

# Stat547L: Spatial Statistics

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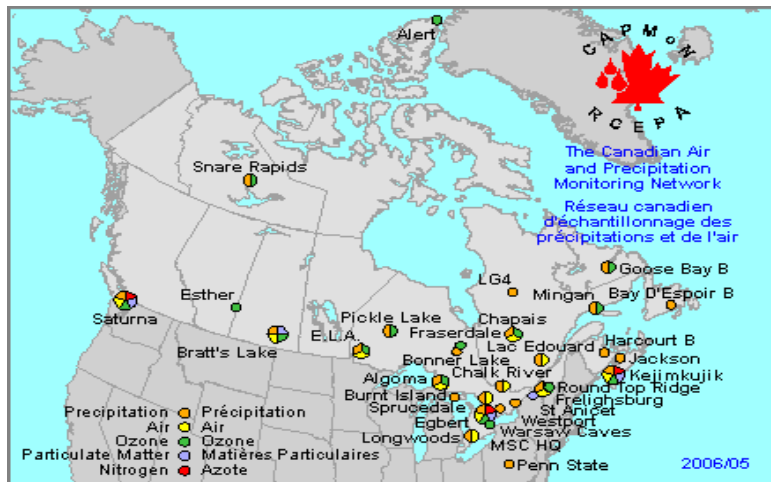
Term 1, 2017/18

# Designing monitoring networks

## In this lecture you will learn:

- Uncertainty, information & **the need for monitoring**
- 
- That getting **more information can increase not decrease your uncertainty** for the phenomenon of interest
- The difference between **aleatory uncertainty** and **epistemic uncertainty**
- How to **optimally redesign a monitoring network**
- **New challenges facing designers**

# Example # 1 Capmon network.



## NOTES on Capmon network

- No sense an “optimal” network for monitoring the environment.
  - For administrative simplicity Capmon was a merger of three networks, each setup to monitor acid precipitation when that topic was fashionable.
  - For simplicity, the sites were then adopted for other things, e.g, air pollution
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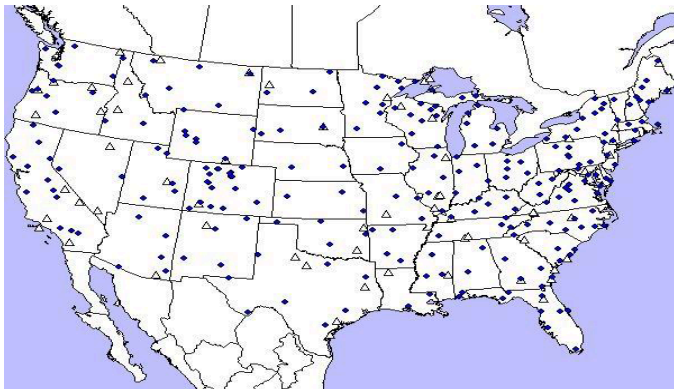
### Lessons learned:

A network's purposes often diverse & unforeseen.



## Example #2. NADP/NTN network

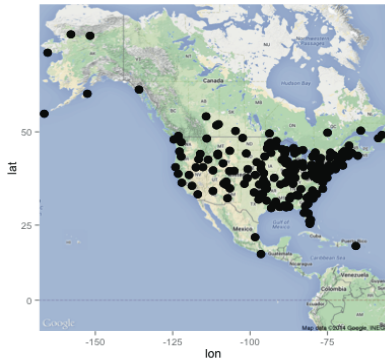
Monitors multivariate responses related to “acid precipitation”—  
another network merger with better defined siting rules!



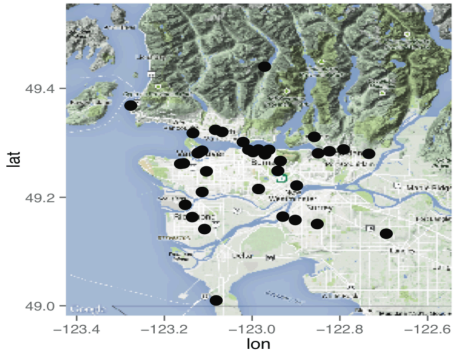
## **Rules governing siting and types of NDP/NTN monitors:**

*“The COLLECTOR should be installed over undisturbed land on its standard 1 meter high aluminum base. Naturally vegetated, level areas are preferred, but grassed areas and up or down slopes up to 15% will be tolerated. Sudden changes in slope within 30 meters of the collector should also be avoided. Ground cover should surround the collector for a distance of approximately 30 meters. In farm areas a vegetated buffer strip must surround the collector for at least 30 meters.” :*

