

# **Spatial statistics**

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# Course description

Lectures (10)

**Practica (8)**

Need to submit solutions to the problems given in at least four of the practica. Can be done in groups of 2-3.

**Homework**

Need to submit eight homework problems. To be done individually. A data analysis can, with permission, replace three problems.

**Office hours**

PG Tu 12-1

PDS Th 11-12

**PASI**

# Outline

- 1. Kriging (9/28)**
- 2. Spatial covariance (10/3)**
- 3. Nonstationary structures I: deformations (10/10)**
- 4. Nonstationary structures II: linear combinations etc. (10/17)**
- 5. Space-time models (10/24)**
- 6. Markov random fields (10/31)**
- 7. Misalignment and use of deterministic models (11/7)**
- 8. Design of monitoring network (11/14)**
- 9. Extremes (11/16)**
- 10. Statistical climatology (11/28)**



# Kriging

# The geostatistical model

Gaussian process  $Z(\mathbf{s}), \mathbf{s} \in D \subseteq \mathbb{R}^2$

$$\mu(\mathbf{s}) = \mathbb{E}Z(\mathbf{s}) \quad \text{Var } Z(\mathbf{s}) < \infty$$

$Z$  is **strictly stationary** if

$$(Z(\mathbf{s}_1), \dots, Z(\mathbf{s}_k)) \stackrel{d}{=} (Z(\mathbf{s}_1 + \mathbf{h}), \dots, Z(\mathbf{s}_k + \mathbf{h}))$$

$Z$  is **weakly stationary** if

$$\mu(\mathbf{s}) \equiv \mu \quad \text{Cov}(Z(\mathbf{s}_1), Z(\mathbf{s}_2)) = \mathbf{C}(\mathbf{s}_1 - \mathbf{s}_2)$$

$Z$  is **isotropic** if weakly stationary and

$$\mathbf{C}(\mathbf{s}_1 - \mathbf{s}_2) = \mathbf{C}_0(\|\mathbf{s}_1 - \mathbf{s}_2\|)$$

# The problem

**Given observations at n locations**

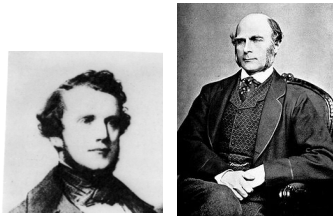
$$Z(s_1), \dots, Z(s_n)$$

**estimate**

$$Z(s_0) \text{ (the process at an unobserved location)}$$

$$\text{or } \int_A Z(s) dv(s) \text{ (an average of the process)}$$

**In the environmental context often time series of observations at the locations.**



## Some history

**Regression (Bravais, Galton, Bartlett)**

**Mining engineers (Krige 1951, Matheron, 60s)**

**Spatial models (Whittle, 1954)**

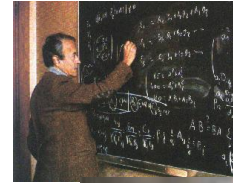
**Forestry (Matérn, 1960)**

**Objective analysis (Gandin, 1961)**

**More recent work:**

**Stein (1999)**

**Gelfand et al. (2010)**



## A Gaussian formula

$$\text{If } \begin{pmatrix} \mathbf{X} \\ \mathbf{Y} \end{pmatrix} \sim \mathbf{N} \left( \begin{pmatrix} \boldsymbol{\mu}_X \\ \boldsymbol{\mu}_Y \end{pmatrix}, \begin{pmatrix} \boldsymbol{\Sigma}_{XX} & \boldsymbol{\Sigma}_{XY} \\ \boldsymbol{\Sigma}_{YX} & \boldsymbol{\Sigma}_{YY} \end{pmatrix} \right)$$

$$\text{then } (\mathbf{Y} | \mathbf{X}) \sim \mathbf{N}(\boldsymbol{\mu}_Y + \boldsymbol{\Sigma}_{YX} \boldsymbol{\Sigma}_{XX}^{-1} (\mathbf{X} - \boldsymbol{\mu}_X), \\ \boldsymbol{\Sigma}_{YY} - \boldsymbol{\Sigma}_{YX} \boldsymbol{\Sigma}_{XX}^{-1} \boldsymbol{\Sigma}_{XY})$$



## Simple kriging

Let  $X = (Z(s_1), \dots, Z(s_n))^T$ ,  $Y = Z(s_0)$ , so that

$$\begin{aligned}\mu_X &= \mu \mathbf{1}_n, \quad \mu_Y = \mu, \\ \Sigma_{XX} &= [C(s_i - s_j)], \quad \Sigma_{YY} = C(0), \text{ and} \\ \Sigma_{YX} &= [C(s_i - s_0)].\end{aligned}$$

Then

$$p(X) \equiv \hat{Z}(s_0) = \mu + [C(s_i - s_0)]^T [C(s_i - s_j)]^{-1} (X - \mu \mathbf{1}_n)$$

This is the best unbiased linear predictor when  $\mu$  and  $C$  are known (simple kriging).

The prediction variance is

$$m_1 = C(0) - [C(s_i - s_0)]^T [C(s_i - s_j)]^{-1} [C(s_i - s_0)]$$

## Some variants

**Ordinary kriging (unknown  $\mu$ )**

$$p(X) \equiv \hat{Z}(s_0) = \hat{\mu} + [C(s_i - s_0)]^T [C(s_i - s_j)]^{-1} (X - \hat{\mu} \mathbf{1}_n)$$

where

$$\hat{\mu} = \left( \mathbf{1}_n^T [C(s_i - s_j)]^{-1} \mathbf{1}_n \right)^{-1} \mathbf{1}_n^T [C(s_i - s_j)]^{-1} X$$

