

# Inference in graphical models

Graphical model is joint distribution over domain  $V$ .

$$V = U \cup H \cup E$$

observed variables (evidence)  
 hidden variables (unobserved, don't care)  
 query variables (visible variables) (unobserved, interesting)

Inference = answering queries about  $U$ , knowing that  $E = e$ .

Therefore, if  $P_V$  is the graphical model, the distribution

$$P_{U|E} = \frac{\int P_{U,H,E}}{\int P_{U,H,E}}$$

$$\Omega(H)$$

$$\Omega(UU)$$

is central to the inference methods that we study.

Types of queries:  $P_{U|E=e} = ?$

$$P_{x|E=e} = ? \quad x \in U$$

$$\underset{\Omega(U)}{\operatorname{argmax}} P_{U|E=e}$$

$$\underset{\Omega(x)}{\operatorname{argmax}} P_{x|E=e} \quad x \in U$$

" $E = e$ " is called evidence. This type of evidence is called **categorical**. Sometimes evidence is given in the form of a likelihood, i.e.  $\tilde{P}_E$  distribution over  $\Omega(E)$ .

$$\text{categorical evidence} \Leftrightarrow \tilde{P}_E = \begin{cases} 1 & \text{for } E = e \\ 0 & \text{otherwise} \end{cases}$$

Brute force computation of  $P_{U|E}$  (binary variables)

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$$P_{U|E}(u) = \frac{\sum_{h \in \Omega(H)} P_{U|H,E}(u, h, e)}{|\Omega(H)| = 2^{|H|}}$$

$$\sum_{h \in \Omega(H)} \sum_{u \in \Omega(U)} P_{U|H,E}(u, h, e) \quad |\Omega(U)| = 2^{|U|}$$

$\Rightarrow O(2^{|H|+|U|})$  operations

exponential in the number of variables that are marginalized over

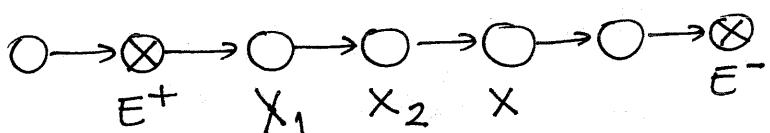
Can we do better by exploiting the structure of graphical models? Yes, if structure is sparse (small no. of parents per variable, small clique size, etc.)

Exploiting structure = exploiting independence

### Pearl's algorithm

- Bayes nets, singly connected
- computes  $P_{X|E}$  for each  $X \in V \setminus E$
- linear time (in  $|V|$ )

### Pearl's algorithm in a chain

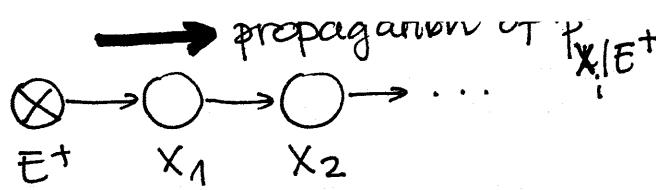


First idea:  $P_{X|E^+E^-} = \frac{P_{X|E^+E^-}}{P_{E^+E^-}} = \frac{P_{E^+}P_{X|E^+}P_{E^-|X,E^+}}{P_{E^+E^-}}$

$$\propto P_{X|E^+}P_{E^-|X}$$

second idea: recursion

Computing  $P_{X|E^+}$



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$P_{X_1|E^+}$  known

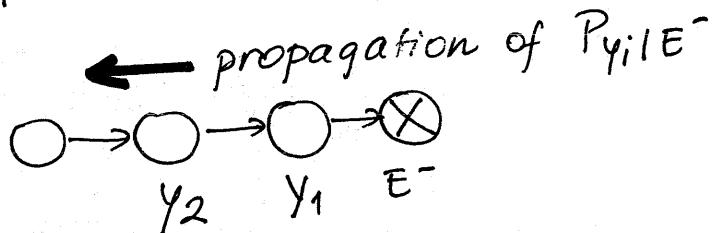
$$P_{X_2|E^+}(x_2|e^+) = \sum_{x_1} P_{X_1|E^+}(x_1|e^+) \cdot P_{X_2|x_1}(x_2|x_1)$$

... and so on, recursively until  $X$

Note that: - computation is recursive, propagating  $P_{X|E^+}$  down the chain

- every stage involves only 2 variables:  $x_i$  and its parent,  $P_{x_i|x_{i-1}}$  and  $P_{X_{i-1}|E^+}$  (the previous "message")  $\Rightarrow$  **LOCALITY**
- total nr. of steps  $\leq |V|=n$
- computations / step  $|\Omega(x_i)| \cdot |\Omega(x_{i-1})|$   
if  $|\Omega(x)| \leq r \Rightarrow$  total computation  $\sim n r^2$

Computing  $P_{E^-|X}$



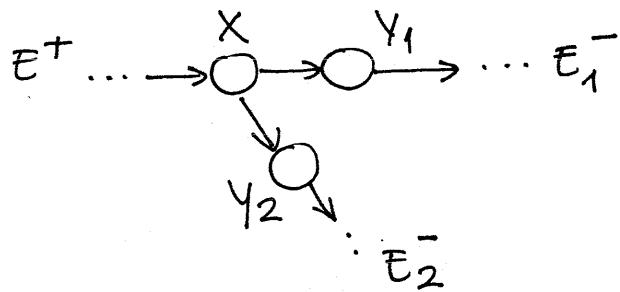
$P_{E^-|Y_1}$  known

$$P_{E^-|Y_2} = \sum_{y_1} P_{E^-|Y_1} \cdot P_{Y_1|Y_2}$$

... and so on

Algorithm

1. Propagate  $P_{X|E^+}$  down the chain
2. Propagate  $P_{E^-|X}$  up the chain
3. Compute  $P_{X|E^+ \cdot E^-}$  by normalizing  $P_{X|E^+} \cdot P_{E^-|X}$

