STAT 340: Introduction to Probability and Mathematical Statistics I

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Class information

Lectures MF TA section W Homework due Mondays Midterm Friday November 5

Grading scheme: 5% class participation 30% homework (drop worst) 30% midterm 35% final (MUST take final)

Probability and statistics

Probability is the science of randomness –regularity in irregularity Build probability models to describe observed phenomena

Statistics is a *system for formal inference* based on probability models

Estimate parameters of model Test hypotheses Assess uncertainty

Climate change?

IPCC

AR4 WG1 Ch. 11 (2007): Fewer cold outbreaks; fewer, shorter, less intense cold spells / cold extremes in winter VL (consistent across model projections) Northern Europe, South Asia, East Asia L (consistent with warmer mean temperatures) Most other regions Reduced diurnal temperature range L (consistent across model projections) Over most continental regions, night temperatures increase faster than day temperatures







London smog 1952

5 days of pea soup fog Visibility at times a few feet 12 000 deaths 100 000 made ill





Smog in Beijing



Health effects of ozone

Decreased lung capacity Irritation of respiratory system Increased asthma hospital admissions Children particularly at risk

Ozone regulation

In each region the expected number of daily maximum 1-hr ozone concentrations in excess of 0.12 ppm shall be no higher than one per year

Implementation: A region is in violation if 0.12 ppm is exceeded at any approved monitoring site in the region more than 3 times in 3 years

Does this rule protect the people?

Rainfall measurement

Rain gauge (1 hr) High wind, low rain rate (evaporation) Spatially localized, temporally moderate Radar reflectivity (6 min) Attenuation, not ground measure Spatially integrated, temporally fine Cloud top temp. (satellite, ca 12 hrs) Not directly related to precipitation Spatially integrated, temporally sparse Distrometer (drop sizes, 1 min) Expensive measurement Spatially localized, temporally fine Radar image











Compartment 2. Differentiating clones





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Photon

counter

Probability theory

EXPERIMENT

Repeatable (controlled conditions)

Measurable well-defined outcomes

In any repetition there is always uncertainty as to which of the possible outcomes will occur

SAMPLE OUTCOME

s denotes a possible outcome of the experiment

SAMPLE SPACE

S denotes the set of sample outcomes (also called the outcome set)

EVENT

A is a subset of S; a description of the result of the experiment

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Some experiments

Experiment: Roll a die and note which face is showing.

Sample space:

Event: Even outcome =

Experiment: A hospital has two sources of emergency power. We observe which (if any) of the emergency sources are functioning in the case of a power failure.

Sample space:

Event: The hospital has power during a power failure =





A and B are mutually exclusive if $A \cap B = \emptyset$

Examples

A dashboard warning light is supposed to flash if the car's oil pressure is too low. We are interested in the event E = "warning light flashes when oil pressure is too low"

S =

A = Oil pressure too low =

B = Warning light flashes =

E =

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Let S be the plane and define the events $A = \{ x, y \mid 0 < x < 3, 0 < y < 3 \}$ $B = \{ x, y \mid 2 < x < 4, 2 < y < 4 \}$ What is $A \cup B$? $A \cap B$?





Probability axioms

Probability is a function from S to the reals such that (1) $P(A) \ge 0$ for all subsets A of S (2) P(S) = 1(3) For mutually exclusive A and B $P(A \cup B) = P(A) + P(B)$

Alexei Kolmogorov 1903-1987 Foundations of the Theory of Probability (1933)



Summary of last lecture

Experiment Outcomes Sample space Events Venn diagrams Probability axioms

Office hours in B302: Bailey F 9-11 Peter Th 8:30-10:30













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Midterm?

Conditional probability

Define $P(A|B) = \frac{P(A \cap B)}{P(B)}$

for B such that P(B) > 0.

Fact: P(A|B) is a probability on the sample space B.

Corollary: $P(A \cap B) = P(A|B)P(B)$

The older child

Pick a family at random among those with two children. Let B = {both children are girls} P(B) =

What is the conditional probability of B, given there is at least one girl?

What is the conditional probability of B, given that the older child is a girl?

Total probability

Let $A_1,...,A_n$ be a *partition* of S, i.e. $A_i \cap A_j = \emptyset$, $P(A_i) > 0$ and $\bigcup_{i=1}^n A_i = S$ Then

$$\mathbf{P}(\mathbf{B}) = \sum_{i=1}^{n} \mathbf{P}(\mathbf{B} | \mathbf{A}_{i}) \mathbf{P}(\mathbf{A}_{i})$$

Let's Make a Deal

<u>Game</u>

If the contestant picks door k, what is the probability that the car is behind door i given that door j is opened?

