

# Design of Engineering Experiments

## Part 9 – Experiments with Random Factors

- Text reference, Chapter 13, Pg. 484
- Previous chapters have considered **fixed** factors
  - A specific set of factor levels is chosen for the experiment
  - Inference confined to those levels
  - Often **quantitative** factors are fixed (why?)
- When factor levels are chosen at random from a larger population of potential levels, the factor is **random**
  - Inference is about the entire population of levels
  - Industrial applications include measurement system studies

# Random Effects Models

- Example 13-1 (pg. 487) – weaving fabric on looms
- Response variable is strength
- Interest focuses on determining if there is difference in strength due to the different looms
- However, the weave room contains many (100s) looms
- Solution – select a (random) **sample** of the looms, obtain fabric from each
- Consequently, “looms” is a **random factor**
- See data, Table 13-1; looks like standard single-factor experiment with  $a = 4$  &  $n = 4$

# Random Effects Models

- The usual single factor ANOVA model is

$$y_{ij} = \mu + \tau_i + \varepsilon_{ij} \begin{cases} i = 1, 2, \dots, a \\ j = 1, 2, \dots, n \end{cases}$$

- Now both the error term and the treatment effects are random variables:

$$\varepsilon_{ij} \text{ is } NID(0, \sigma^2) \text{ and } \tau_i \text{ is } NID(0, \sigma_\tau^2)$$

- **Variance components:**  $V(y_{ij}) = \sigma^2 + \sigma_\tau^2$

# Relevant Hypotheses in the Random Effects (or Components of Variance) Model

- In the fixed effects model we test equality of treatment means
- This is no longer appropriate because the treatments are randomly selected
  - the individual ones we happen to have are not of specific interest
  - we are interested in the **population** of treatments
- The appropriate hypotheses are

$$H_0 : \sigma_{\tau}^2 = 0$$

$$H_1 : \sigma_{\tau}^2 > 0$$

# Testing Hypotheses - Random Effects Model

- The standard ANOVA partition of the total sum of squares still works; leads to usual ANOVA display
- Form of the hypothesis test depends on the **expected mean squares**

$$E(MS_E) = \sigma^2 \text{ and } E(MS_{Treatments}) = \sigma^2 + n\sigma_\tau^2$$

- Therefore, the appropriate test statistic is

$$F_0 = MS_{Treatments} / MS_E$$

# Estimating the Variance Components

- Use the **ANOVA** method; equate expected mean squares to their observed values:

$$\hat{\sigma}^2 = MS_E \text{ and } \hat{\sigma}^2 + n\hat{\sigma}_\tau^2 = MS_{Treatments}$$

$$\hat{\sigma}_\tau^2 = \frac{MS_{Treatments} - MS_E}{n}$$

$$\hat{\sigma}^2 = MS_E$$

- Potential problems with these estimators
  - Negative estimates (woops!)
  - They are moment estimators & don't have best statistical properties

# Minitab Solution (Balanced ANOVA)

Factor            Type Levels Values

Loom            random            4            1            2            3            4

Analysis of Variance for y

Source	DF	SS	MS	F	P
Loom	3	89.188	29.729	15.68	0.000
Error	12	22.750	1.896		
Total	15	111.938			

Source            Variance Error Expected Mean Square for Each Term  
 component term (using unrestricted model)

1 Loom            6.958    2    (2) + 4(1)

2 Error            1.896            (2)

# Confidence Intervals on the Variance Components

- Easy to find a  $100(1-\alpha)\%$  CI on  $\sigma^2$  :

$$\frac{(N-a)MS_E}{\chi_{\alpha/2, N-a}^2} \leq \sigma^2 \leq \frac{(N-a)MS_E}{\chi_{1-(\alpha/2), N-a}^2}$$

- Other confidence interval results are given in the book
- Sometimes the procedures are not simple



# Extension to Factorial Treatment Structure

- Two factors, factorial experiment, both factors random (Section 13-2, pg. 490)

$$y_{ijk} = \mu + \tau_i + \beta_j + (\tau\beta)_{ij} + \varepsilon_{ijk} \begin{cases} i = 1, 2, \dots, a \\ j = 1, 2, \dots, b \\ k = 1, 2, \dots, n \end{cases}$$

$$V(\tau_i) = \sigma_\tau^2, V(\beta_j) = \sigma_\beta^2, V[(\tau\beta)_{ij}] = \sigma_{\tau\beta}^2, V(\varepsilon_{ijk}) = \sigma^2$$

$$V(y_{ijk}) = \sigma_\tau^2 + \sigma_\beta^2 + \sigma_{\tau\beta}^2 + \sigma^2$$