Pearl's algorithm in a tree

up-propagation  $E^-$  is now  $\{E_1^-, E_2^-\}$   
but  $E_1^- \perp E_2^- | X$

$$P_{E_1^- E_2^- | X} = P_{E_1^- | X} \cdot P_{E_2^- | X}$$

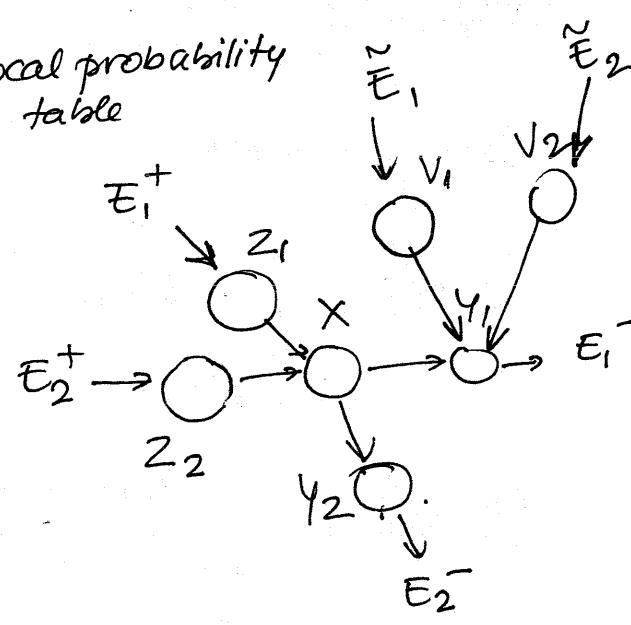
down propagation assume we have  $P_{X | E^+}$

what is  $P_{Y_1 | \text{"evidence above } y_1\text{"}}$ ?

$$P_{Y_1 | E^+, E_2^-} = \sum_x P_{X | E^+, E_2^-} \cdot P_{Y_1 | X}$$

$$\propto \sum_x P_{X | E^+} \cdot P_{E_2^- | X} \cdot P_{Y_1 | X}$$

involves: downward, upward, local probability table  
msg. of parent msgs of siblings

Pearl's algorithm in polytrees

polytree = singly connected network

(see [Pearl88] chapter 4, section 5)

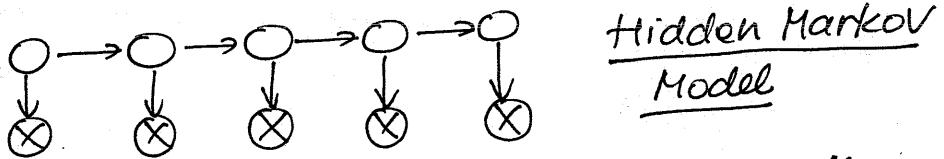
Running time / step no.  $\propto |pa(x)|$

## Conclusion about Pearl's algorithm

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- Computes  $P_{X|E}$  for all  $X \in V \setminus E$
- running time :- linear in  $n = |V|$ 
  - polynomial in  $k = \max|\Omega(x)|$  if  $\max \# \text{parents bounded}$
  - exponential i.e.  $\max \# \text{parents of } X \in V$
- involves only local computations - local variables
  - local prob. tables (parameters)
  - messages from parents, children
- makes use of independence relationships
- works exactly on singly connected Bayes nets (polytrees)

### Special cases



Pearl's alg  $\Leftrightarrow$  Forward - Backward Alg.  
(can do max propagation as well  $\Leftrightarrow$  Viterbi)

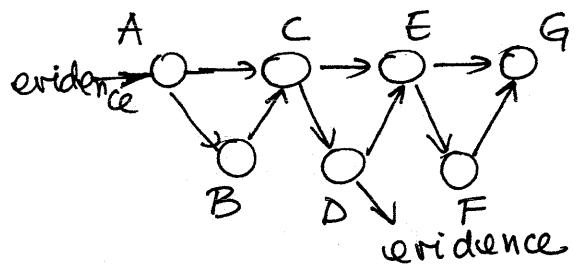
HMM + continuous variables + gaussian distribution + linear dependence  
= Linear system (discrete time)

$\Downarrow$   
Pearl's algorithm  $\Leftrightarrow$  [batch version of] Kalman filter

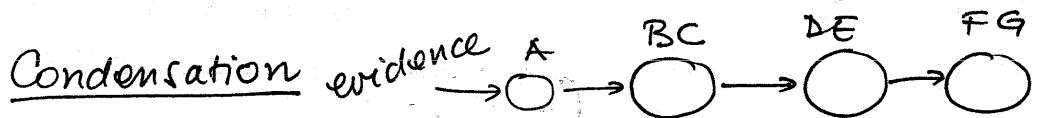
Pearl's alg. itself is a special case of the junction tree alg that will be discussed later on.

# What to do on Bayes nets with loops?

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- Pearl's algorithm - approximation!
- conditioning = breaking the loops
- condense variables
- convert to junction tree -  
- general method

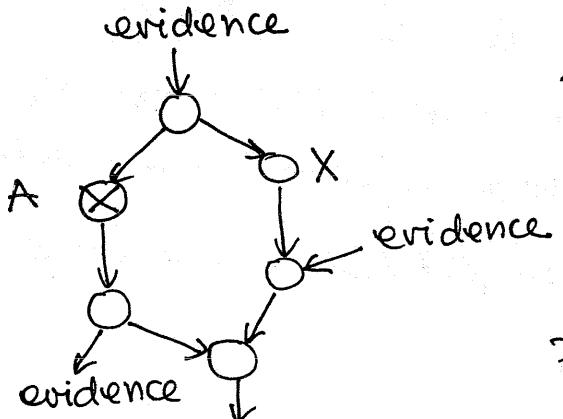


computation cost  $\sim$  exponential in the size of largest <sup>set of</sup> condens. variables

Junction tree alg is a special case of condensation + propagation

## Conditioning

assume A known  $\Rightarrow$  blocks a path and breaks cycle



$$P_{X \mid \text{evidence}} = P_{X \mid A=0, \text{evid}} \cdot P_{(A=0 \mid \text{evidence})} + \\ + P_{X \mid A=1, \text{evid}} \cdot P_{(A=1 \mid \text{evidence})}$$

$$P_{(A=0 \mid \text{evidence})} \propto P(\text{evidence} \mid A=0) \cdot P_A$$

computation cost  $\sim$  exponential in the number of instantiated variables

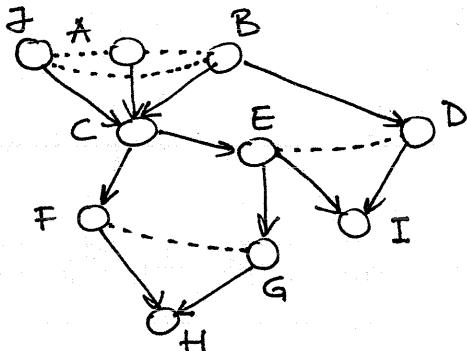
# Converting a Bayes net into a junction tree