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A con game

Three cards are put into a hat. One is white on both sides, one is red on both sides, and one is red on one and white on the other side. A card is drawn from the hat and placed on a table, without showing the underside. Suppose the top of it is red. The con man offers to bet even money on the bottom also being red. Why should you not take the bet?

Last Friday

Showed that conditional probability is probability on a new sample space Proved the theorem of total probability Examples Problems

Solutions to problems

1. 3,3,3 can only come out in one way, 3,3,4 in three ways (think of the dice as different color). All the other outcomes can be matched up so there are equally many outcomes (6 when all the dice are different, 3 when two dice are the same). In total there are 27 outcomes that yield 10 and 25 that yield 9 (out of 216). (Galileo, ~1640)

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2. (a) .25+.05+.05 = .35 (b) .04+.10+.05+.04+.05+.03 = .31 (c) .25/(.25+.04+.10) = .64 (d).25/(.25+.05+.05) = .71
3. (a) P(A _k)=c/k, and A ₁ =S so c=1. A _k = {s _k } + A _{k+1} so P(A _k) = P({s _k }) + P(A _{k+1}), or P({s _k }) = 1/k - 1/(k+1) [= 1/(k(k+1))]. (b) P(B) = $\sum_{j=0}^{\infty} P(\{s_{2j+1}\}) = \sum_{j=0}^{\infty} \left[\frac{1}{2j+1} - \frac{1}{2j+2}\right]$ = $1 - \frac{1}{2} + \frac{1}{3} - \frac{1}{4} + \dots = \ln 2 = .69$
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Constructing the tree

We are told the following three probabilities: P(heard the advertisement) = 0.35 P(bought the product) = 0.23 P(heard the ad and bought the product) = 0.15





Rules for probability trees Probabilities add to one at each level of the tree Conditional probabilities add to one for each group of branches radiating from a single point

The probability at a branch point times the conditional probability along the branch is the probability at the end of the branch

The probability at a branch point is the sum of all probabilities at the end of branches to the right of the point



Bayes' theorem

Let $A_1,...,A_n$ be a partition of S. Assume P(B) > 0. Then

$$\mathbf{P}(\mathbf{A}_{i}|\mathbf{B}) = \frac{\mathbf{P}(\mathbf{B}|\mathbf{A}_{i})\mathbf{P}(\mathbf{A}_{i})}{\sum_{i=1}^{n}\mathbf{P}(\mathbf{B}|\mathbf{A}_{i})\mathbf{P}(\mathbf{A}_{i})}$$

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 $\begin{array}{l} \textbf{Screening for} \\ \textbf{cervical cancer} \\ <section-header>$ $\begin{array}{l} A = \{positive Pap smear\} \\ B = \{cervical cancer\} \\ P(B) = 0.00081 (2002-2005) \\ P(A|B) = 0.72 \\ P(A|B^{\circ}) = 0.06 \\ \end{array}$ $\begin{array}{l} P(B|A) = \frac{0.000058}{0.0601} = 0.0010 \\ 0.0601 \\ \end{array}$ $\begin{array}{l} \textbf{What is the error rate for the Pap smear} \\ \textbf{smear} \end{array}$

Monday's lecture

Solutions to Friday's problem set Probability trees Bayes' formula

Asking sensitive questions

How can you get accurate answers to questions about drug use, cheating etc.

A: Is your student number even? B: Have you ever cheated at an exam at the University of Washington? Flip a coin. Heads answer A, tails answer B. DO NOT SAY which question you are answering! P(even student number) = 24/46 Result: Yes 18 No 20 Invalid 1

US Census data

Race	Male	Female	All
White	1,487	1,412	2,899
Nonwhite	366	348	714
Total	1,853	1,760	3,613

Data on newborns from 1980 US census in thousands of individuals. Consider a randomly chosen baby.

P(female) =

P(female | nonwhite) =

Independence

A and B are independent if P(A|B) = P(A). Equivalently $P(A \cap B) = P(A)P(B)$



Genetics, cont.

By total probability $P(A_w) = P(A_w | WW) P (WW)$ $+ P (A_w | Ww) P (Ww)$ $+ P (A_w | ww) P (ww)$ = 0 x p + 1/2 x q + 1 x r $= r + q/2 = P (B_w)$

Genetics, cont.

By the independence assumption the probability of the offspring receiving the recessive trait, i.e., genotype ww, is

$$P(A_{w} \cap B_{w}) = P(A_{w})P(B_{w}) = (r + q / 2)^{2}$$

How about WW?

One of each allele?