# Example # 3. NA mercury monitoring sites



# Example # 4. Vancouver air quality monitors

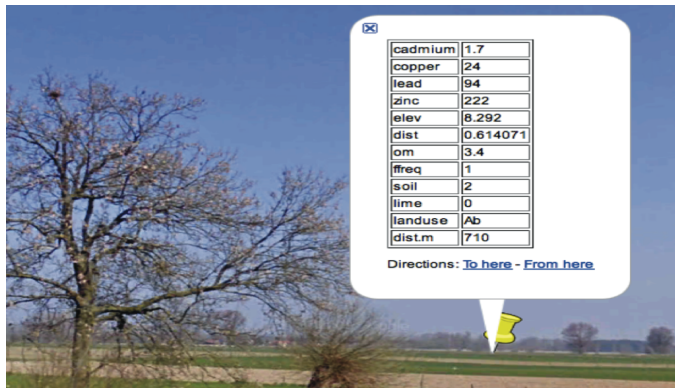


## Example # 5. Meuse river bank soil sampling sites



# The “street view”


Click on yellow **tack**:



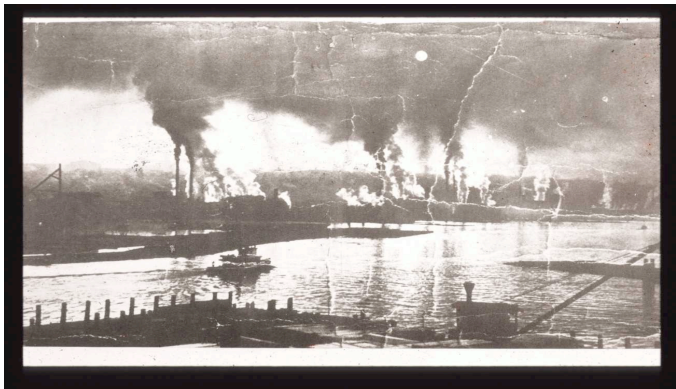
✕

cadmium	1.7
copper	24
lead	94
zinc	222
elev	8.292
dist	0.614071
om	3.4
ffreq	1
soil	2
lime	0
landuse	Ab
dist.m	710

Directions: [To here](#) - [From here](#)



## False creek circa 1900 before monitoring



*Courtesy of Professor Douw Steyn*

## False creek circa 1990



*Courtesy of Professor Douw Steyn*



# What do monitoring sites look like?

## At Kitsilano High School



*Courtesy of Professor Douw Steyn*

# What do monitoring sites look like?

## Near Robson Square



*Courtesy of Professor Douw Steyn*

## Some specific objectives of monitoring

- **Measure process responses** at critical points:
  - Near a new smelter using arsenic
- **Enable predictions** of unmeasured responses
- **Enable future forecasts**
- **Estimation** of process parameter
  - physical model parameters
  - stochastic model parameters eg. covariance parameters
- **Address societal concerns**

- **Non-compliance detection** given regulatory standards
- **Health risk analysis**
  - & provide good estimates of relative risk
  - determine how well sensitive sub-populations are protected
  - can include all life, not just human
- **Time trend analysis**
  - are things getting worse?
  - is climate changing?

# General purposes

## Explore/reduce uncertainty re the environment:

- One form of uncertainty (***aleatory***) is irreducible (e.g. outcome of fair die toss)
- The other (***epistemic***) (e.g. whether the die is fair) can increase or decrease. Implication: even an optimum design must be regularly revisited

## But what is “uncertainty”?

- **Laplace:** “Probability is the language of uncertainty”
- **DeFinetti:** “In life uncertainty is everything”
- **Kolmorov & Renyi:** “Entropy”
- **Statisticians:** “variance” or “standard error”

# Exercises

**8.1** Suppose  $X \sim N(0, 1)$ . Prove that uncertainty about  $X$ , i.e.  $\text{Var}(X \mid |X| < C)$  is increasing function of  $C$ .

**8.2** Suppose  $X \sim N(\eta, 1)$ . Prove that uncertainty about  $X$ , i.e.  $\text{Var}(X \mid |X| < C)$  is increasing function of  $C$ . Warning: Very hard! <sup>1</sup>

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<sup>1</sup>I posed the problem years ago with a prize of \$100 but it was not solved until Jiahua Chen (UBC) did so Chen et al. [2010].

# Possible design criteria

## Add monitors to (i.e. “Gauge”) sites:

- **that maximally reduce uncertainty about their responses** at their space–time points [because then measuring their responses eliminates their uncertainty]
- **best minimize uncertainty** about ungauged site responses
- give best **process parameter estimates**
- best **detect non-compliers**



## Designer challenges:

- **Multiplicity of objectives**
- **Unforeseen & changing objectives**
- **Multiple responses** at each site: which to monitor?
- Including **prior knowledge**
- **Good process models**
- **Fit with existing networks**
- **Fit with reality!!!**

# Exercises

**8.3** How might design criteria be arrived at in practice? Who should be responsible for setting them?

**8.4** Monitor placement should recognize such things as the geographical distribution of impacted populations (eg trees or fish). How can an optimal design be determined in such a context?  
Research question!