**Universal kriging ( $\mu(s) = A(s)\beta$  for some spatial variable A)**

$$\hat{\beta} = \left( [A(s_i)]^T [C(s_i - s_j)]^{-1} [A(s_i)] \right)^{-1} [A(s_i)]^T [C(s_i - s_j)]^{-1} X$$

$$p(X) \equiv \hat{Z}(s_0) = \hat{\mu}(s_0) + [C(s_i - s_0)]^T [C(s_i - s_j)]^{-1} (X - [\hat{\mu}(s_i)])$$

where  $\hat{\mu}(s) = A(s)\beta$

**Still optimal for known C.**

## Universal kriging variance

$$E\left(\hat{Z}(s_0) - Z(s_0)\right)^2 = \mathbf{m}_1 +$$

simple kriging variance

$$\begin{aligned} & \left( \mathbf{A}(s_0) - [\mathbf{A}(s_i)]^T [\mathbf{C}(s_i - s_j)]^{-1} [\mathbf{C}(s_i - s_0)] \right)^T \\ & \times \left( [\mathbf{A}(s_i)]^T [\mathbf{C}(s_i - s_j)]^{-1} [\mathbf{A}(s_i)] \right)^{-1} \\ & \times \left( \mathbf{A}(s_0) - [\mathbf{A}(s_i)]^T [\mathbf{C}(s_i - s_j)]^{-1} [\mathbf{C}(s_i - s_0)] \right) \end{aligned}$$

variability due to estimating  $\mu$

# The (semi)variogram

$$\gamma(\|h\|) = \frac{1}{2} \text{Var}(Z(s+h) - Z(s)) = C(0) - C(\|h\|)$$

**Intrinsic stationarity**

**Weaker assumption (C(0) needs not exist)**

**Kriging predictions can be expressed in terms of the variogram instead of the covariance.**

## The exponential variogram

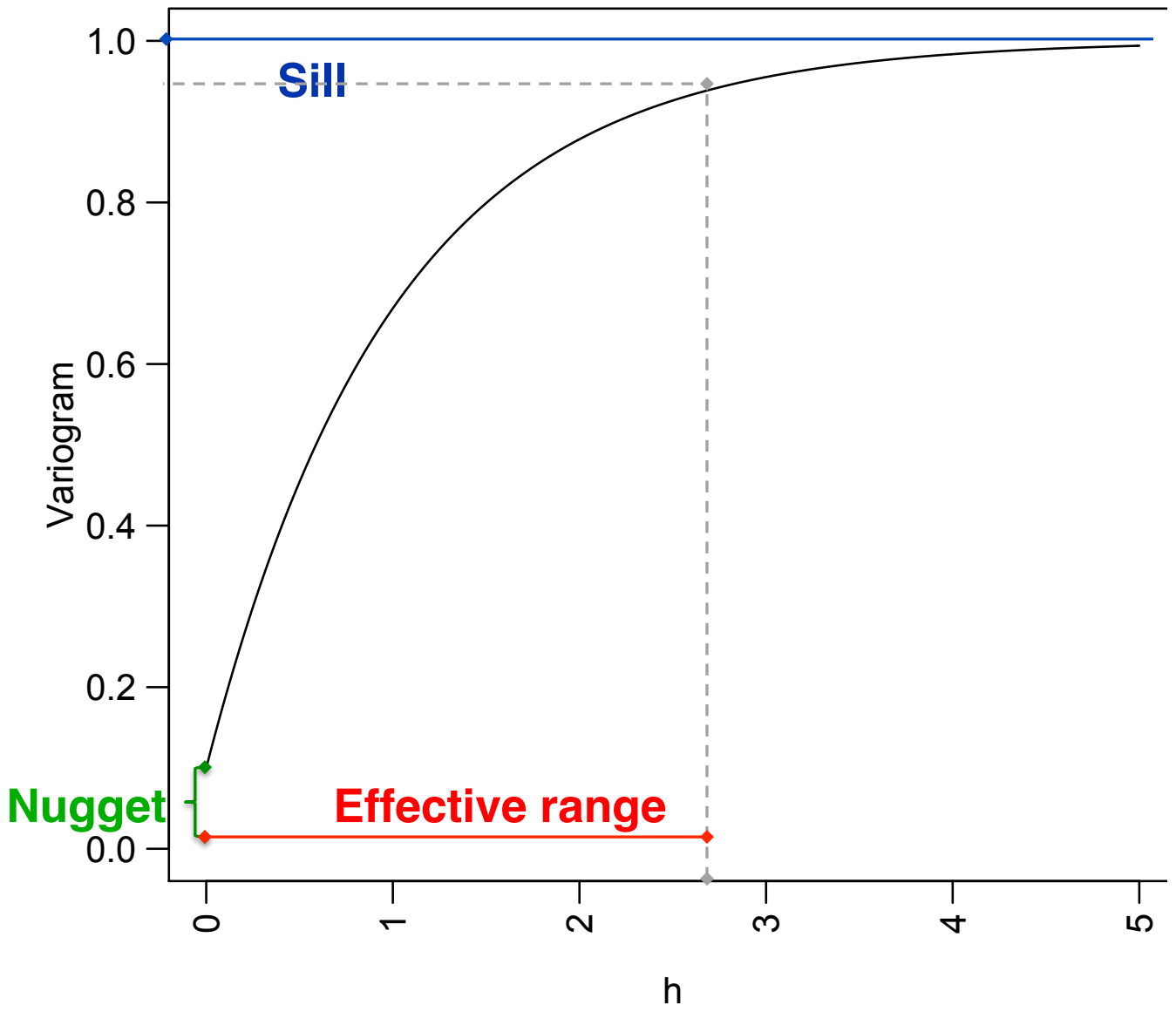
A commonly used variogram function is  $\gamma(h) = \sigma^2 (1 - e^{-h/\phi})$ .

Corresponds to a Gaussian process with continuous but not differentiable sample paths.

More generally,

$$\gamma(h) = (\sigma^2 - \tau^2)(1 - e^{-h/\phi}) + \tau^2$$

has a nugget  $\tau^2$ , corresponding to measurement error and spatial correlation at small distances.