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- by adding edges  $\Rightarrow$  hides some independence relationships  
( $\Rightarrow$  they won't be exploited in the inference procedure)

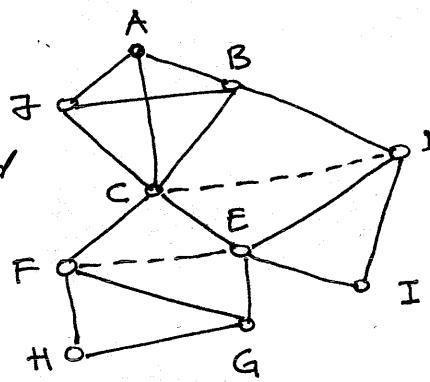
## 1. Moralization

- "marry" parents of common child
- then drop direction of edges



## 2. Triangulation

- add edges until graph triangulated
- triangulation is not unique!



## 3. Junction tree construction

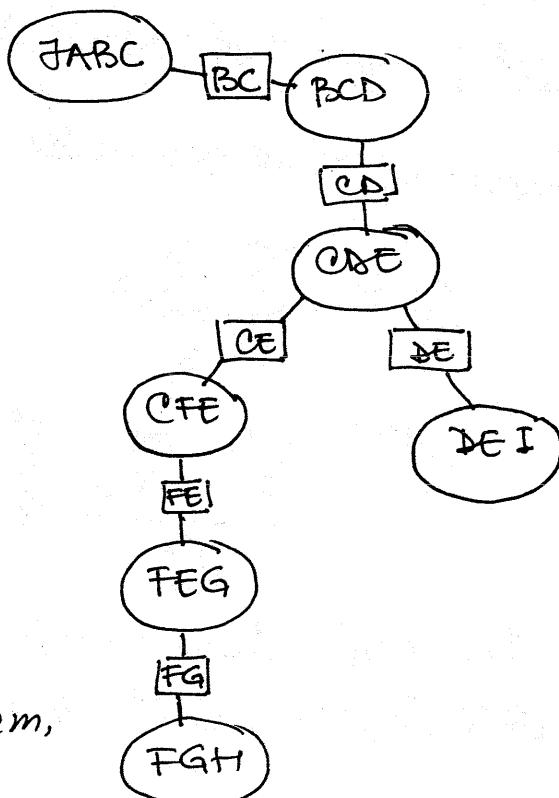
- junction tree may not be unique  
(same cliques, same separators but different topologies)

Total "state space" size =

$$= \sum_{\text{cliques}} |\Omega(c)|$$

indicator of total memory,  
propagation time

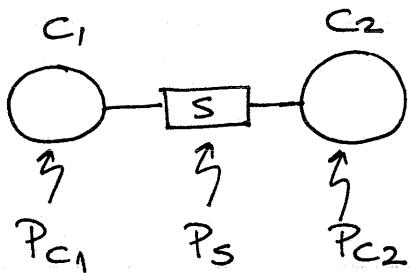
Wanted triangulations that give  
small total state space. Hard problem,  
only heuristics available.



# The Junction tree algorithm

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Passing messages between two cliques



before propagation:

$$P_S = \sum_{\Omega(C_1 \setminus S)} P_{C_1} = \sum_{\Omega(C_2 \setminus S)} P_{C_2}$$

$$\text{new info: } P_{C_1} \leftarrow P_{C_1}^*$$

we say that junction tree is "balanced"

Propagating this new info to  $C_2$  is called absorption.

$$1. P_S^* = \sum_{\Omega(C_1 \setminus S)} P_{C_1}^* \quad \text{new separator potential}$$

$$2. P_{C_2}^* = P_{C_2} \cdot \frac{P_S^*}{P_S} \quad \text{replaces } P_{C_2}; \text{ then } P_S^* \text{ replaces } P_S$$

Why does this operation make sense?

First, note that absorption does not change the joint distribution.

Supposing there are only two cliques, i.e.  $V = C_1 \cup C_2$ :

$$P_V = \frac{P_{C_1}^* \cdot P_{C_2}}{P_S} \quad \text{before absorption}$$

$$P_V^{\text{new}} = \frac{P_{C_1}^* \cdot \left( P_{C_2} \cdot \frac{P_S^*}{P_S} \right)}{P_S^*} = P_V \quad \text{after absorption}$$

Thus, absorption distributes info that is already in the distribution. But how does this info get incorporated?

Second Assume that  $P_V = \frac{P_{C_1} \cdot P_{C_2}}{P_S}$  and

you observe some [subset of] variables [ $\supset$ ]  $A \in C_1 \setminus C_2$ .

Hence  $A=a$ . Before the observation we had

$P_A = \sum_{S(C_1 \setminus A)} P_{C_1}$  = the marginal of  $A$  computed in  $P_{C_1}$

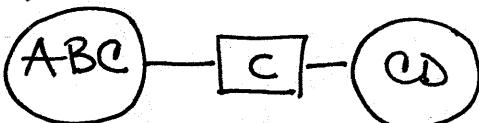
After the observation, we have a new " $P_A$ ":

$$P_A^* = \delta_a = \begin{cases} 1 & A=a \\ 0 & \text{otherwise} \end{cases}$$

We incorporate this into  $C_1$  by the law of conditional probability.  
(let's denote  $C_1 \setminus A = B$ )

$$P_{C_1}^* = P_{C_1 | \text{observation}}(b, a' | A=a) = \begin{cases} \frac{P_{C_1}(b, a)}{P_A} & \text{if } a'=a \\ 0 & \text{if } a' \neq a \end{cases}$$

Example:



$A, B, C, D$  binary variables

		$A=0$	$A=1$
		00	01
$P_{ABC}$	00	0.2	0.1
	01	0.15	0.05
$P_{BC}$	11	0.2	0.2
	10	0.05	0.05

$$\Rightarrow P_C = \begin{bmatrix} 0 & 1 \\ 0.4 & 0.6 \end{bmatrix}$$

		$C=0$	$C=1$
		0	1
$P_{CD}$	0	0.1	0.5
	1	0.3	0.1

Observed  $A=0$

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$$P_A(A=0) = 0.2 + 0.15 + 0.2 + 0.05 = 0.6$$

$$P_{ABC}^* = \frac{P_{ABC} \cdot \delta_{A,0}}{P_A} = \frac{1}{0.6}$$

0.2	0
0.15	0
0.2	0
0.05	0

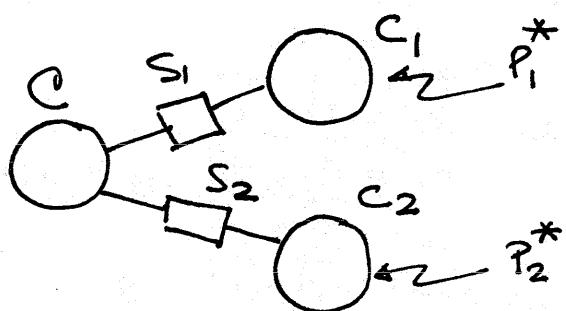
$$P_C^* = \sum_{a,b} P_{ABC}^*(a,b,c) = \begin{bmatrix} \frac{0.25}{0.6} & \frac{0.35}{0.6} \end{bmatrix}$$

$$P_{CD}^* = P_{CD} \cdot \frac{P_C^*}{P_C} = \begin{array}{|c|c|} \hline & 0 & 1 \\ \hline 0 & 0.10 & 0.486 \\ \hline 1 & 0.312 & 0.097 \\ \hline \end{array}$$

$$\text{for exs } P_{CD}(0,1) = 0.3$$

$$P_{CD}^*(0,1) = P_{CD}(0,1) \cdot \frac{P_C^*(0)}{P_C(0)} = 0.3 \cdot \frac{\frac{0.25}{0.6}}{0.4} = 0.312$$

How to absorb from several cliques?



$$1. P_{S_i}^* = \sum_{C_i \setminus S_i} P_{C_i}^* \quad i=1,2,\dots$$

$$2. P_C^* = P_C \cdot \frac{P_{S_1}^*}{P_{S_1}} \cdot \frac{P_{S_2}^*}{P_{S_2}} \dots$$

This is justified by the requirement that  $P_V$  remain constant after absorption.

The general case

Arbitrary junction tree

$$P_V = \frac{\prod P_{C_i}}{\prod P_{S_j}}$$

Arbitrary categorical evidence

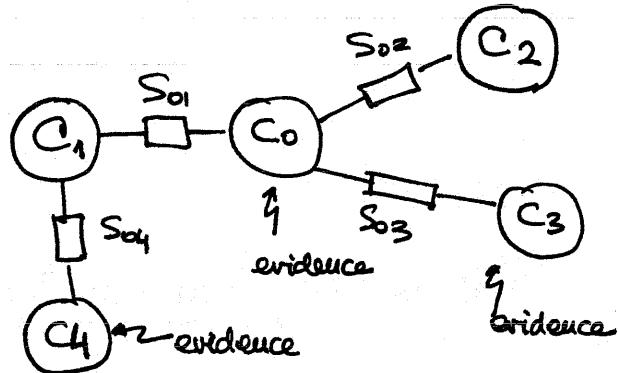
$$x_k = x_k^* \quad k=1, \dots, n$$

1. Incorporate evidence:for each  $k$  find  $C_i$  clique such that  $x_k \in C_i$ 

$$P_{C_i}^* = \frac{P_{C_i} \cdot \delta_{x_k, x_k^*}}{P_{x_k}(x_k^*)}$$

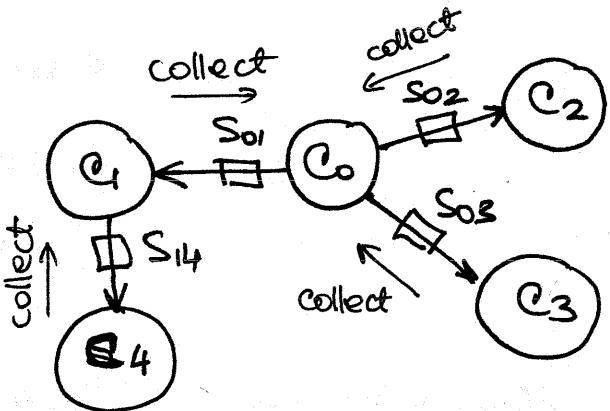
2. "Direct" the junction tree.choose [any clique]  $C_0$  as root

direct all edges away from root

3. Collect evidence

starting from root do recursively:

- call "collect evidence" in all children
- then absorb from all children

(this results in a series of absorptions from leaves towards the root. For ex:  $C_1$  absorbs from  $C_4$ 

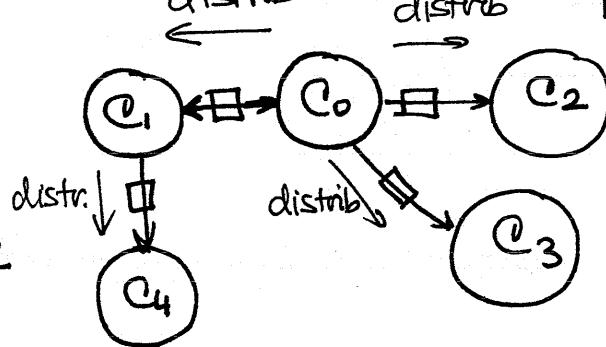
$$C_0 \longrightarrow C_1, C_2, C_3$$

#### 4. "Distribute evidence"

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starting from root do recursively  
for each clique  $C$ :

- each child of  $C$  absorbs from  $C$
- each child calls "distribute evidence"



For ex, here :  $C_1, C_2, C_3$  absorb from  $C_0$   
 $C_4$  absorbs from  $C_1$

END

This is the Junction Tree Algorithm.

It ensures that :

- the tree is balanced
- every  $P_C^*, P_S^*$  is normalized
- every  $P_C^*$  is the marginal of  $C$  in  $P_V^*$
- $P_V^*$  represents  $\text{Prob}[V \mid \text{evidence}]$

Non-categorical evidence :

- info from external source tells us

$$P_{X_k}^* \quad (\neq P_{X_k})$$

AND

$$(\neq S_{X_k, x_k^*})$$

• then  $P_{C_i}^* = \frac{P_{C_i} \cdot P_{X_k}^*}{P_{X_k}}$  in (1.) and the other

steps remain unchanged

# Computational complexity of the junction tree algorithm

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Absorption  $C_1 \rightarrow C_2$  takes  $\sim |\Omega(C_1)|$  additions

$|\Omega(C_2)|$  multiplications

Collect evidence : absorptions children  $\rightarrow$  parent  $\Rightarrow$   
distribute evidence : parent  $\rightarrow$  child

whole alg  $\approx 2 \cdot \sum_{\text{cliques}} |\Omega(C)|$  additions and  
multiplications  
 $\underbrace{2 \cdot \sum_{\text{cliques}}}_{2T}$

$T$  is called "total state space"; it's an approximate measure of the cost of inference in the junction tree.