- The model parameters are *NID* random variables
- Random effects model

# Testing Hypotheses - Random Effects Model

- Once again, the standard ANOVA partition is appropriate
- Relevant hypotheses:

$$H_0 : \sigma_\tau^2 = 0 \quad H_0 : \sigma_\beta^2 = 0 \quad H_0 : \sigma_{\tau\beta}^2 = 0$$

$$H_1 : \sigma_\tau^2 > 0 \quad H_1 : \sigma_\beta^2 > 0 \quad H_1 : \sigma_{\tau\beta}^2 > 0$$

- Form of the test statistics depend on the **expected mean squares**:

$$E(MS_A) = \sigma^2 + n\sigma_{\tau\beta}^2 + bn\sigma_\tau^2 \Rightarrow F_0 = \frac{MS_A}{MS_{AB}}$$

$$E(MS_B) = \sigma^2 + n\sigma_{\tau\beta}^2 + an\sigma_\beta^2 \Rightarrow F_0 = \frac{MS_B}{MS_{AB}}$$

$$E(MS_{AB}) = \sigma^2 + n\sigma_{\tau\beta}^2 \Rightarrow F_0 = \frac{MS_{AB}}{MS_E}$$

$$E(MS_E) = \sigma^2$$

# Estimating the Variance Components – Two Factor Random model

- As before, use the **ANOVA** method; equate expected mean squares to their observed values:

$$\hat{\sigma}_{\tau}^2 = \frac{MS_A - MS_{AB}}{bn}$$

$$\hat{\sigma}_{\beta}^2 = \frac{MS_B - MS_{AB}}{an}$$

$$\hat{\sigma}_{\tau\beta}^2 = \frac{MS_{AB} - MS_E}{n}$$

$$\hat{\sigma}^2 = MS_E$$

- Potential problems with these estimators

## Example 13-2 (pg. 492)

# A Measurement Systems Capability Study

- Gauge capability (or R&R) is of interest
- The gauge is used by an operator to measure a critical dimension on a part
- **Repeatability** is a measure of the variability due only to the gauge
- **Reproducibility** is a measure of the variability due to the operator
- See experimental layout, Table 13-3. This is a two-factor factorial (completely randomized) with both factors (operators, parts) **random – a random effects model**

# Example 13-2 (pg. 493)

## Minitab Solution – Using Balanced ANOVA

Table 13-4 Analysis of Variance (Minitab Balanced ANOVA) for Example 13-2

### Analysis of Variance (Balanced Designs)

Factor	Type	Levels	Values						
part	random	20	1	2	3	4	5	6	7
			8	9	10	11	12	13	14
			15	16	17	18	19	20	
operator	random	3	1	2	3				

### Analysis of Variance for y

Source	DF	SS	MS	F	P
part	19	1185.425	62.391	87.65	0.000
operator	2	2.617	1.308	1.84	0.173
part*operator	38	27.050	0.712	0.72	0.861
Error	60	59.500	0.992		
Total	119	1274.592			

Source	Variance component	Error term	Expected Mean Square for Each Term (using unrestricted model)
1 part	10.2798	3	(4) + 2(3) + 6(1)
2 operator	0.0149	3	(4) + 2(3) + 40(2)
3 part*operator	-0.1399	4	(4) + 2(3)
4 Error	0.9917		(4)

## **Example 13-2 (pg. 493)**

### **Minitab Solution – Balanced ANOVA**

- There is a large effect of parts (not unexpected)
- Small operator effect
- No Part – Operator interaction
- Negative estimate of the Part – Operator interaction variance component
- Fit a reduced model with the Part – Operator interaction deleted

# Example 13-2 (pg. 493)

## Minitab Solution – Reduced Model

Table 13-5 Analysis of Variance for the Reduced Model, Example 13-2

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### Analysis of Variance (Balanced Designs)

Factor	Type	Levels	Values							
part	random	20	1	2	3	4	5	6	7	
			8	9	10	11	12	13	14	
			15	16	17	18	19	20		
operator	random	3	1	2	3					

### Analysis of Variance for y

Source	DF	SS	MS	F	P
part	19	1185.425	62.391	70.64	0.000
operator	2	2.617	1.308	1.48	0.232
Error	98	86.550	0.883		
Total	119	1274.592			

Source	Variance component	Error term	Expected Mean Square for Each Term (using unrestricted model)
1 part	10.2513	3	(3) + 6(1)
2 operator	0.0106	3	(3) + 40(2)
3 Error	0.8832		(3)

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## Example 13-2 (pg. 493)

### Minitab Solution – Reduced Model

- Estimating gauge capability:

$$\begin{aligned}\hat{\sigma}_{gauge}^2 &= \hat{\sigma}^2 + \hat{\sigma}_{\beta}^2 \\ &= 0.88 + 0.01 \\ &= 0.89\end{aligned}$$

- If interaction had been significant?



