

# Spatio-temporal processes, including dynamic linear models

Dani Gamerman

Departamento de Métodos Estatísticos (DME)

Instituto de Matemática - UFRJ

dani@im.ufrj.br

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Presentation based on GAMERMAN, D. (2010). Dynamic Spatial Models, Including Spatial Time Series. Chapter 24 of Handbook of Spatial Statistics (Eds.: A. E. Gelfand, A. E., Diggle, P., Guttorp, P. e Fuentes, M.). London: Chapman & Hall.

# 1. Introduction

Consider a process  $X(s, t)$  varying with  $s \in S \subset R^d$  and  $t \in T \subset R$ .

I will be concentrating on  $d = 2$  and  $T = \{t_1, \dots, t_n\}$

[Theoretically, continuous  $T$  simply means that  $d \rightarrow d + 1$  ]

2 routes are usually taken for modeling:

- direct (observation-driven)
- latent (parameter-driven)

Routes are somewhat disjoint

## 2. Direct approach

Covariance function (CF):  $Cov[X(s, t), X(s', t')] = K(s, s'; t, t')$

Too general for reliable estimation without massive replications over  $S \times T$

Simplifying assumptions:

1) Temporal stationarity:  $K(s, s'; t, t') = f(s, s'; t - t')$

2) Spatial stationarity:  $K(s, s'; t, t') = f(s - s'; t, t')$

3) Spatio-temporal stationarity:  $K(s, s'; t, t') = f(s - s'; t - t')$

4) Separability:  $K(s, s'; t, t') = f_1(s, s')f_2(t, t')$

Cressie and Huang (1999), Gneiting (2002), many others:

stationary, non-separable CF's

### 3. Latent approach

Explain  $X$  behaviour through model components

CF's appear as a consequence

Knowledge on how  $X$  varies over  $(s, t)$ , through known functions  $\eta_i(s, t)$ ,

$$\Rightarrow X(s, t) = \sum_i \beta_i \eta_i(s, t) \quad ( +e(s, t) )$$

errors  $e(s, t)$  do not carry any spatio-temporal correlation

Usually not feasible  $\rightarrow$  knowledge rarely exists

Even when it does, there may remain substantial correlation

Idea can be used to span the space of possible representations

$$\Rightarrow X(s, t) = \sum_i \beta_i \eta_i(s, t) \quad ( +e(s, t) )$$

Now: unknown  $\eta_i$  functions form a basis for (smooth) functions of  $S \times T$

These functions adequately chosen to cover the entire space  $S \times T$

This may require too many  $\eta_i$ 's

Alternative: tensor product

$\{\eta_i(s, t)\}$  replaced by  $\{\phi_i(s)\}$  and  $\{\alpha_j(t)\}$

$$\Rightarrow X(s, t) = \sum_i \phi_i(s) \alpha_i(t) \quad ( +e(s, t) )$$

Substantial dimension reduction, but there is a cost

More importantly, relevant correlation may still remain

General approach: allow  $\{\eta_i(s, t)\}$  to be stochastic

How? to be detailed later

Simplified approach: allow  $\{\phi_i(s)\}$  and/or  $\{\alpha_j(t)\}$  to be stochastic

$$\Rightarrow X(s, t) = \sum_i \phi_i(s) \alpha_i(t) \quad ( +e(s, t) )$$

$\{\phi_i(s)\}$  are Gaussian processes (GP) and/or  $\{\alpha_j(t)\}$  are time series

Usual time series model: autoregressive

Example - AR(1):  $\alpha(t) = G(t)\alpha(t - 1) + w(t)$ ,  $w(t) \text{ iid } N(0, \Sigma)$

Lots of references: Mardia, Goodall, Redfern & Alonso (1998), Wikle & Cressie (1999), Stroud, Muller & Sansó (2001), Calder (2007), Sansó, Schmidt & Nobre (2008), ...

Also leads to non-separable CF's

## 4. Dynamic Gaussian processes

Set  $\eta(s, t) = (\eta_1(s, t), \dots, \eta_m(s, t))$

DGP:  $\eta(s, t) = G(t) \eta(s, t - 1) + w(s, t)$ ,  $w(\cdot, t) \text{ iid } mGP(0, \rho)$

Process is completed with initialization  $\eta(s, 1) \sim mGP$

Leads to non-separable but now also temporally non-stationary CF's

Simplest DGP: univariate spatio-temporal random walk

$\eta(s, t) = \eta(s, t - 1) + w(s, t)$ ,  $w(\cdot, t) \text{ iid } GP$

Useful model for smooth spatio-temporal variation of the data  $Y(s, t)$

$Y(s, t) = \eta(s, t) + WN \text{ error}$

Can handle spatio-temporal heterogeneity of other model components



## 4.1. Regression

Assume the presence of covariates  $Z(s, t)$  associated with data  $Y(s, t)$

Standard approach:  $Y(s, t) = Z(s, t)^T \eta + X_0(s, t) + WN \text{ error}$

with  $X_0 \sim DGP$

Spatio-temporal heterogeneity may also be present in the regression part

Revised approach:  $Y(s, t) = Z(s, t)^T \eta(s, t) + X_0(s, t) + WN \text{ error}$

with  $X = [\eta, X_0] \sim DGP$

The mean of the DGP may be separated from  $X(s, t)$

Now,  $X(s, t) = X(t) + X^*(s, t)$  where  $X^*$  is a zero-mean DGP

Model completed with an evolution for time-varying mean  $X(t)$

## 4.2. Trend

Assume the DGP  $X(s, t)$  is subject to spatio-temporal variations

These are modeled with an auxiliary growth process  $\gamma(s, t)$

$$X(s, t) = X(s, t - 1) + \gamma(s, t) + w_X(s, t), \quad w_X(\cdot, t) \text{ iid GP}$$

$$\gamma(s, t) = \gamma(s, t - 1) + w_\gamma(s, t), \quad w_\gamma(\cdot, t) \text{ iid GP}$$