**8.5** Develop a design theory in a non-Gaussian context. Research question!

# Approaches to design

A big field [Zidek and Zimmerman, 2009]

- **Space-filling designs**
- **Probability based designs**
  - simple random sampling
  - **stratified, multistage designs**
  - e.g. (1) EPA's survey of lakes; (2) the EMAP project
- **Model Based**
  - **Regression model** approach
    - eg to estimate the slope put 1/2 the data at each end of the data range
  - Random fields (prediction, e.g. **entropy**) approach
- **Other.** In particular Zhengyuan Zhu (UNC) incorporates both of the latter, prediction and parameter estimation.

# Entropy based approach

## **“Gauges” sites with greatest “uncertainty”**

- **uncertainty = entropy**
- **maximally reduces uncertainty about “ungauged” sites**
- **best estimates predictive posterior distribution under entropy utility**

**By - passes specification of objectives**

**Long history<sup>2</sup>, currently popular**

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<sup>2</sup>General: Good [1952], Lindley [1956], Shewry and Wynn [1987].  
Network design: Caselton and Zidek [1984], Sebastiani and Wynn [2002], Zidek et al. [2000]

# What is entropy?

Let **probability of uncertain event E** (e.g. heads on bent coin) be:

$$p = P(E)$$

That uncertainty becomes 0 when outcome becomes known.

Let **reduction in uncertainty** be

$\phi(p)$  if  $E$  occurs

$\phi(1 - p)$  if  $E$  does not occur

**Expected reduction in uncertainty:**

$$p\phi(p) + (1 - p)\phi(1 - p)$$

## Simple assumptions imply

:

$$\phi(p) = \log(p)$$

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**Conclusion:** knowing  $E$  occurred reduces entropy by

$$p \log(p) + (1 - p) \log(1 - p)$$

Thus “**uncertainty**” about  $E$  can be quantified as the **entropy** of the two point distribution  $(p, 1 - p)$ :

## Relative entropy

But how much is that entropy?

Needs a **reference level**. Complete uncertainty about  $E$  would be the two point distribution  $(q, 1 - q)$  with  $q = 1/2$ . Thus the relative entropy would be

$$I(p, q) = p \log \left\{ \frac{p}{q} \right\} + (1 - p) \log \left\{ \frac{(1 - p)}{(1 - q)} \right\}$$

This is f the **Kullback-Leibler** measure of the deviation of  $(p, 1 - p)$  from that reference level (the “state of equilibrium” in physics (thermodynamics)).

## Multiple events

$$I(p, q) = \sum_i p_i \log \{p_i/q_i\}$$



## Continuous variables

Start with  $p_i \sim f(x_i)dx_i$  &  $q_i \sim g(x_i)dx_i$  as approximations. Then as  $dx_i \rightarrow 0$ , this entropy converges to

$$I(f, g) = \int f(x) \log \left\{ \frac{f(x)}{g(x)} \right\} dx > 0$$

Commonly  $g \equiv 1$  (*unitsoff*). In any event,  $f/g$  is a unitless quantity. Moreover Jacobean cancels under transformations of  $x$  making entropy an “intrinsic” measure of uncertainty – not scale dependent.

# Entropy framework

Adopt a Bayesian framework.

- $\theta$ : process parameter vector.
- $\mathbf{Y}$  : process response vector at future time  $T+1$  including all sites (gauged & ungauged) conditional on
- $D_T$  set of all available data at time  $T$
- $h_1$  &  $h_2$ : baseline reference densities against which to measure uncertainty.

- Finally compute the entropies **with  $\theta$  random**

$$H(\mathbf{Y} | \theta) = E\left[-\log \left\{ \frac{f(\mathbf{Y} | \theta, D)}{h_1(\mathbf{Y})} \right\} \mid D_T\right]$$

$$H(\theta) = E\left[-\log \left\{ \frac{f(\theta | D_T)}{h_2(\theta)} \right\} \mid D_T\right]$$

Then we get **fundamental entropy identity**

(**Exercise**):

$$H(\mathbf{Y}, \theta) = H(\mathbf{Y} | \theta) + H(\theta)$$

**In other words**, at time  $T$

**TOTAL UNCERTAINTY =  
UNCERTAINTY ABOUT RESPONSE GIVEN MODEL  
+ MODEL UNCERTAINTY**

## Design goal: add sites

Add (or subtract) sites at time  $T + 1$ . Let's focus on  
adding new sites to an existing network

- $\mathbf{Y} = (\mathbf{Y}^{(1)}, \mathbf{Y}^{(2)})$  = all site responses, time  $T + 1$
  - $\mathbf{Y}^{(2)}$  for site currently gauged (time  $T$ )
  - $\mathbf{Y}^{(1)}$  for sites currently ungauged (time  $T$ )
- 

**DESIGN GOAL:** Partition  $\mathbf{Y}^{(1)} = (\mathbf{Y}^{(rem)}, \mathbf{Y}^{(add)})$  at time  $T$  so that

- $\mathbf{Y}^{(rem)}$ : future ungauged sites
- $\mathbf{Y}^{(add)}$  future new network stations.

