exists \lambda = 0 \text{ e-value},$ $\downarrow 0 \qquad n = e \text{-vector}$ $|\mathbf{x}^3|$ $\mathbf{x}^{\star} \qquad \mathbf{x}^{\star}$

Local and global minima

A local minimum for f is point x^* for which

 $f(x^*) \leq f(x)$ whenever $||x - x^*|| < \epsilon$

loca

global mir

A global minimum for f is point x^* for which

 $f(x^*) \leq f(x)$ for all x in the domain of f

We say x^* is a strict local/global minimum when the above inequalities are strict for $x \neq x^*$.

- A minimum is **isolated** if it is the only local minimum in an ϵ -ball around itself.
- A stationary point for f is a point x^* for which $\nabla f(x^*) = 0$.

Theorem

If f has continuous second derivative everywhere in D, and $x^* \in D$ is a point for which $\nabla f(x^*) = 0$ and $\nabla^2 f(x^*) \ge 0$ ($\nabla^2 f(x^*) > 0$) then x^* is a (nonsingular) local minimum for f.

In what follows, we will deal only with non-singular local minima. A non-singular local minimum is strict and isolated.

Descent methods

Many unconstrained optimization methods for finding a local minimum are of the form:

$$x^{k+1} = x^k + \eta^k d^k \tag{13}$$

where $d^k \in \mathbb{R}^d$ represents an (unnormalized) direction and $\eta^k > 0$ is a scalar called the step size.

Direction choice

- gradient based $d^k = -D^k \nabla f(x^k)$ with $D^k \in \mathbb{R}^{n \times n}$
 - **•** steepest descent $D^k = I$
 - stochastic gradient (more about it later)
 - Newton-Raphson $D^k = \nabla^2 f(x^k)^{-1}$
 - conjugate gradient implicity multistep rescaling of the axes "equivalent" to $D^k = \nabla^2 f(x^k)^{-1}$
 - quasi-Newton implicit multistep approximation of $D^k = \nabla^2 f(x^k)^{-1}$
- non-gradient based
 - coordinate descent d^k = one of the basis vectors in \mathbb{R}^d

Step size choice

- line minimization $\eta^k = \min_{\eta} f(x^k + \eta d^k)$
- Armijo rule (also called Backtracking) = search but not minimization
- constant step size $\eta^k = s$
- diminishing step size $\eta^k \to 0$; $\sum_k \eta^k = \infty$