## Ordinary kriging

$$\hat{\mathbf{Z}}(\mathbf{s}_0) = \sum_{i=1}^n \lambda_i \mathbf{Z}(\mathbf{s}_i)$$

where

$$\boldsymbol{\lambda}^T = \left( \boldsymbol{\gamma} + \mathbf{1} \frac{\mathbf{1}^T \boldsymbol{\Gamma}^{-1} \boldsymbol{\gamma}}{\mathbf{1}^T \boldsymbol{\Gamma}^{-1} \mathbf{1}} \right)^T \boldsymbol{\Gamma}^{-1}$$

$$\boldsymbol{\gamma} = (\gamma(\mathbf{s}_0 - \mathbf{s}_1), \dots, \gamma(\mathbf{s}_0 - \mathbf{s}_n))^T$$

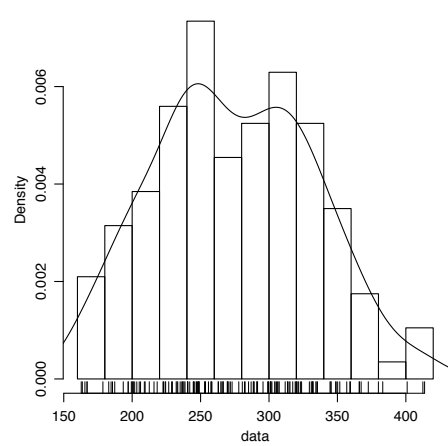
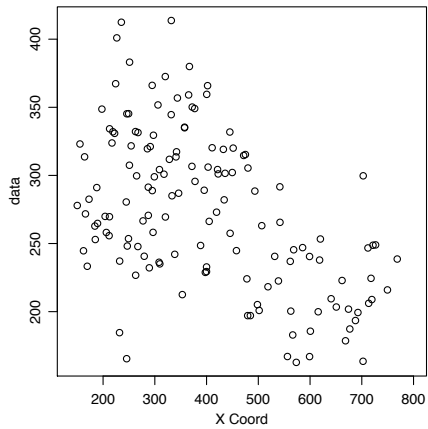
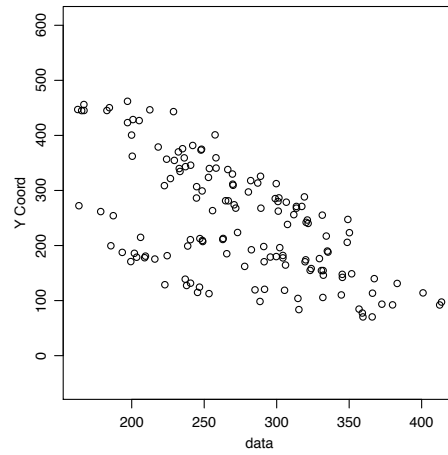
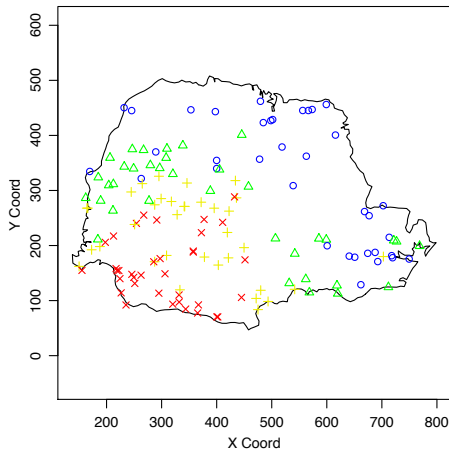
$$\Gamma_{ij} = \gamma(\mathbf{s}_i - \mathbf{s}_j)$$

and kriging variance

$$\mathbf{m}_1(\mathbf{s}_0) = 2 \sum_{i=1}^n \lambda_i \gamma(\mathbf{s}_0 - \mathbf{s}_i) - \sum_{i=1}^n \sum_{j=1}^n \lambda_i \lambda_j \gamma(\mathbf{s}_i - \mathbf{s}_j)$$

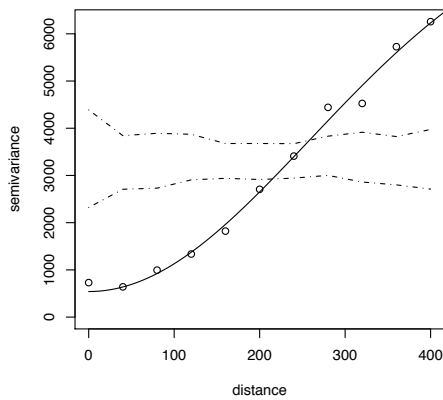
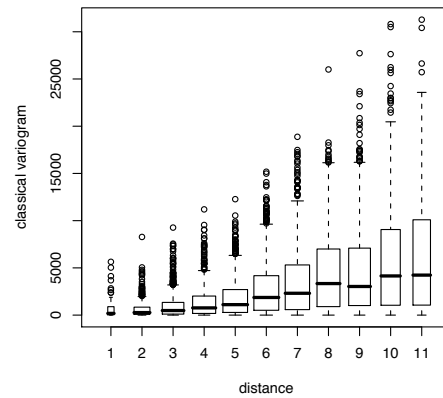
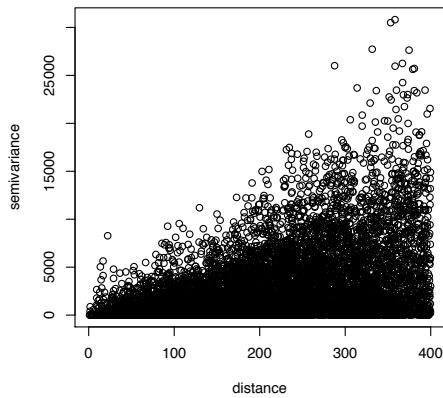
# An example