# Entropy decomposition thm

Let  $\mathbf{U} = \mathbf{Y}^{(rem)}$ ;  $\mathbf{G} = (\mathbf{Y}^{(add)}, \mathbf{Y}^{(2)})$ ;  $\mathbf{Y} = [\mathbf{U}, \mathbf{G}]$

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**Fundamental identity:**

$$\boxed{\text{TOT} = \text{PRED} + \text{MODEL} + \text{MEAS}}$$

where

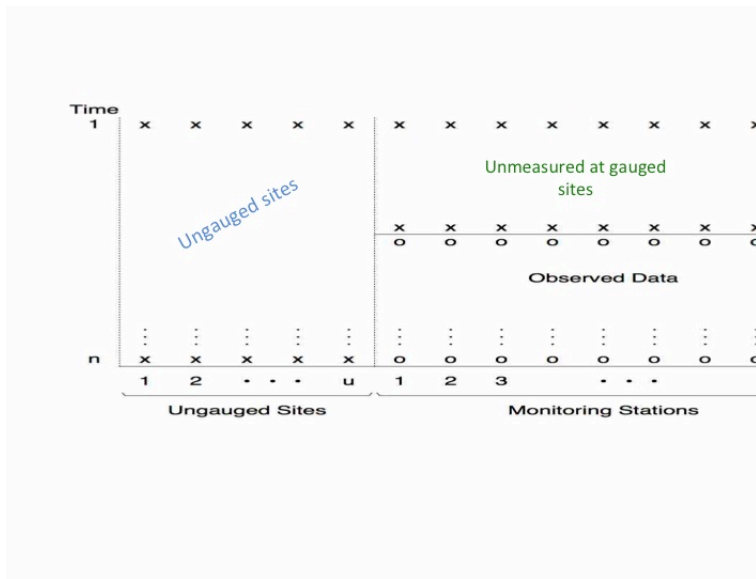
$$\begin{aligned} \text{PRED} &= E[-\log(f(\mathbf{U} | \mathbf{G}, \theta, D_T)/h_{11}(\mathbf{U} | D_T)), \\ \text{MODEL} &= E[-\log(f(\theta | \mathbf{G}, D_T)/h_2(\theta)) | D_T], \end{aligned}$$

and

$$MEAS = E[-\log f(\mathbf{G} | D_T)/h_{12}(\mathbf{G} | D_T)]$$

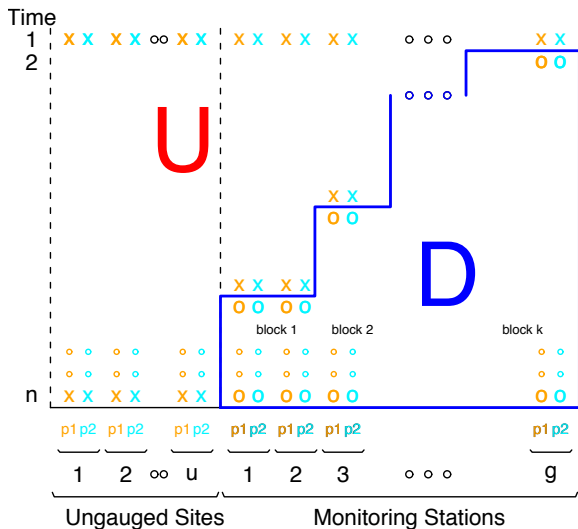
**Theorem:** Maximizing MEAS = Minimizing MODEL + PRED

# The environmental process





# The multivariate data staircase



## Possible inferential objectives

- **Forecasting:** process values at time  $T + 1$
- **Spatial prediction:** process values in U at time  $T$
- **Hindcasting:** past values of the process at times  $t < T$  both in D sites as well as U sites, e.g. cancer vs air pollution.

## Design objective for expanding the network:

Choose sites from **U** to “add” to the existing network at time  $T + 1$ .

**NOTATION:**  $u$ : ungauged sites at time  $T$

$g$ : gauged sites at time  $T$   $p = u + g$ : total number of sites

## The process model

**Transform responses** as necessary. **Remove regular temporal & trend components.** Assume **EnviroStat model:** at time  $t = 1, \dots, T + 1$

$$\mathbf{Y}_t^{1 \times p} \mid \beta, \Sigma \stackrel{ind}{\sim} N(\mathbf{x}_t^{1 \times k} \beta^{k \times p}, \Sigma^{p \times p})$$

$$\beta \mid \Sigma, \beta_0, F \sim N(\beta_0, F^{-1} \otimes \Sigma)$$

$$\Sigma \sim GIW(\Psi, \delta) \text{ \# Inverted Wishart distribution}$$

where  $\otimes$  denotes **Kronecker product**,  $F^{-1}$  covariance between rows;  $\Sigma$  covariance between columns.