- Note that :
- $|\Omega(C)|$  is exponential in  $|C|$
  - therefore  $T \ll |\Omega(V)|$  if tree is sparse
    - ↑ cost of "brute force" inference
  - it is important to minimize  $T$  during triangulation. (but this is hard!)
  - a good approximation to  $T$  is  $\max_C |\Omega(C)|$   
(minimizing this is also NP-hard)

The JT algorithm • is exact and general

- any exact, general algorithm for inference has to perform a [possibly hidden] triangulation; hence, it is essentially no better than the JT algorithm
- unfortunately, for many practical problems the JT is intractable : image segmentation, max likelihood decoding, medical diagnosis in QMR-AT

# Approximate inference methods

- Pearl's (polytree) algorithm ← for Bayes nets
  - comes with no guarantees
  - seems to work in practice
- Markov Chain Monte Carlo methods (MCMC)
  - i.e. Sampling← for Markov fields mainly
- variational methods ← for Bayes nets, Markov fields approximating  $P_{V/evidence}$

# Graphical Probability Models (aka Belief Nets)

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Pearl (188):

"probability theory is unique in its ability to process context-sensitive beliefs, and what makes the process computationally feasible is that the information needed for specifying context dependencies is represented by graphs and manipulated by local propagation!"

Graphical model = graphical representation of (conditional) independence relationships in a joint distribution  
= the distribution itself

graphical model

- structure (a graph)
- parametrization (depends on graph,  
parameters are "local")

graph =  $(V, E)$

- ↓
- nodes
- variables
- ↓
- edges
- "dependencies"

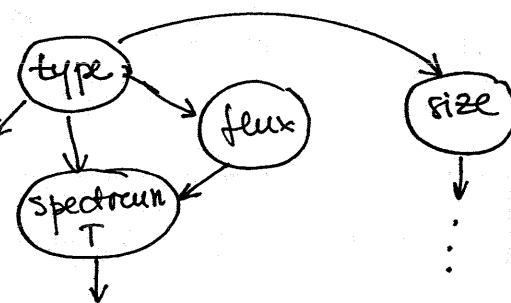
More precisely: a missing edge represents an independence relationship.

But what is "Independence"?

Example: "Redshift" model

$$V = \{ \text{type}, \text{flux}, \text{size}, \dots \}$$

$$E = \{ \text{type} \rightarrow \text{flux}, \text{type} \rightarrow \text{size}, \dots \}$$



# Probabilistic independence

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$$A \perp B \iff P_{AB} = P_A \cdot P_B$$

"A independent B"

(marginal formulation)

Example 1. A, B binary variables,  $A \perp B$

$$P_{AB}:$$

		B/A		$P_B$
		0	1	
$P_A$	0	0.12	0.28	0.4
	1	0.18	0.42	0.6
		0.3	0.7	

Example 2. A, B real, jointly gaussian

$$[A \ B] \sim N([μ_A \ μ_B], Σ_{AB})$$

$$P_{AB}(a, b) = \frac{1}{|\Sigma|^{1/2} \cdot 2\pi} \exp \left\{ -\frac{1}{2} \begin{pmatrix} a - \mu_A & b - \mu_B \end{pmatrix} \Sigma^{-1} \begin{pmatrix} a - \mu_A \\ b - \mu_B \end{pmatrix} \right\}$$

$$A \perp B \iff \Sigma_{AB} = \begin{bmatrix} \tau_A^2 & 0 \\ 0 & \tau_B^2 \end{bmatrix} \quad (\text{diagonal})$$

Ex 1. Prove this!

$$\begin{array}{ccc} \text{Ex 2. } A \perp B & \xrightarrow{?} & A \perp B, C \\ A \perp C & \xleftarrow{?} & \end{array}$$

Is any of the implications  $\iff$  true?  
Prove or give counterexample!

$$A \perp B \iff P_{A|B} = P_A \iff$$

knowing B brings no additional information about A

( $\uparrow$   
conditional formulation)

## Conditional independence

$$A \perp B | C \Leftrightarrow P_{ABC} = P_{A|C} \cdot P_{B|C} \Leftrightarrow P_{A|BC} = P_{A|C}$$

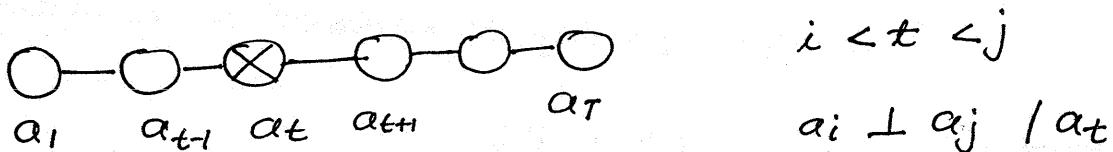
"once  $C$  is known, knowing  $B$  brings no additional information about  $A$ "

→ has to hold for all values of  $A, B, C$

## Representing independences in graphs

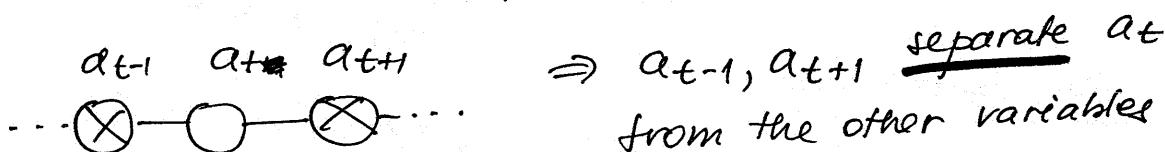
$$\begin{array}{ccc} \text{Independence} & \longleftrightarrow & \text{Separation} \\ (\text{in joint distribution}) & & (\text{in graph}) \end{array}$$

### Markov chain

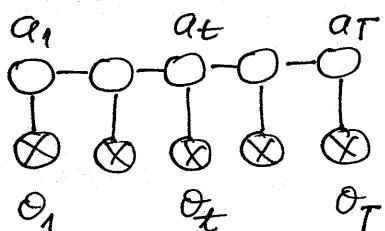


Knowing a state makes future independent from past.