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These probabilities are called the *Hardy-Weinberg* proportions.

More than two events

We can't define independence between three events by pairwise independence, or by saying $P(A \cap B \cap C) = P(A)P(B)P(C)$

We need this: $A_1,...,A_n$ are independent if for all distinct $i_1,...,i_k$ with $1 \le i_j \le n$ $P(A_{i_k} \cap \cdots \cap A_{i_k}) = P(A_{i_k}) \cdots P(A_{i_k})$ Breathing Ceasar's last breath

The number of molecules in the earth's atmosphere is about 10⁴⁴. The molecules present today are pretty much the same as those present thousands of years ago. A typical breath consists of about 10²² molecules. What is the chance that your next breath will contain at least one molecule from Caesar's last breath?

Friday's lecture

Randomized response survey Independence Examples Problems

Based on Bailey's survey I have moved my office hours to Thursday 1:30-3:20 pm in B302. Bailey's stay the same (Friday 9-11 in B302).

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Problem solutions

- 1. The easiest way to do this is with Venn diagrams. Answers: T F F F T T T
- 2. (a) No. Let A=C with P(A) = .5
 (b) No. Consider tossing two fair coins. A="heads on first", C="heads on second", B="same outcome on both"
 P(B) = 1/4+1/4 = 1/2

 $P(A \cap B) = P(A \cap C) = 1/4 = P(A)P(B)$

so A and B are independent, as are A and C.
$P(A \cup C) = P(A) + P(C) - P(A \cap C) = 3/4$
$P(B \cap (A \cup C)) = P(A \cap C) = 1/4 \neq 1/2 \times 3/4$ (c) No. Same setup as (b).
$P(B \cap A \cap C) = P(A \cap C) = 1/4 \neq 1/2 \times 1/4$ 3.(a) If B attracts A then
$P(A \cap B) > P(A)P(B)$
so A attracts B. Now write
$P(A B^{c})P(B^{c}) = P(A \setminus B) = P(A) - P(A \cap B)$
= P(A) - P(A B)P(B) < P(A) - P(A)P(B)
= P(A)P(B ^c)
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(b) Since P(F_j | B_i) = 1 for i ≠ j
Bayes' theorem yields
$$P(B_j | F_j) = \frac{P(F_j | B_j)P(B_j)}{\sum_{i=1}^{n} P(F_j | B_i)P(B_i)} = \frac{f_i b_j}{1 - b_j + f_i b_j}$$
$$P(B_j) - P(B_j | F_j) = \frac{b_j (1 - b_j)(1 - f_j)}{1 - f_j + b_j f_j} > 0$$
and for i ≠ j
$$P(B_i | F_j) - P(B_i) = \frac{b_i}{1 - b_j - b_j f_j} - b_i = \frac{b_i b_j (1 - f_j)}{1 - b_j - b_j f_j} > 0$$



5. A={0 recieved}, B={0 sent}						
Sent	Relay 1	Relay 2	Relay 3	Probability		
0	0	1	0	3/64		
0	1	1	0	3/64		
0	1	0	0	3/64		
0	0	0	0	27/64		
$P(A B) = \frac{(36 / 64)p}{(36 / 64)p + (28 / 64)(1-p)}$ where p is the frequency of 0. 6. A={6 on first} B= {6 on second} F={fair die picked} L={loaded die}						
$P(A B) = P(A B F)P(F) + P(A B L)P(L)$ $= \frac{1}{36} \times \frac{3}{9} + \frac{1}{4} \times \frac{6}{9} = \frac{19}{108}$						

$$P(A) = P(B) = \frac{1}{6} \times \frac{3}{9} + \frac{1}{2} \times \frac{6}{9} = \frac{7}{18}$$

$$P(A)P(B) = \frac{49}{324} \neq \frac{57}{324} = P(A \cap B)$$

so the two events are not independent.

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People vs. Collins

1964 LA purse snatching by a young blond female with ponytail, escaping with a black male with moustache and beard in a yellow car. A suspect with these characteristics was found. The state argued that it would be very unlikely that there would be more than one such couple.

Characteristic	Probability
Female with ponytail	1/10
Yellow automobile	1/10
Man with moustache	1/4
Blond female	1/3
Black male with beard	1/10
Interracial couple	1/1000
The product is 1/12 million.	

Bertillon system of identification

Predates fingerprints

11 anatomical variables, such as height, head width, ear length, which do not change much for adults

Each classified as small, medium, or large: e.g.,

(s, m, m, s, l, s, m, m, s, s, m)

How likely would it be that in Seattle there are two individuals with the same Bertillon configuration?