# The predictors

The  $x_t$  are assumed to be the same for all sites in the region

**Question:** what about site specific predictors?

- Random predictor at site  $j$   $W^j$  of  $Y^j$ :
- model  $[Y^j, W^j]$  then  $[Y^j | W^j]$
- Nonrandom predictor  $w^j$ : fit  $\beta_0^j = \gamma w_j$  via empirical Bayes

## Matric-normal extension

$$: \quad \mathbf{Y}^{(T+1) \times p} \mid \boldsymbol{\beta}, \boldsymbol{\Sigma} \sim N(\mathbf{x}^{(T+1) \times k} \boldsymbol{\beta}^{k \times p}, \mathbf{I}_{T+1} \otimes \boldsymbol{\Sigma}^{p \times p})$$

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**NOTE:** The IW distribution is the inverse  $\chi^2$  for matrices. It can be generalized to the GIW - it had different numbers of degrees of freedom for different steps in the data staircase for example.

## Bartlett decomposition:

**Reading Notation:** “u” means “ungauged” sites; “g” means “gauged” sites (with  $p = u + g$ ):

$$\Sigma = \begin{pmatrix} \Sigma^{[u]} & \Sigma^{[ug]} \\ \Sigma^{[gu]} & \Sigma^{[g]} \end{pmatrix}$$

**Bartlett decomposition:**

$$\Sigma = T \Delta T'$$

where

$$T = \begin{pmatrix} I & \Sigma^{[ug]}(\Sigma^{[g]})^{-1} \\ 0 & I \end{pmatrix}$$
$$\Delta = \begin{pmatrix} \Sigma^{[u]} - \Sigma^{[ug]}(\Sigma^{[g]})^{-1}\Sigma^{[gu]} & 0 \\ 0 & \Sigma^{[g]} \end{pmatrix}$$



# Generalized Inverted Wishart Distribution

## Reading

Let

$$\Gamma^{[u]} = \Sigma^{[u|g]} = \Sigma^{[u]} - \Sigma^{[ug]}(\Sigma^{[g]})^{-1}\Sigma^{[gu]}$$

$$\tau^{[u]} = (\Sigma^{[g]})^{-1}\Sigma^{[gu]}.$$

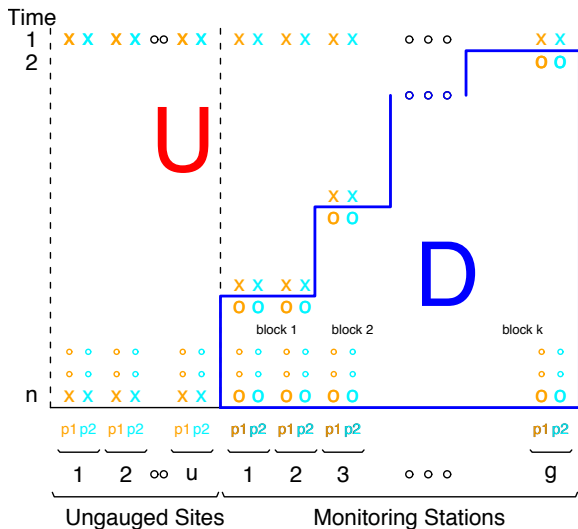
**Definition:** For  $\Sigma \sim GIW(\Psi, \delta)$ , iterate the following decomposition starting with  $\Sigma^{[g]}$  to get successive deg's freedom  $\delta_0, \delta_1, \dots$ :

$$\Sigma^{[g]} \sim GIW(\Psi^{[g]}, \delta^{[g]})$$

$$\Gamma^{[u]} \sim IW(\Lambda_0 \otimes \Omega, \delta_0)$$

$$\tau^{[u]} | \Gamma^{[u]} \sim N(\tau_{0u}, H_0 \otimes \Gamma^{[u]})$$

# Reminder: the multivariate data staircase



## Additional notation

$\mathbf{g}_i^o$  denotes **observed responses at gauged site  $i$**  sites

$\mathbf{g}_i^m$  denotes **missing data at gauged site  $i$** ,  $i = 1, \dots, g$ .

Then  $D_T = \{\mathbf{g}_1^o, \dots, \mathbf{g}_g^o\}$ .