## Precipitation data from Parana state in Brazil (May-June, averaged over years)



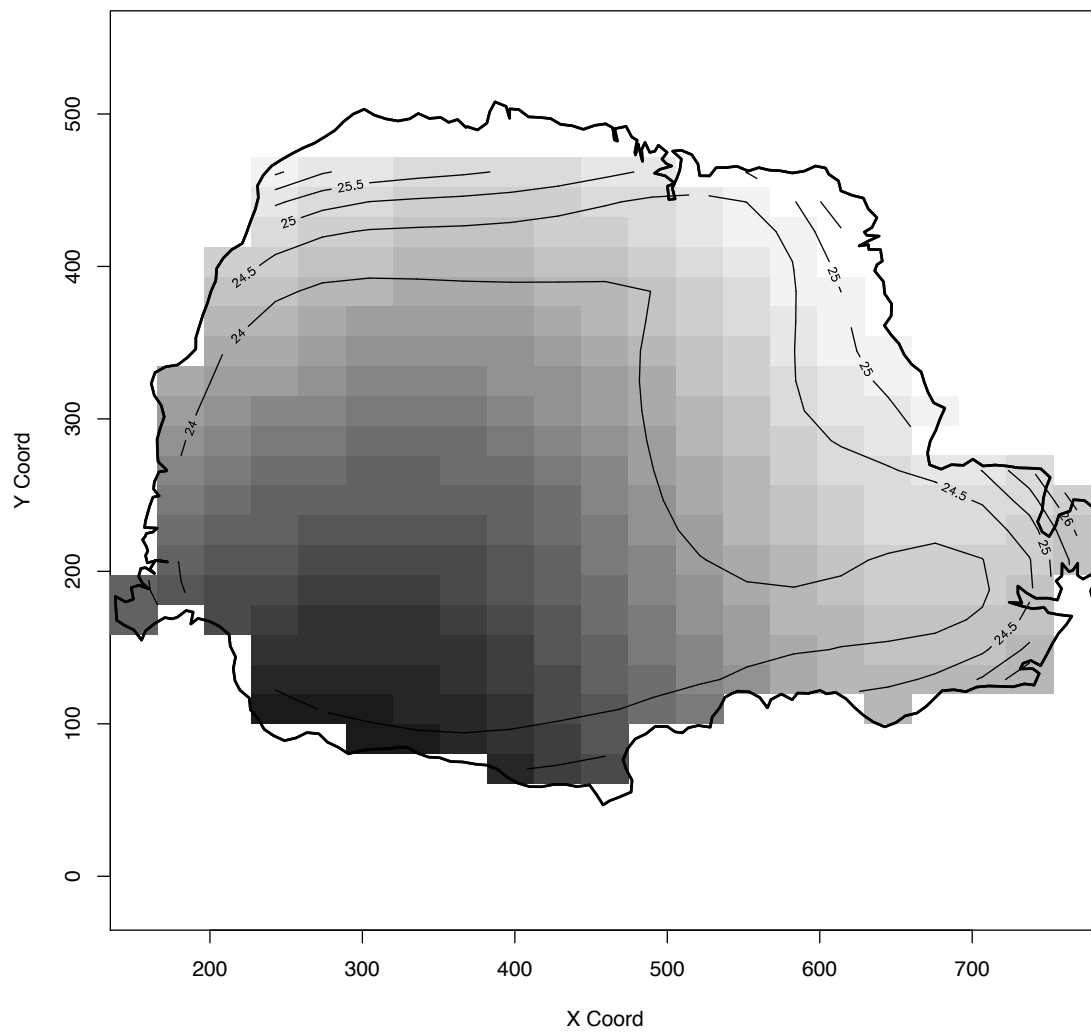


# Variogram plots



$$\gamma(h) = 542 + 8141 \left( 1 - \exp\left(-\left(\frac{h}{365}\right)^2\right) \right)$$

# Kriging surface



# Ozone data set

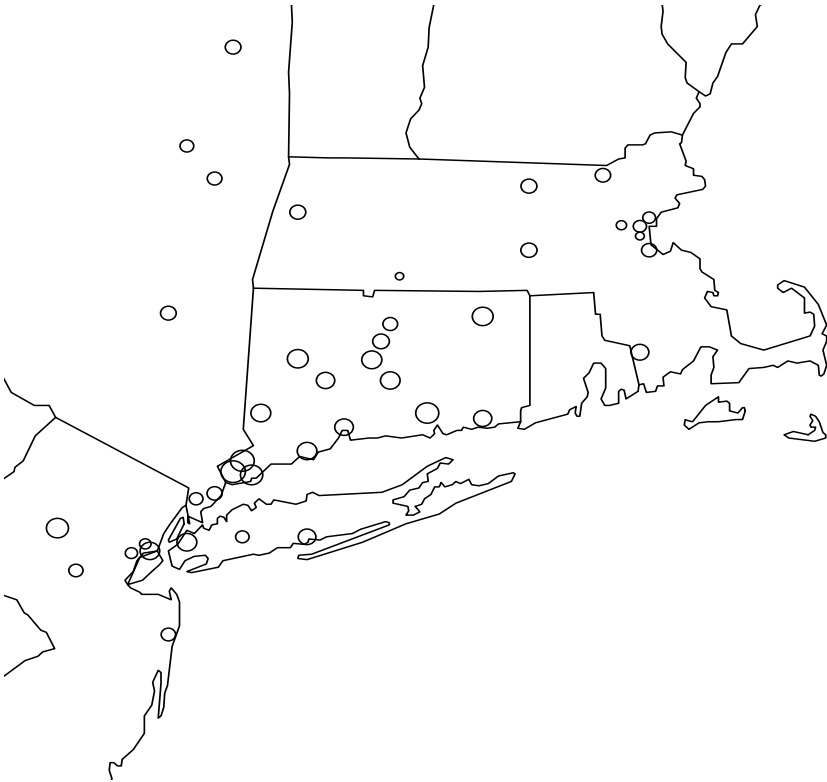
**Built in data set in “maps” library in R  
NW US ozone data**