Knowing  $a_{t+1}, a_{t-1} \Rightarrow$  no other state can give additional info about  $a_t$



### Variant Hidden Markov Model



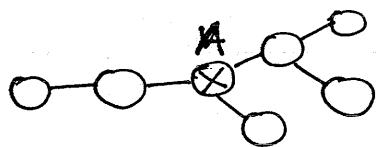
separation properties allow for efficient algorithms for HMMs (forward-backward, Viterbi, Kalman filter, Baum-Welch)

There are  $T$  "hidden" variables.  $\Rightarrow 2^T$  hidden states

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But, forward-backw, etc, ... are  $O(T)$  (i.e. linear!)

Tree



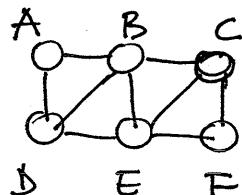
graph with no cycles  
(or singly connected)

Property - every variable  $A$  breaks the tree into 2 or more subtrees

- between two variables there is exactly 1 path

Trees are the only structure for which learning can be done efficient

Markov field



Separation

•  $V_1 \perp V_2 \mid V_3 \Leftrightarrow V_3$  "blocks" all the paths from  $V_1$  to  $V_2$

• Variable  $\perp$  everything | neighbors

(i.e. all info about a variable  $X$  that can be obtained from  $V \setminus X$  is contained in the neighbors of  $X$ )

Examples:  $A, D \perp C, F \mid B, E$

$B \perp F \mid C, E$

$B \not\perp F \mid D, E$

$A \perp C, E, F \mid \begin{matrix} B \\ \text{neighbors of } A \end{matrix}$

Parametrization

Clique = fully connected subgraph of a graph

Ex: • all edges are size 2 cliques

•  $ABD, BDE, BCE, CEF$  are size 3 cliques. They are maximal.

•  $BCDE$  is not a clique

Denote by  $C_G$  the set of cliques of a graph  $G$ .  
Maximal

$P_V$  Markov Field  
over graph  $G$

$\Leftrightarrow$

$$P_V(x) = \prod_{c \in C_G} \psi_c(x_c)$$

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$$\text{Ex: } P_{ABCDEF} = \underbrace{\psi_{ABD}(abd)}_{\downarrow} \psi_{BDE}(bde) \psi_{BCE}(bce) \psi_{CEF}(cef)$$

"local" function/parametrization - depends only  
on the variables ABD

How much memory is saved?

- assume  $A, B, \dots F$  binary variables

- assume  $\psi$  are [probability] tables

$P_{ABCDEF}$  = prob. table with  $2^6$  entries,  $2^6 - 1$  free parameters

$\psi_{ABD}$  = prob. table with  $2^3$  entries,  $2^3 - 1$   $\rightarrow$

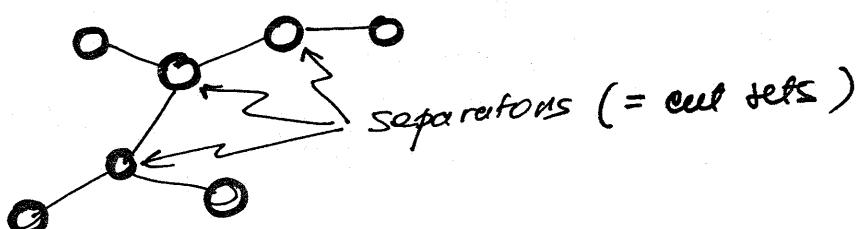
saving:  $2^6 - 1 - 4(2^3 - 1) = 35$  out of 63 (more than 50%)

Ex: Tree over  $n$  variables (binary)

$$T \equiv P_V = \prod_{ABEE} \psi_{AB} \quad \text{Memory: } 3(n-1)$$

For a tree,  $\psi_{AB}$  can be related to the corresponding marginals  $P_A$

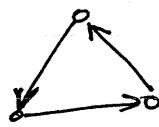
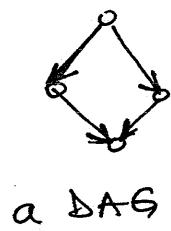
$$T = \frac{\prod_{ABEE} P_{AB}(a, b)}{\prod_{AEV} P_A(a)} = \frac{\prod \text{Pclique}}{\prod \text{Pseparators}}$$



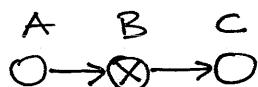
# Bayes Nets

Oct 6, 1998  
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DAG (Directed Acyclic Graph) = directed graph with no directed cycles

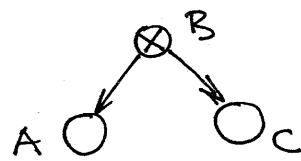


## Separation in a DAG ( $d$ -separation)

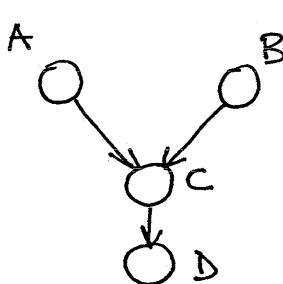


$A \perp C \mid B$

(path  $A \rightarrow B \rightarrow C$  blocked)



$A \perp C \mid B$ , but  $A \not\perp C$



$A \perp B$  but  $A \not\perp B \mid C$

$A \not\perp B \mid D$

(path  $A \rightarrow B$  blocked by  $C$  or  $D$ )

In general  $A \perp B \mid C$  in a DAG iff all paths btw. A and B are blocked when C is instantiated.