Then [Le and Zidek, 2006]:

$$\begin{aligned} \left[ \mathbf{Y}^{[u]} \mid D_T, \mathcal{H} \right] &\sim \left[ \mathbf{Y}^{[u]} \mid \mathbf{Y}^{[g_1^m, \dots, g_k^m]}, D_T, \mathcal{H} \right] \times \\ &\prod_{j=1}^{k-1} \left[ \mathbf{Y}^{[g_j^m]} \mid \mathbf{Y}^{[g_{j+1}^m, \dots, g_k^m]}, D_T, \mathcal{H} \right] \\ &\times \left[ \mathbf{Y}^{[g_k^m]} \mid D_T, \mathcal{H} \right]. \end{aligned}$$

and  $\mathbf{Y}^{[u]}$  only appears in first factor:

$$\left( \mathbf{Y}^{[u]} \mid \mathbf{Y}^{[g_1^m, \dots, g_k^m]}, D, \mathcal{H} \right) \sim$$

$$t_{n \times u} \rho \left( \mu^{[u|g]}, \text{Dispersion}, \delta_0 - u + 1 \right).$$

$$\text{Dispersion} = (\delta_0 - u + 1)^{-1} \Phi^{[u|g]} \otimes (\Lambda_0 \otimes \Omega)$$

- Here  $\Lambda_0$  represents spatial covariance that links  $U$  to  $G$
- $\Omega$  represents within site covariance e.g. a four hour block of ozone concentrations or several different chemical species
- $\delta_0 - u + 1 > 0$  is required to avoid a degenerate t-distribution. No such thing as a free lunch. Must keep  $u$  to realistic size. Don't see this with purely Gaussian process models.

$$\text{Dispersion} = (\delta_0 - u + 1)^{-1} \phi^{[u|g]} \otimes (\Lambda_0 \otimes \Omega)$$

## Entropy calculation

The relevant entropy for  $Y^{[u]}$  is from that first factor in the entropy decomposition:

$$H \left[ Y^{[u]} \mid Y^{[g_1^m, \dots, g_k^m]}, D \right] = \frac{\rho}{2} \log |\Lambda_0| + \text{irrelevant terms}$$

- Ungauged sites  $u$  must be partitioned into ‘add’ and ‘rem’ sites in optimal way.
- Applying the Bartlett decomposition tells us we must find the submatrix of  $|\Lambda_0[add, add]|$  whose sub-determinant  $|\Lambda_0|$  corresponds to the ‘add’ sites in the partitioned  $Y^{[u]}$ .

**NOTE:** Will simultaneously minimize the entropy left in the ‘rem’ sites.

## Finding the 'add' sites

- **NP-Hard:** No exact algorithms for big networks
- Inexact Methods:
  - **Greedy** - add (or subtract) one at a time
  - **Reverse Greedy** - subtract one at a time
- Exact Methods:
  - Complete **enumeration**
  - **Branch and bound**



# How many sites?

**Compute:**

$$\frac{\text{Entropy}(n)}{n}$$

where  $n$  is the number of sites for success  $n$ 's. The ratio will initially increase than decrease. Stop at the maximum (bang for the buck)  $n$ .

## Application: Redesign Vancouver's hourly $PM_{10}$ monitoring field.

- Increase **10** monitoring sites to **16**—add **6** new stations from among **20** possible sites <sup>3</sup>.
- Use entropy approach with **Normal–GIW predictive distribution**.

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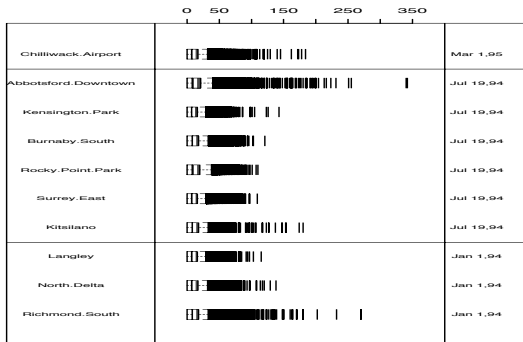
<sup>3</sup>Resembles actual study Nhu Le & I did [Ainslie et al., 2009]

## Method (Will see details in the Practicum:

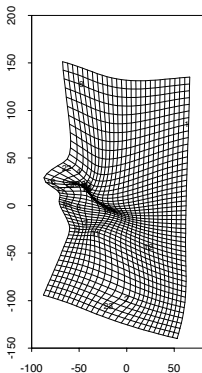
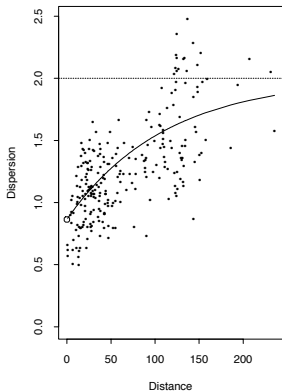
- **Transform data**
- **Remove all-site (common) spatio-temporal trends**  
This means fit the same model to all sites.
- **Whiten” residuals** with common 24 (hour) dimensional multivariate AR(1) model  
same idea - result is approximately unautocorrelated temporal process
- 10 different station startup times means **monotone (“staircase”) data pattern** – use GIW distribution with different degrees of freedom ( $\delta$ s ) for each staircase step
- **Select the 6 new stations** that jointly maximize conditional entropy