# The Two-Factor Mixed Model

- Two factors, factorial experiment, factor  $A$  fixed, factor  $B$  random (Section 12-3, pg. 522)

$$y_{ijk} = \mu + \tau_i + \beta_j + (\tau\beta)_{ij} + \varepsilon_{ijk} \begin{cases} i = 1, 2, \dots, a \\ j = 1, 2, \dots, b \\ k = 1, 2, \dots, n \end{cases}$$

$$V(\beta_j) = \sigma_\beta^2, V[(\tau\beta)_{ij}] = [(a-1)/a]\sigma_{\tau\beta}^2, V(\varepsilon_{ijk}) = \sigma^2$$

$$\sum_{i=1}^a \tau_i = 0, \sum_{i=1}^a (\tau\beta)_{ij} = 0$$

- The model parameters  $\beta_j$  and  $\varepsilon_{ijk}$  are *NID* random variables, the interaction effect is normal, but not independent
- This is called the **restricted model**

# Testing Hypotheses - Mixed Model

- Once again, the standard ANOVA partition is appropriate
- Relevant hypotheses:

$$H_0 : \tau_i = 0 \quad H_0 : \sigma_\beta^2 = 0 \quad H_0 : \sigma_{\tau\beta}^2 = 0$$

$$H_1 : \tau_i \neq 0 \quad H_1 : \sigma_\beta^2 > 0 \quad H_1 : \sigma_{\tau\beta}^2 > 0$$

- Test statistics depend on the **expected mean squares**:

$$E(MS_A) = \sigma^2 + n\sigma_{\tau\beta}^2 + \frac{bn \sum_{i=1}^a \tau_i^2}{a-1} \Rightarrow F_0 = \frac{MS_A}{MS_{AB}}$$

$$E(MS_B) = \sigma^2 + an\sigma_\beta^2 \Rightarrow F_0 = \frac{MS_B}{MS_E}$$

$$E(MS_{AB}) = \sigma^2 + n\sigma_{\tau\beta}^2 \Rightarrow F_0 = \frac{MS_{AB}}{MS_E}$$

$$E(MS_E) = \sigma^2$$

# Estimating the Variance Components

## – Two Factor Mixed model

- Use the **ANOVA** method; equate expected mean squares to their observed values:

$$\hat{\sigma}_{\beta}^2 = \frac{MS_B - MS_E}{an}$$

$$\hat{\sigma}_{\tau\beta}^2 = \frac{MS_{AB} - MS_E}{n}$$

$$\hat{\sigma}^2 = MS_E$$

- Estimate the fixed effects (treatment means) as usual

## **Example 13-3 (pg. 497)**

### **The Measurement Systems Capability Study Revisited**

- Same experimental setting as in example 13-2
- Parts are a random factor, but Operators are fixed
- Assume the restricted form of the mixed model
- Minitab can analyze the mixed model

# Example 13-3 (pg. 497)

## Minitab Solution – Balanced ANOVA

Table 13-6 Analysis of Variance (Minitab) for the Mixed Model in Example 13-3.  
The Restricted Model is Assumed.

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### Analysis of Variance (Balanced Designs)

Factor	Type	Levels	Values							
part	random	20	1	2	3	4	5	6	7	
			8	9	10	11	12	13	14	
			15	16	17	18	19	20		
operator	fixed	3	1	2	3					

### Analysis of Variance for y

Source	DF	SS	MS	F	P
part	19	1185.425	62.391	62.92	0.000
operator	2	2.617	1.308	1.84	0.173
part*operator	38	27.050	0.712	0.72	0.861
Error	60	59.500	0.992		
Total	119	1274.592			

Source	Variance component	Error term	Expected Mean Square for Each Term (using restricted model)
1 part	10.2332	4	$(4) + 6(1)$
2 operator		3	$(4) + 2(3) + 40Q[2]$
3 part*operator	-0.1399	4	$(4) + 2(3)$
4 Error	0.9917		$(4)$

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## **Example 13-3**

### **Minitab Solution – Balanced ANOVA**

- There is a large effect of parts (not unexpected)
- Small operator effect
- No Part – Operator interaction
- Negative estimate of the Part – Operator interaction variance component
- Fit a reduced model with the Part – Operator interaction deleted
- This leads to the same solution that we found previously for the two-factor random model