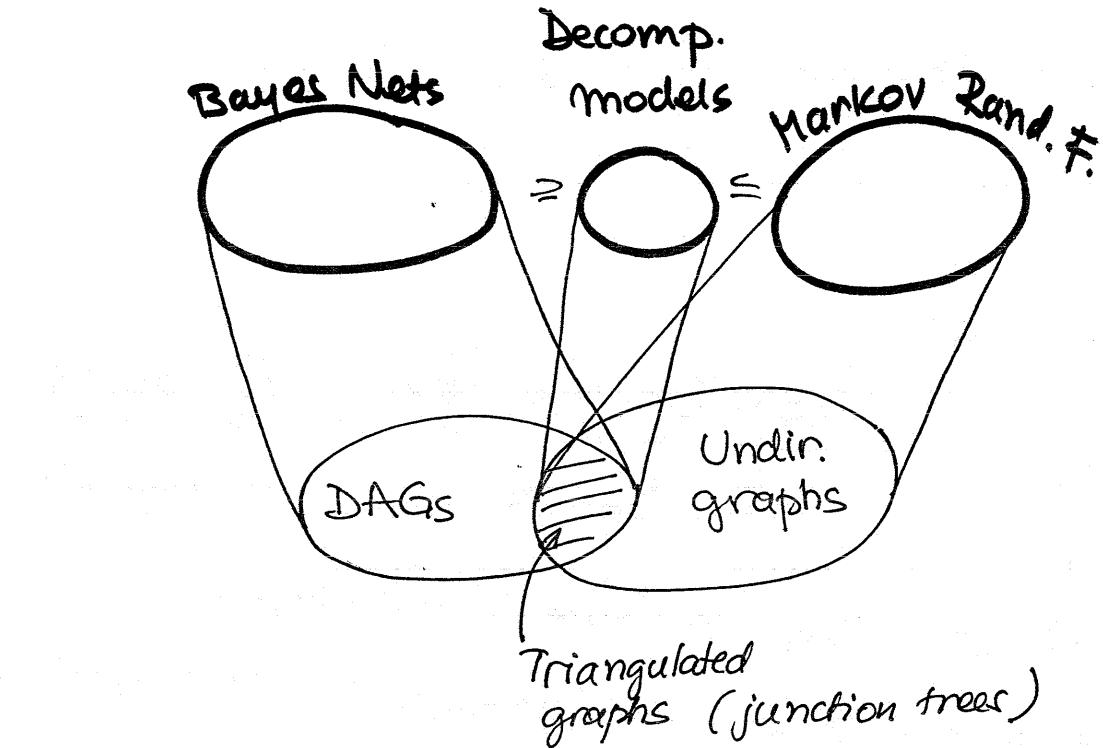
Understanding separation is important for inference: if two variables are separated we can ignore one when reasoning about the other.

## Parametrization

$$P_V = \prod_{A \in V} \underbrace{P}_{\text{local fctns}} \underbrace{\alpha_{\text{pa}(A)}}_{\text{(depend only on } A, \text{ pa}(A))}$$

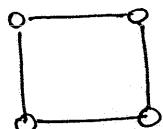
$\alpha(A)$  = parents of variable A

local fctns (depend only on  $A, \text{pa}(A)$ )

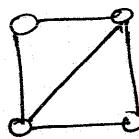


## Triangulated graphs / decomposable models

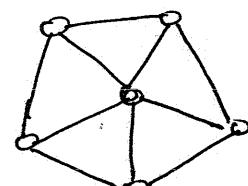
Graph  $G$  triangulated (or chordal)  $\Leftrightarrow$  every cycle  $\geq 4$  in  $G$  has a chord



not triangulated

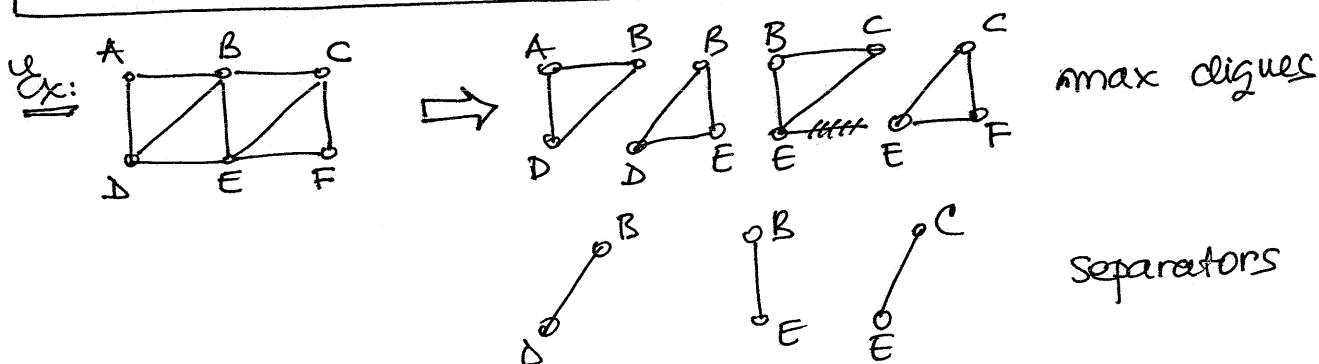


triangulated



NOT triangulated

Triangulated graphs can be represented as a tree of max cliques.



## Junction tree



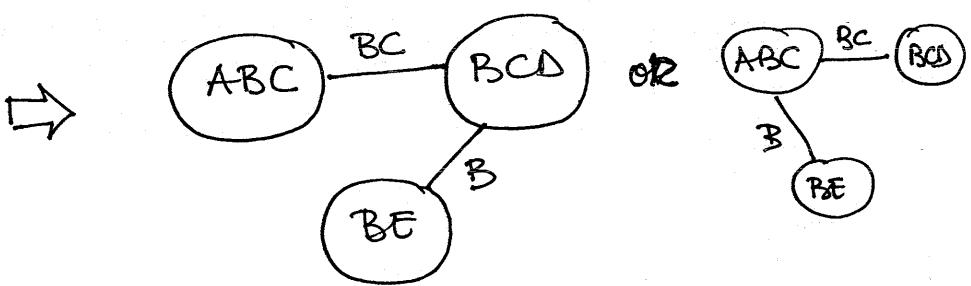
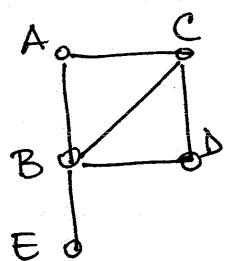
⇒ parametrization:

$$P_{ABCDEF} = \frac{P_{ABD} \cdot P_{BDE} \cdot P_{BCE} \cdot P_{CEF}}{P_{BD} \cdot P_{BE} \cdot P_{CE}} = \frac{\prod_c P_{\text{cliques}}}{\prod_s P_{\text{separators}}}$$

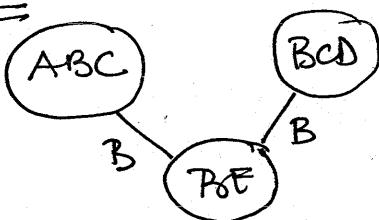
Remarks: • Junction tree structure is not unique

- Junction tree is a max spanning tree w.r.t. separator size

Ex:



but NOT:



# The Modified Pearl Algorithm

(After Peot & Schachter)