**1974 June-August median daily  
maximum ground level ozone data from  
41 stations in New Jersey, New York,  
Connecticut and Massachusetts**

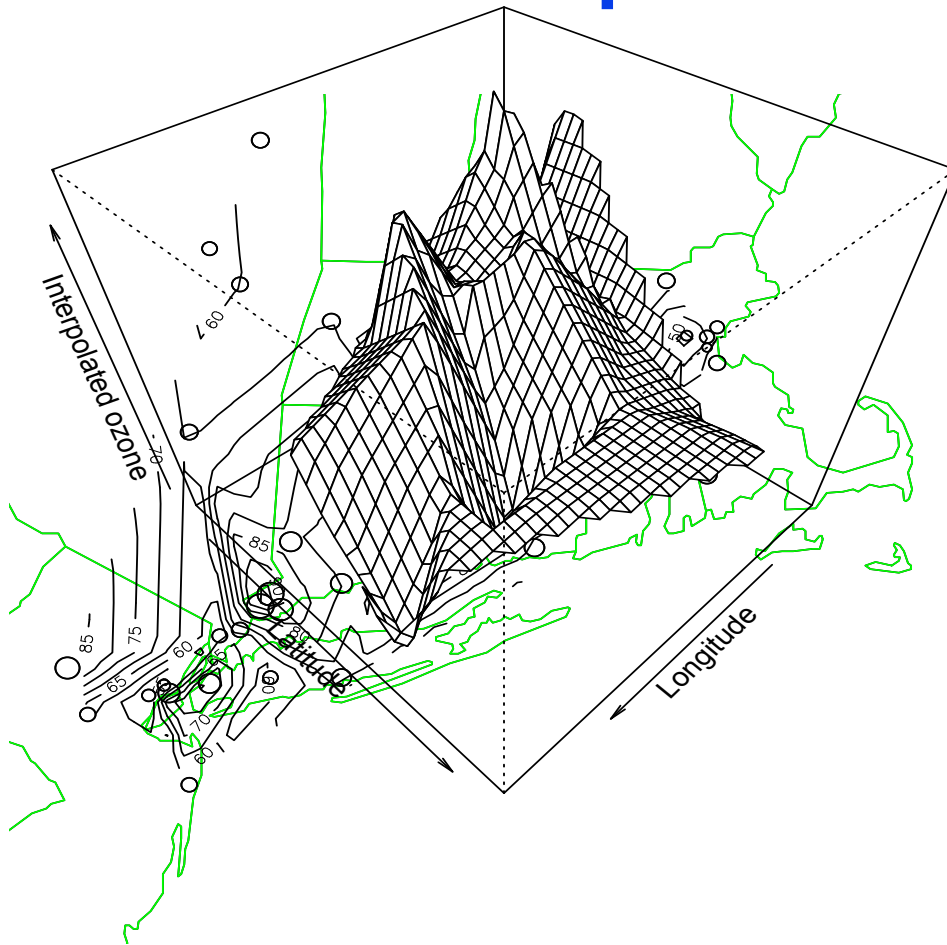
**Contour plot using bilinear  
interpolation**

**Kriging with exponential covariance  
function and nugget**

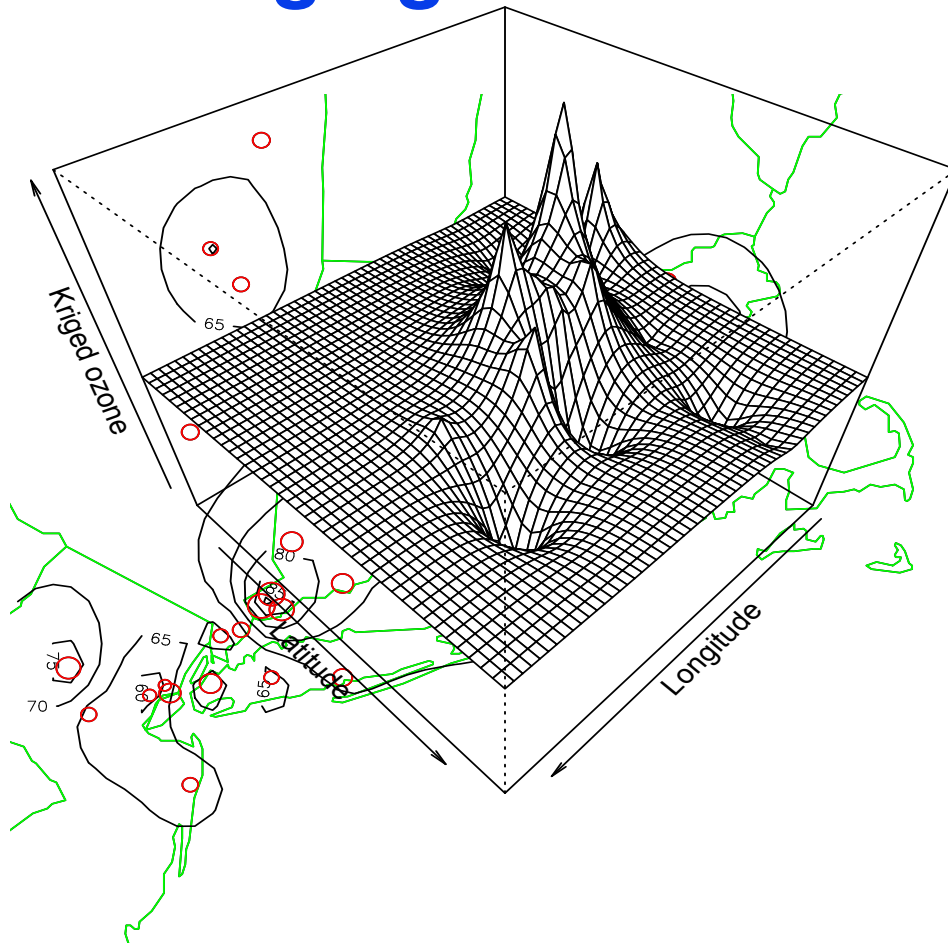
# Data



# Bilinear interpolation



# Kriging the ozone



## How many contours?

A Gaussian prediction  $\hat{Z}(s)$  falls between contour lines between a and b with probability