# The Unrestricted Mixed Model

- Two factors, factorial experiment, factor  $A$  fixed, factor  $B$  random (pg. 526)

$$y_{ijk} = \mu + \alpha_i + \gamma_j + (\alpha\gamma)_{ij} + \varepsilon_{ijk} \begin{cases} i = 1, 2, \dots, a \\ j = 1, 2, \dots, b \\ k = 1, 2, \dots, n \end{cases}$$

$$V(\gamma_j) = \sigma_\beta^2, V[(\alpha\gamma)_{ij}] = \sigma_{\alpha\gamma}^2, V(\varepsilon_{ijk}) = \sigma^2$$

$$\sum_{i=1}^a \alpha_i = 0$$

- The random model parameters are now **all** assumed to be *NID*

# Testing Hypotheses – Unrestricted Mixed Model

- The standard ANOVA partition is appropriate
- Relevant hypotheses:

$$H_0 : \alpha_i = 0 \quad H_0 : \sigma_\gamma^2 = 0 \quad H_0 : \sigma_{\alpha\gamma}^2 = 0$$

$$H_1 : \alpha_i \neq 0 \quad H_1 : \sigma_\gamma^2 > 0 \quad H_1 : \sigma_{\alpha\gamma}^2 > 0$$

- **Expected mean squares** determine the test statistics:

$$E(MS_A) = \sigma^2 + n\sigma_{\alpha\gamma}^2 + \frac{bn \sum_{i=1}^a \alpha_i^2}{a-1} \Rightarrow F_0 = \frac{MS_A}{MS_{AB}}$$

$$E(MS_B) = \sigma^2 + n\sigma_{\alpha\gamma}^2 + an\sigma_\gamma^2 \Rightarrow F_0 = \frac{MS_B}{MS_{AB}}$$

$$E(MS_{AB}) = \sigma^2 + n\sigma_{\alpha\gamma}^2 \Rightarrow F_0 = \frac{MS_{AB}}{MS_E}$$

$$E(MS_E) = \sigma^2$$



# Estimating the Variance Components

## – Unrestricted Mixed Model

- Use the **ANOVA** method; equate expected mean squares to their observed values:

$$\hat{\sigma}_{\gamma}^2 = \frac{MS_B - MS_{AB}}{an}$$

$$\hat{\sigma}_{\alpha\gamma}^2 = \frac{MS_{AB} - MS_E}{n}$$

$$\hat{\sigma}^2 = MS_E$$

- The only change compared to the restricted mixed model is in the estimate of the random effect variance component

# Example 13-4 (pg. 499)

## Minitab Solution – Unrestricted Model

Table 13-7 Analysis of the Experiment in Example 13-3 Using the Unrestricted Model

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### Analysis of Variance (Balanced Designs)

Factor	Type	Levels	Values						
part	random	20	1	2	3	4	5	6	7
			8	9	10	11	12	13	14
			15	16	17	18	19	20	
operator	fixed	3	1	2	3				

### Analysis of Variance for y

Source	DF	SS	MS	F	P
part	19	1185.425	62.391	87.65	0.000
operator	2	2.617	1.308	1.84	0.173
part*operator	38	27.050	0.712	0.72	0.861
Error	60	59.500	0.992		
Total	119	1274.592			

Source	Variance component	Error term	Expected Mean Square for Each Term (using unrestricted model)
1 part	10.2798	3	(4) + 2(3) + 6(1)
2 operator		3	(4) + 2(3) + Q[2]
3 part*operator	-0.1399	4	(4) + 2(3)
4 Error	0.9917		(4)

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# Finding Expected Mean Squares

- Obviously important in determining the form of the test statistic
- In fixed models, it's easy:

$$E(MS) = \sigma^2 + f(\text{fixed factor})$$

- Can always use the “brute force” approach – just apply the expectation operator
- Straightforward but tedious
- Rules on page 502-504 work for any balanced model
- Rules are consistent with the **restricted** mixed model – can be modified to incorporate the **unrestricted** model assumptions

# Approximate $F$ Tests

- Sometimes we find that there are no exact tests for certain effects (page 505)
- Leads to an approximate  $F$  test (“pseudo”  $F$  test)
- Test procedure is due to Satterthwaite (1946), and uses **linear combinations** of the original mean squares to form the  $F$ -ratio
- The linear combinations of the original mean squares are sometimes called “synthetic” mean squares
- Adjustments are required to the degrees of freedom
- Refer to Example 13-7, page 507
- Minitab will analyze these experiments, although their “synthetic” mean squares are not always the best choice