# Preliminary data analysis:

$PM_{10}$  levels at the 10 existing stations. **Note differing startup times.**

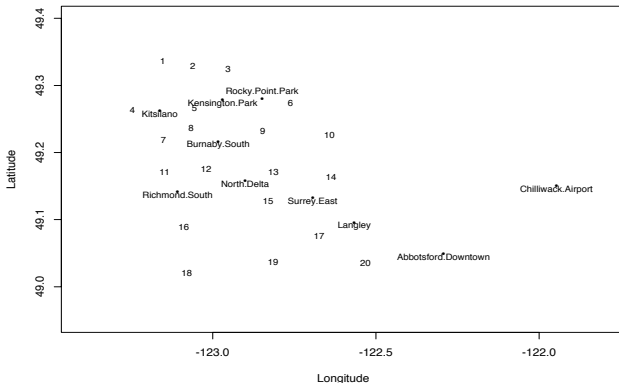


Vancouver's ozone field is clearly non-stationary and the Sampson–Guttorp method was used to estimate the hypercovariance in the GIW matrix.

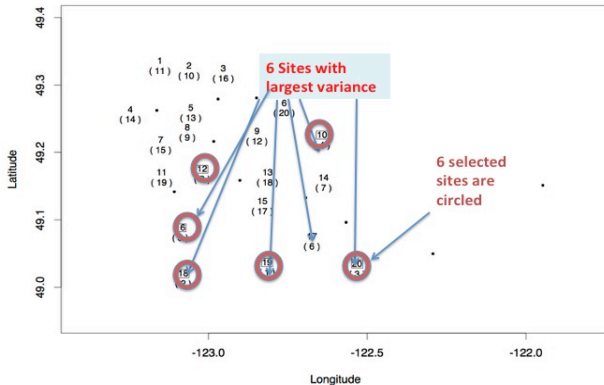


# The original 10 PM<sub>10</sub> monitoring sites

Also **20** possible new locations – subregions with no existing monitors & big populations.



Locations of the old and **newly selected 'add' sites (square brackets)**. The ranks of the **20** sites by estimated variance is in curved brackets and those selected. Notice **6<sup>th</sup>** selected site.



# Alternative design strategies: #1

**Non-Compliance detection design:** Select sites with max prob of finding sites out of compliance with regulatory standards.

## **Probability & best design is day-dependent!**

- which day?
- a simulated future day? Average day? Bad day?

## **How to implement?**

- monitor sites most likely to noncomply?
- do not monitor sites least likely?
- what about existing sites?



**Example:** What if the 6 sites had been selected to catch noncompliance rather than to maximize entropy?

Say we use 10 station hourly data for  $PM_{10}$ ,  
February 28, 1999 & hierarchical Bayes predictive  
distribution

**DETECTION CRITERION:**

$$\text{argmax PR}\{\text{daily max } PM_{10} Y^{6\text{added}} \geq 50 (\mu\text{g m}^{-3})\}$$

**RESULT:** Same as for entropy based design!!!

## NOTES:

- 1 Entropy does not work for detecting noncompliance nearly as well on August 1, 1998!
- 2 The selected new sites are now determined pretty much by their posterior estimated variances. That is because the spatial correlation is now quite small – strength hard to borrow.
- 3 Designing for noncompliance seems largely unexplored issue

## Alternative design strategy #2

**Site selection for fields of extremes:** Very difficult since spatial dependence declines when looking say at monthly maxima rather than daily maxima.

Yet regulatory criteria metrics (risk) usually base on extremes!

**Example:** EPA'S  $PM_{10}$  criterion:

*For particles of diameters of 10 micrometers or less:*

Annual Arithmetic Mean:  $50 \mu\text{g m}^{-3}$

24 - hour Average:  $150^{FN} \mu\text{g m}^{-3}$

The three year average of 98-th annual percentiles of 24 hour averages must be  $\leq 150 (\mu\text{g m}^{-3})$  at all sites in an urban area.  
Complex metric  $\Rightarrow$  need predictive distribute to simulate its distribution!

# The bad news for fields of extremes

- 1 **Insufficient data, spatial and temporal.**
- 2 **Extremes have small inter - site dependence**
  - **between some site pairs, not others**
- 3 **Conventional approaches fail**
- 4 **Multivariate extreme value distributions - not tractable**
  - **conditional computation (e.g. entropy) difficult**
  - **simulating extreme fields hard**
- 5 **Elusive design objective**

# The good news

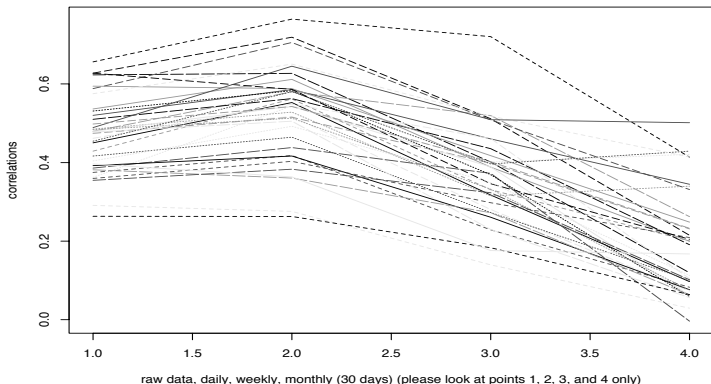
**Joint distribution of extremes approximately a log multivariate t distribution. Hence can:**