$$\begin{aligned} p(s) &= \Pr(a < \hat{Z}(s) < b) \\ &= \Phi((1-r)q) - \Phi(-rq) \end{aligned}$$

where  $q=(b-a)/\sigma$  and  $r= (\hat{Z}(s) - a)/(b - a)$

If  $q=.5$  the probability is at most 0.2 that a statement about the level of  $Z(s)$  is correct (Polfeldt, 1999).

If  $q=2$  it is at most 2/3

If  $q=4$  it is at most 0.95.

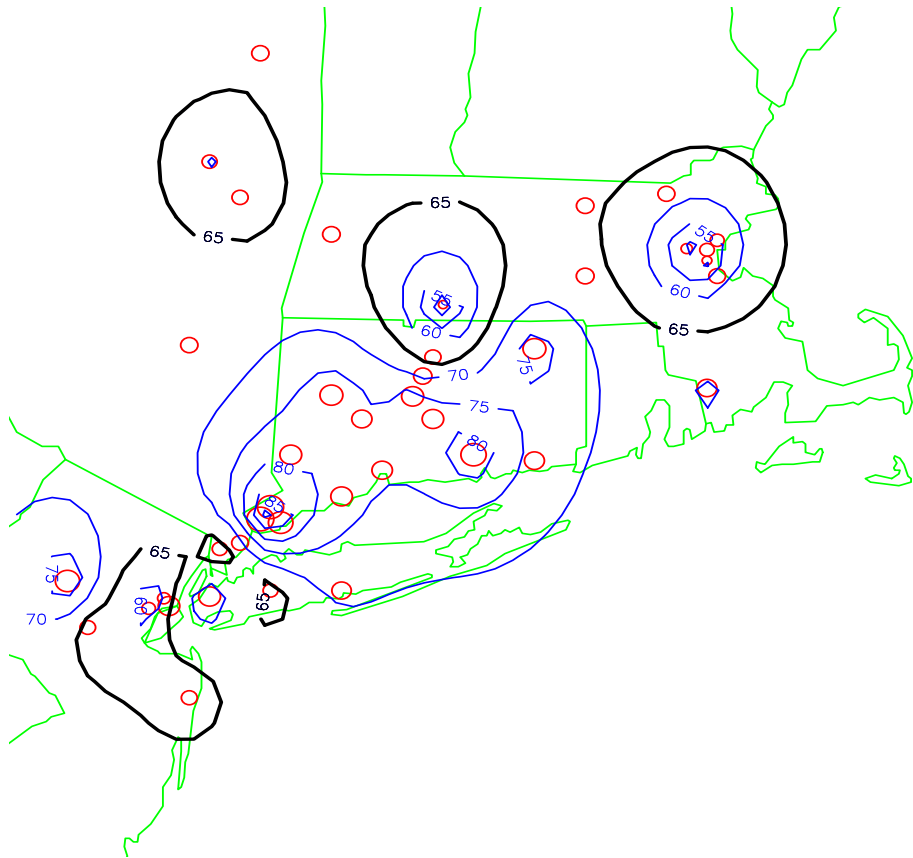
# Consequences

**Points close to contour lines are always very uncertain as to whether they should be above or below the line.**

**If the contour lines are well separated there are high probabilities of correctness in the middle between them.**



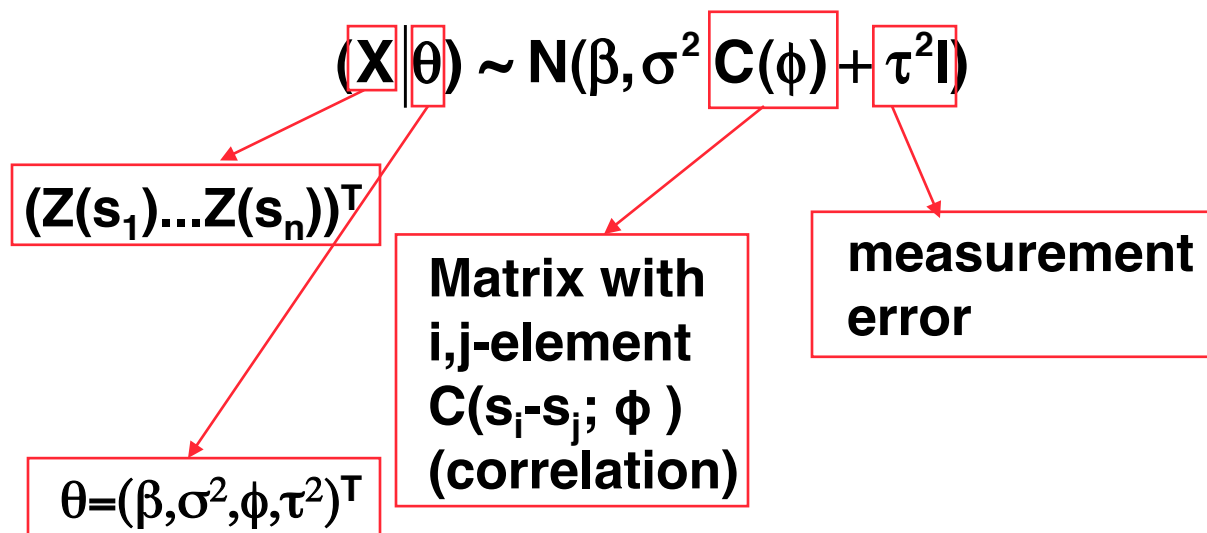
# Revisit kriging contours



# Bayesian kriging

Instead of estimating the parameters, we put a prior distribution on them, and update the distribution using the data.