At each node  $X$ :

$U = \{U_1, \dots, U_m\}$  parents

$V = \{V_1, \dots, V_n\}$  children

$P_{X|U}$  conditional prob.  
table

$\lambda_X(x)$  = evidence for  $X$

$$= \begin{cases} \delta_x & \text{if } X=x \text{ observed} \\ 1 & \text{if } X \text{ unobserved} \end{cases}$$

local data tables :  $\bar{\pi}_x(x), \lambda(x)$   
 $P(X, e^+), P(e^- | X=x)$

Algorithm (Modified Pearl)

Initialize  $\bar{\pi}_x(x) = \bar{\pi}_{X|U_j}(x) = P(X=x)$  .... obtained by a propagation without evidence

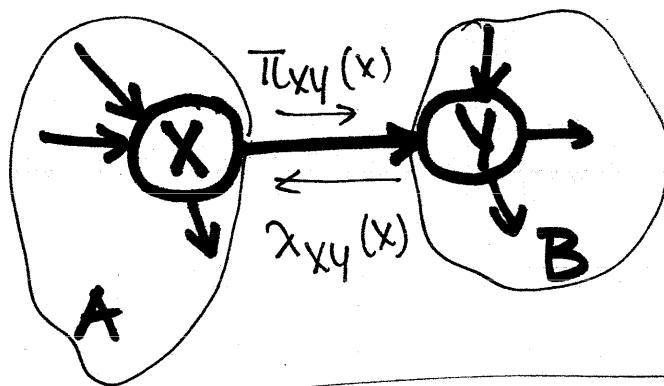
$$\lambda = 1$$

Update at node  $X$  :

$$P(X) \propto \bar{\pi}_x(x) \lambda(x)$$

$$\bar{\pi}_x(x) = \sum_{u \in \Omega_{U \setminus x}} P(X=x | U=u) \prod_{k=1}^m \bar{\pi}_{U_k X}(u_k)$$

$$\lambda(x) = \lambda_X(x) \prod_{i=1}^m \lambda_{X|U_i}(x)$$



$$\bar{\pi}_{xy}(x) = P(X=x, e_A)$$

$$\lambda_{xy}(x) = P(e_B | X=x)$$

"MESSAGES"

# [Modified Pearl Algorithm]

STAT 593A  
Spring 05

To parent  $U_i$

$$\lambda_{U_i X}(u_i) = \sum_x \lambda(x) \sum_{u' \in \Omega \setminus U_i} P(x | u', u_i) \prod_{k \neq i} \pi_{U_k X}(u_k)$$

[2]

To child  $Y_j$

$$\pi_{X Y_j}(x) = \pi(x) \lambda_x(x) \prod_{k \neq j} \lambda_{X Y_k}(x)$$

# Sum-Product Algorithm on Trees

(After Wainwright & Jordan)

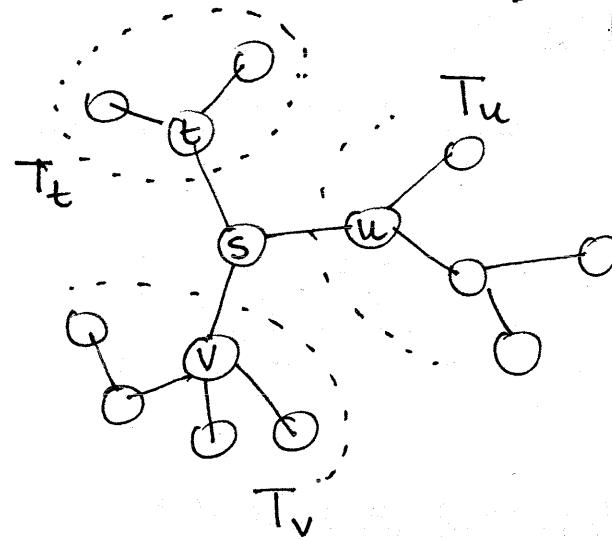
Tree distribution

$$P_X = \frac{\prod_{st \in E} P_{st}(x_s, x_t)}{\prod_{s \in V} P_s(x_s)^{deg(s)-1}}$$

"potentials"  $\psi_{st} = P_{st}$   
 $\psi_s = 1/P_s^{deg(s)-1}$

$$\Rightarrow P_X(x) = \prod_{s \in V} \psi_s(x_s) \cdot \prod_{st \in E} \psi_{st}(x_s, x_t)$$

enter evidence by multiplication



Wanted:  $\mu_s(x_s) = \sum_{x' : x'_s = x_s} P_X(x')$  Marginal of node  $s$

Remarks: Node  $s$  separates tree into  $s$ , and disjoint trees  $T_u, T_v, T_t$    
 $n(s)$

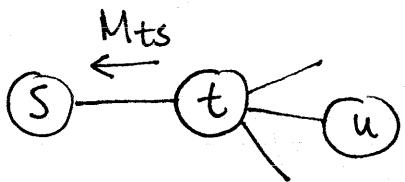
Denote  $T_u = (V_u, E_u)$

$$\begin{aligned} \mu_s(x_s) &= \sum_{x' : x'_s = x_s} \psi_s(x_s) \cancel{\prod_{u \in n(s)} P_{X_{Vu}}(x'_{vu})} \\ &= \psi_s(x_s) \cdot \prod_{u \in n(s)} \left[ \sum_{x'_{vu}} P_{X_{Vu}}(x'_{vu}) \cdot \psi_{su}(x_s, x'_u) \right] \\ &\quad \text{--- "message" from } T_u \\ &= M_{us}^*(x_s) \end{aligned}$$

Algorithm

Message updating:  $M_{ts}(x_s) \leftarrow \sum_{x'_t} [\psi_{st}(x_s, x'_t) \psi_t(x'_t)]$

$$\prod_{u \in N(t) \setminus s} M_{ut}(x'_t)$$



Message propagation:

- Each node  $s$  gets messages from all its neighbours, and sends messages to all neighbors (asynchronous, naive)
- OR:
  - Choose a root  $s$
  - Collect messages recursively from leaf nodes to  $s$
  - Distribute messages recursively from  $s$  toward leaf nodes

**Sum-Product:** in  $\textcircled{*}$ , each ~~message~~ message is computed as a product of messages received, followed by a sum.

**Max-Product:** Replace " $\sum$ " by "max" in  $\textcircled{*}$   $\Rightarrow$  dynamic programming algorithm to compute the most likely configuration