- 1 have convenient conditional, marginal distributions**
- 2 accommodate existing sites and historical data**
- 3 permit simulation of complex metric distributions**
- 4 have explicitly computable entropy's, regression models, etc**
- 5 can enable “elusive objectives issue” to be bypassed**

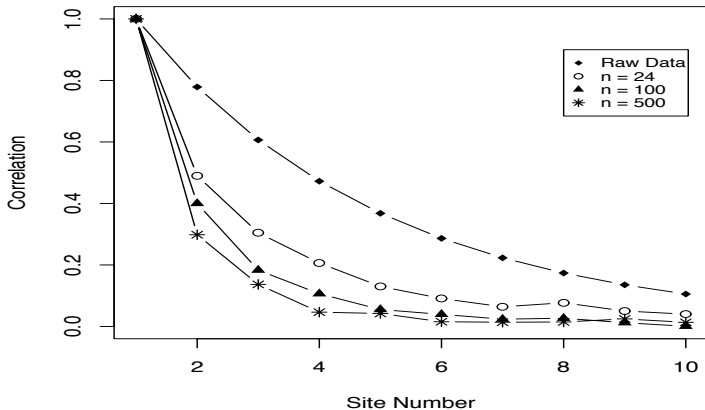
## Some detail: Inter-site correlations

**Inter-site dependence declines with time span** of extremes for many but not all site pairs

**Example:** The figure shows Vancouver's  $PM_{10}$  intersite correlation between maxima computed for various time spans.



**Simulation study<sup>4</sup>**: multinormal responses; maxima with varying ranges at 10 sites. Multivariate  $t$  results show smaller loss of dependence. Inter-site correlations for maxima for simulated fields of extremes. Big  $n$  = light tails.



<sup>4</sup>Chang et al. [2007]

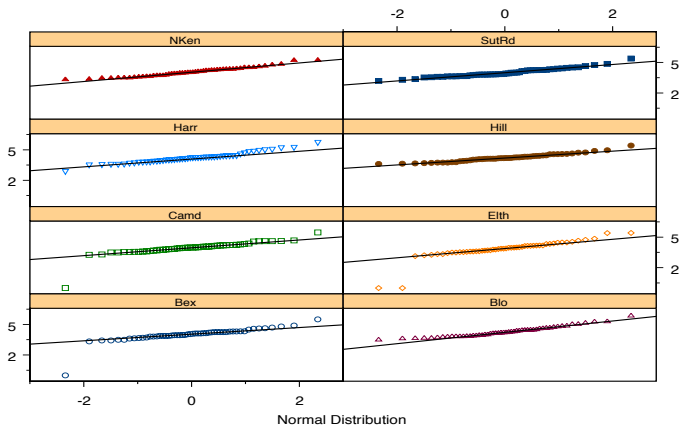
# Exercises

**23.6** Try the simulation experiment yourself and confirm that for heavier tails lead to increased intersite correlations



# An approach to monitoring extremes

Empirical results  $\mapsto$  log multivariate - t distribution as approximation to joint distribution of extremes field. QQplots for weekly maxima of hourly log  $PM_{10}$  London 1997 data  $\rightarrow$  **marginal normality of extremes:**



## T approx approach cont'd

**Empirical results** → **well-calibrated 95% (etc) prediction intervals. Supports use of multivariate approximation.**

Credibility Level	Mean	Median
30%	35	35
95%	96	97
99.9%	99.9	1

**Table : Summary of coverage probabilities at different credibility levels for the simulated precipitation data over 319 grid cells, Canadian Climate Model**

# Room for optimism?

**Use of log multivariate t distribution for extreme fields promising. But:**

- **How far can approximation go? Need new theory**
- **Must test approximation on case-by-case basis**
- **need to compare regular - and extreme-entropy designs.**

# Summary

In this lecture we have seen:

- Why environmental monitoring networks are needed and what they look like
- Why spatio-temporal theory is needed in their design
- A class of models with associated software that provide a platform for designing them.
- The entropy and other criteria for design.

# Regulation and control

We saw that networks are needed for control and mitigation.  
Also saw some issues arising:

- Current designs need to be revisited - fitness for use.
- Simpler control metrics than the current ones would be preferable for transparency and analysis.
- The current design criteria have not been well spelled out.

## Some conclusions

- Current urban networks may be inadequate for surveillance of extremes. Much more attention needs to be paid
- But the MaxEnt strategy may be a way out of this problem.
- The EnviroStat approach needs an upgrade that takes account of recent theoretical developments .

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