Model:



## Prior and posterior

**Prior:**  $f(\theta) = f(\beta)f(\sigma^2)f(\phi)f(\tau^2)$

**Posterior:**  $f(\theta|X = x) \propto f(\theta)f(x|\theta)$

$$f(\phi|X = x) \propto f(\phi) \iiint f(x|\theta)f(\beta)f(\sigma^2)f(\tau^2) d\beta d\sigma^2 d\tau^2$$

**Predictive distribution:**

$$f(Z(s_0) | X = x) = \int \mathbf{p}(x; \theta) f(\theta | x) d\theta$$

kriging predictor

## Specifically

**Exponential isotropic correlation function:**

$$\mathbf{C}(\mathbf{h}; \phi) = \exp(-\phi h)$$

**Default correlation model in geoR.**

**Prior on  $\beta$  defaults to flat, but can also be normal or fixed**

**Prior on  $\sigma^2$  defaults to reciprocal, but can be scaled inverse chisquare or flat**

**Prior on  $\phi$  can be exponential, uniform, reciprocal, squared reciprocal, or user specified (discrete)**

## More priors

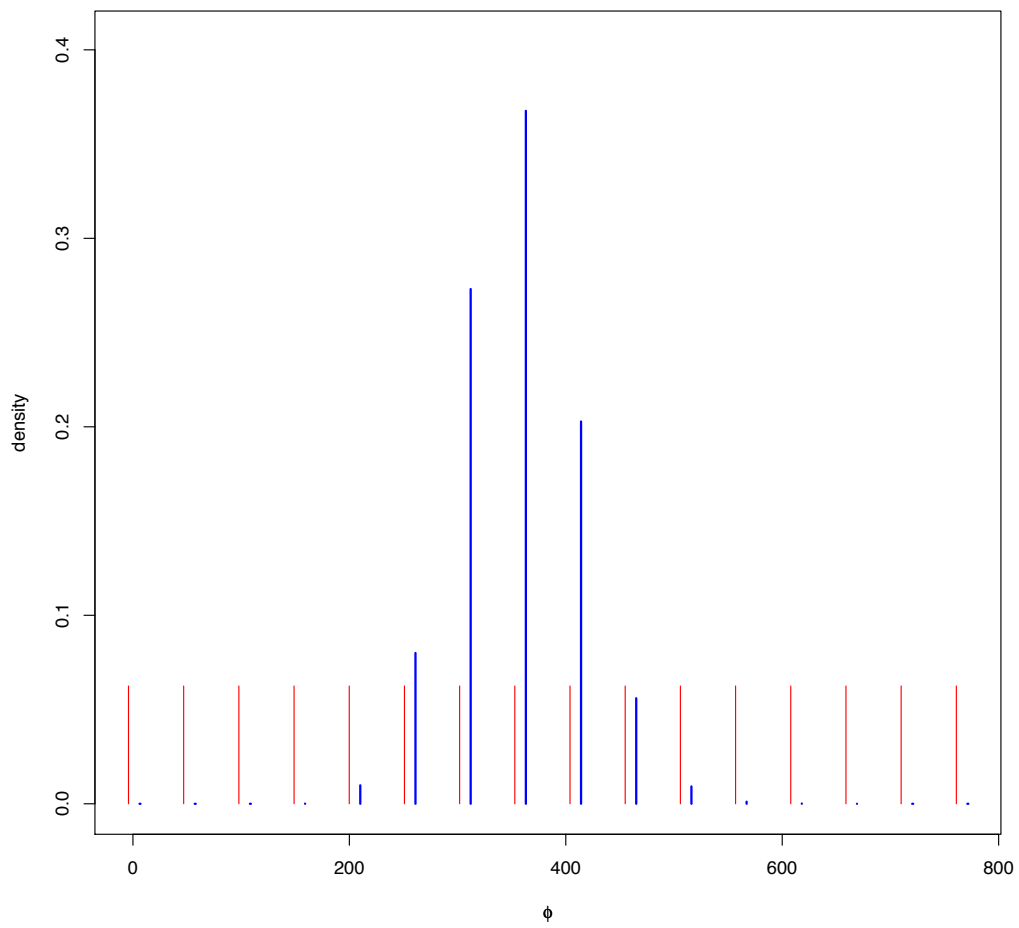
**A prior is assigned to  $\tau/\sigma$ . Defaults to fixed=0, but can also be uniform or user specified (discrete).**

**These choices are made for computational reasons. For example, the posterior distribution of  $\sigma^2$  is inverse chisquared.**

