

STAT 534

Lecture

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1 Rejection Sampling

- Goal: Get independent samples from $f(x)$ on domain Ω
- Given: $g(x), \frac{f(x)}{g(x)} \leq M, g \gg f \Leftrightarrow f > 0 \Rightarrow g > 0$
- Algorithm

Algorithm 1 Rejection sampling algorithm

```
1: procedure REJECTION SAMPLING()  
2:   sample  $x \sim g$ ,  
3:   sample  $u \sim \text{uniform}[0,1]$   
4:   until  $u \leq \frac{f(x)}{g(x)} \leq 1$   
5:   return  $x$ 
```

Notice that the acceptance probability $P[\text{accept}] \propto \frac{1}{M}$:

$$P[\text{accept}] = \int_{\Omega} g(x)P[\text{accept}|x]dx = \int_{\Omega} \frac{f(x)}{Mg(x)} = \frac{1}{M}$$

This property has implication of how to choose proper $g(x)$. A proper $g(x)$ will yield a smaller M value so that the acceptance probability will not be too small.

Also, It can be confirmed that the sample generated using the procedure above is actually from the true $f(x)$ distribution:

$$\hat{f}(x) = \frac{g(x)a(x)}{P[\text{accept}]} = g(x)\frac{f(x)}{Mg(x)} = f(x)$$

- Example: Sampling from "tail" of $f(x)$
Define $I_{(x \in A)}$ where $A = \{x | |x| > c\}$ for some constant c . Then from the relationship between $f(x)$ and $g(x)$:

$$f(x) \propto g(x)I_{[x \in A]}$$

So the target $f(x)$ can be calculated as below:

$$f(x) = \frac{g(x)I_{x \in A}}{P_g[A]}$$

Then M can be calculated as:

$$M = \frac{1}{P_g[A]}$$

It can be seen that if the tail probability of $f(x)$ is very small, then M is a very large number which means that the acceptance probability will be very small. The algorithm will not be efficient.

2 Importance Sampling

- Goal: $E_f[h(x)]$ ($x \in \Omega, h : \Omega \in \mathbb{R}$)
- Given: $g(x), g > 0$ whenever $f > 0$
- Algorithm

Algorithm 2 Importance sampling algorithm

- 1: **procedure** IMPORTANCE SAMPLING()
 - 2: **for** $t=1:T$ **do**
 - 3: sample $x_t \sim g$
 - 4: $\bar{h} = \frac{1}{T} \sum_{t=1}^T h(X_t) \frac{f(x)}{g(x)}$
-

It can be confirmed that $E_g[\bar{h}]$ is actually $E_f[h]$:

$$E_f[\bar{h}] = \int_{\Omega} h(x)f(x)dx \approx \bar{h} = \frac{1}{T} \sum_{t=1}^T h(X_t)$$

$$\begin{aligned} E_g[\bar{h}] &= \frac{1}{T} \sum_{t=1}^T E_g[h(X_t) \frac{f(x)}{g(x)}] \\ &= E[h(X_t) \frac{f(x)}{g(x)}] \\ &= \int_{\Omega} h(x) \frac{f(x)}{g(x)} dx \\ &= E_f[h] \end{aligned}$$

Notice that $Var(\bar{h}) \propto \frac{1}{T}$:

$$Var(\bar{h}) = \frac{1}{T^2} \times T \times Var\left(\frac{h(x)f(x)}{g(x)}\right) \propto \frac{1}{T}$$

Since we want the variance to be small, then ideally: $g(x) \propto h(x)f(x)$

3 Metropolis-Hastings Sampling

- Goal: Simulate samples from target distribution π
- Given: Proposal distribution $g(x' | x)$
- Algorithm

Algorithm 3 Metropolis-Hastings algorithm

```

1: procedure METROPOLIS-HASTINGS()
   Initialize  $x^{(0)} \sim q(x)$ 
2: for iteration  $t=1,2,3\dots$  do
3:   sample  $x' \sim q(x' | x_{t-1})$  ▷ proposal distribution
4:    $a = \min(1, \frac{\pi(x')q(x_{t-1}|x')}{\pi(x_{t-1})q(x'|x_{t-1})})$  ▷ Acceptance Probability
5:    $x_t = \begin{cases} x' & \text{with probability } a \\ x_{t-1} & \text{with probability } 1 - a \end{cases}$ 

```

- Example: Bayesian estimation for Probit Model
Given that $\mathbf{x} \in \mathbb{R}^d, y \in \{0, 1\}$, and $P[y = 1 | \mathbf{x}] = \Phi(\beta^T \mathbf{x})$ where Φ is the cdf of standard normal distribution.

The goal is to find the posterior distribution of β

$$P(\beta | (x_i, y_i)_{i=1:n}, \pi_0)$$

where $(x_i, y_i)_{i=1:n}$ is the data and π_0 is the prior distribution.

Notice that choosing the uniform distribution as the prior distribution is not a good idea. The better choice is $\pi_0 \sim N(0, \Sigma)$ with large Σ

First, we initiate the proposal distribution

$$q(\beta' | \beta) = N(\beta, \tau^2 \Sigma)$$

Compute the maximum likelihood estimator $\hat{\beta}$ and $var(\hat{\beta}) = \Sigma$.

Algorithm 4 Bayesian Estimation for Probit Model

```

1: procedure METROPOLIS-HASTINGS()
    $\beta_0 = \hat{\beta}$ 
2: for iteration  $t=1,2,3\dots$  do
3:   sample  $\beta' \sim N(\beta, \tau^2 \Sigma)$  ▷ proposal distribution
4:    $a = \min(1, \frac{\pi(\beta')N(\beta_{t-1}, \beta', \tau^2 \Sigma)}{\pi(\beta_{t-1})N(\beta', \beta_{t-1}, \tau^2 \Sigma)})$  ▷ Acceptance Probability
5:    $\beta_t = \begin{cases} \beta' & \text{with probability } a \\ \beta_{t-1} & \text{with probability } 1 - a \end{cases}$ 

```