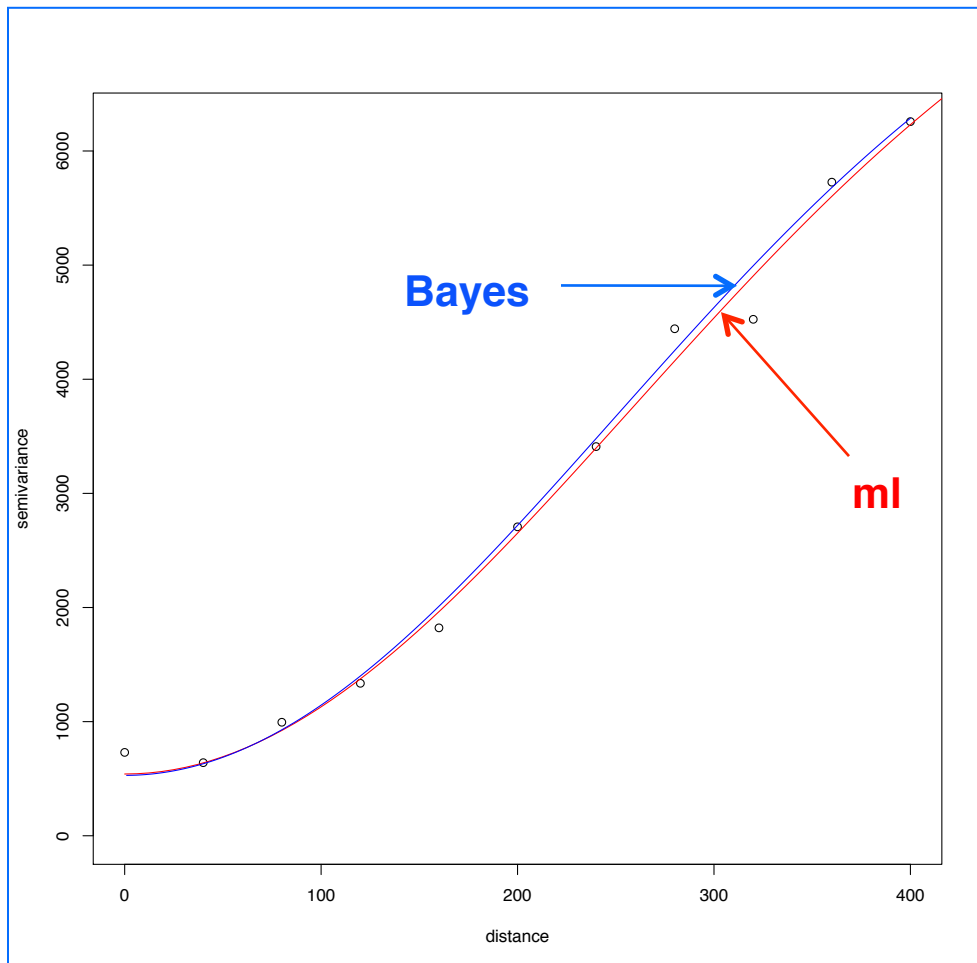
**For details, see**

**<http://www.leg.ufpr.br/geoR/geoRdoc/bayeskrige.pdf>**

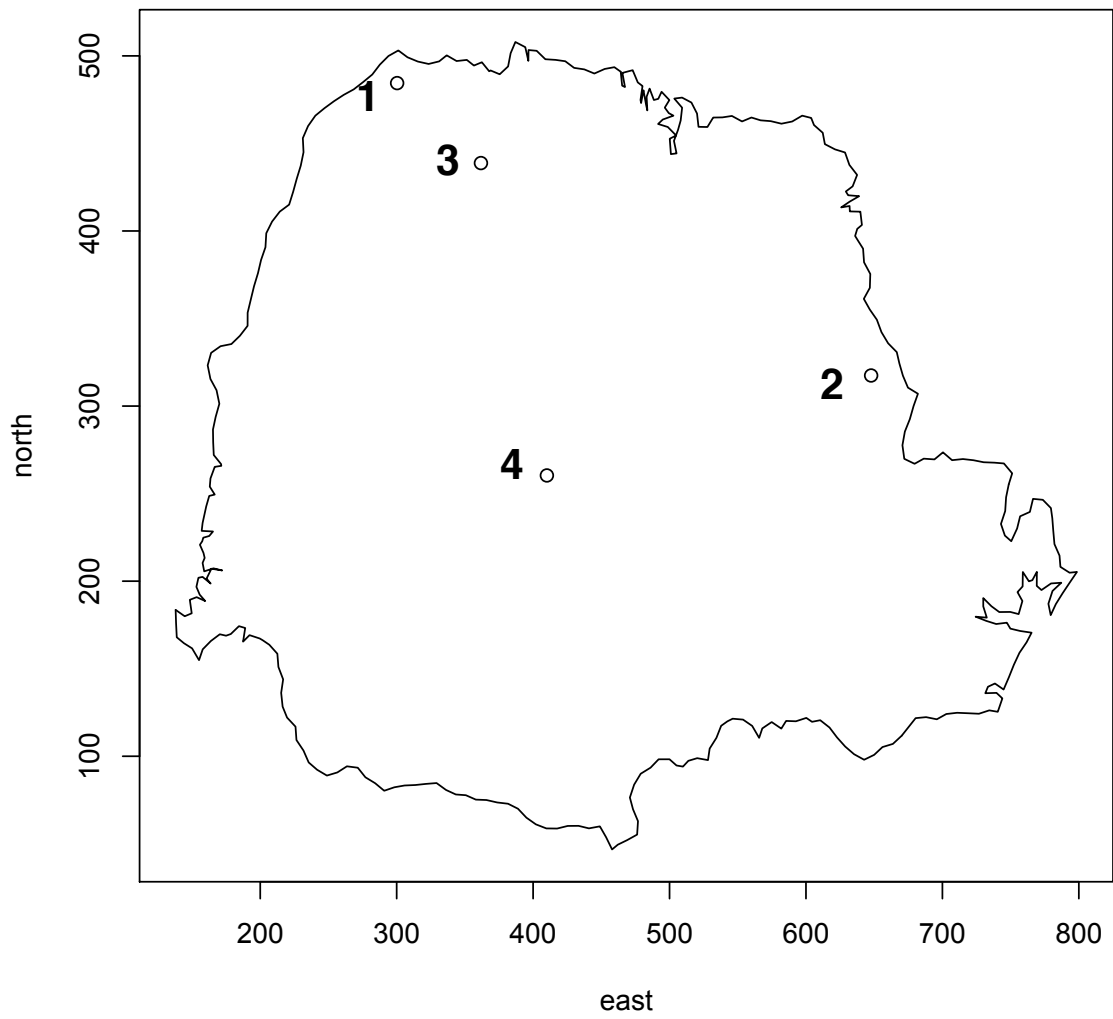
# Prior/posterior of $\phi$



# Estimated variogram



# Prediction sites





# Predictive distribution