This is in the form of a bivariate DGP with  $G(t) = \begin{pmatrix} 1 & 1 \\ 0 & 1 \end{pmatrix}$

Acceleration and higher order variations can be equally defined

Useful for estimation of trend features and specially for prediction

### 4.3. Seasonality

Assume the DGP  $X(s, t)$  is subject to seasonal variations

Seasonality given by a sinusoidal wave (temperature): single harmonic

Requires an auxiliary process  $\gamma(s, t)$

$$\begin{pmatrix} X(s, t) \\ \gamma(s, t) \end{pmatrix} = \begin{pmatrix} \cos(2\pi/q) & \sin(2\pi/q) \\ -\sin(2\pi/q) & \cos(2\pi/q) \end{pmatrix} \begin{pmatrix} X(s, t-1) \\ \gamma(s, t-1) \end{pmatrix} + w_S(s, t),$$

where  $w_S(\cdot, t) \stackrel{iid}{\sim} \text{DGP}$  and  $q$  is the seasonal cycle length.

More elaborate seasonal patterns  $\rightarrow$  additional bivariate processes

$(X_2, \gamma_2), \dots, (X_{[q/2]}, \gamma_{[q/2]})$  associated with 2nd,  $\dots$ ,  $[q/2]$ th harmonics

Simpler form:  $\sum_{j=1}^q X(s, t-j) = w(s, t)$ , where  $w(\cdot, t) \stackrel{iid}{\sim} \text{DGP}$

but requires  $q$ -dimensional DGP

## 5. Applications

DGP can be used in models for 3 types of spatial data

Directly applied to the mean structure of spatio-temporal data  $Y(s, t)$

### 5.1. Geostatistics - continuous space

Spatial version of generalized dynamic linear models

$$Y(s, t) \sim \mathcal{F} \text{ with mean } \mu(s, t) \quad [ \text{eg } N( \mu(s, t), \sigma^2 ) ]$$

$$g[\mu(s, t)] = F(s, t)^T X(s, t) \quad [F^T = (Z^T, 1) \text{ and } X = (\eta, X_0)]$$

$$X(s, t) = G(t)X(s, t - 1) + w_X(s, t), \quad w_X(\cdot, t) \text{ iid GP}$$

## Illustration: Effect of precipitation (Prec) over temperature (Temp)

$$\text{Temp}(s, t) = \beta_0(s, t) + \beta_1(s, t) \text{Prec}(s, t) + v(s, t)$$

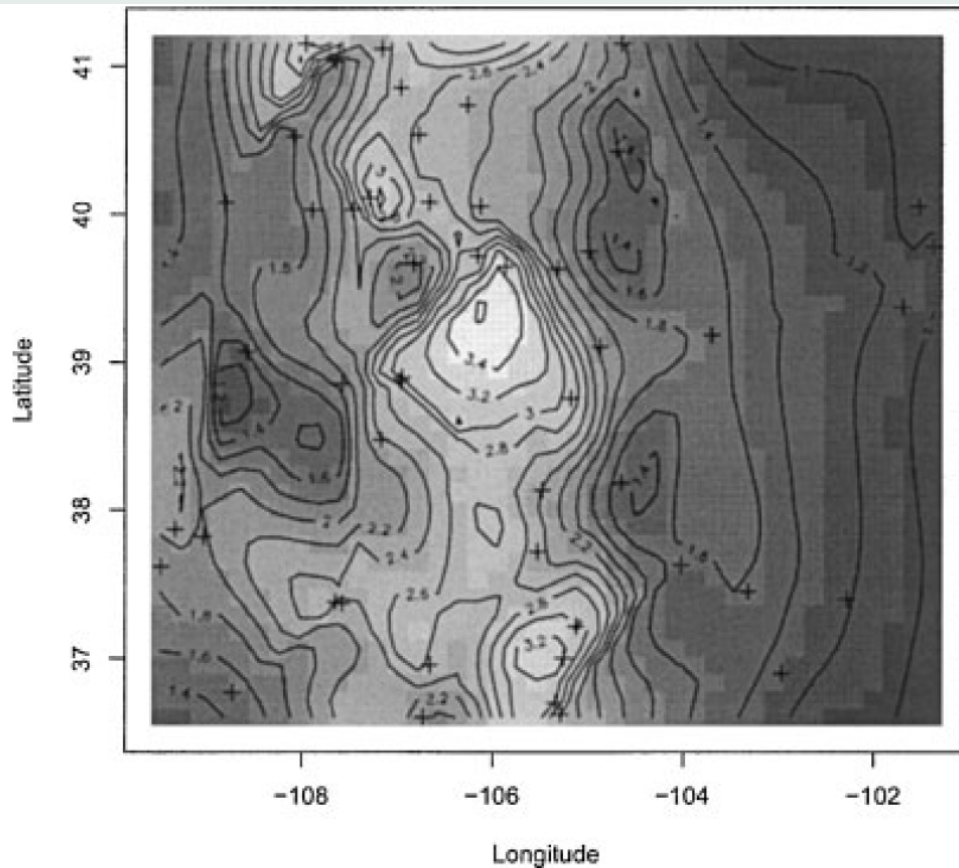
$$\beta_0(s, t) = \beta_0(s, t - 1) + w_0(s, t), \quad w_0(\cdot, t) \text{ iid GP}$$

$$\beta_1(s, t) = \beta_1(s, t - 1) + w_1(s, t), \quad w_1(\cdot, t) \text{ iid GP}$$

Intercept  $\beta_0$  and regression coef  $\beta_1$  are spatio-temporally varying

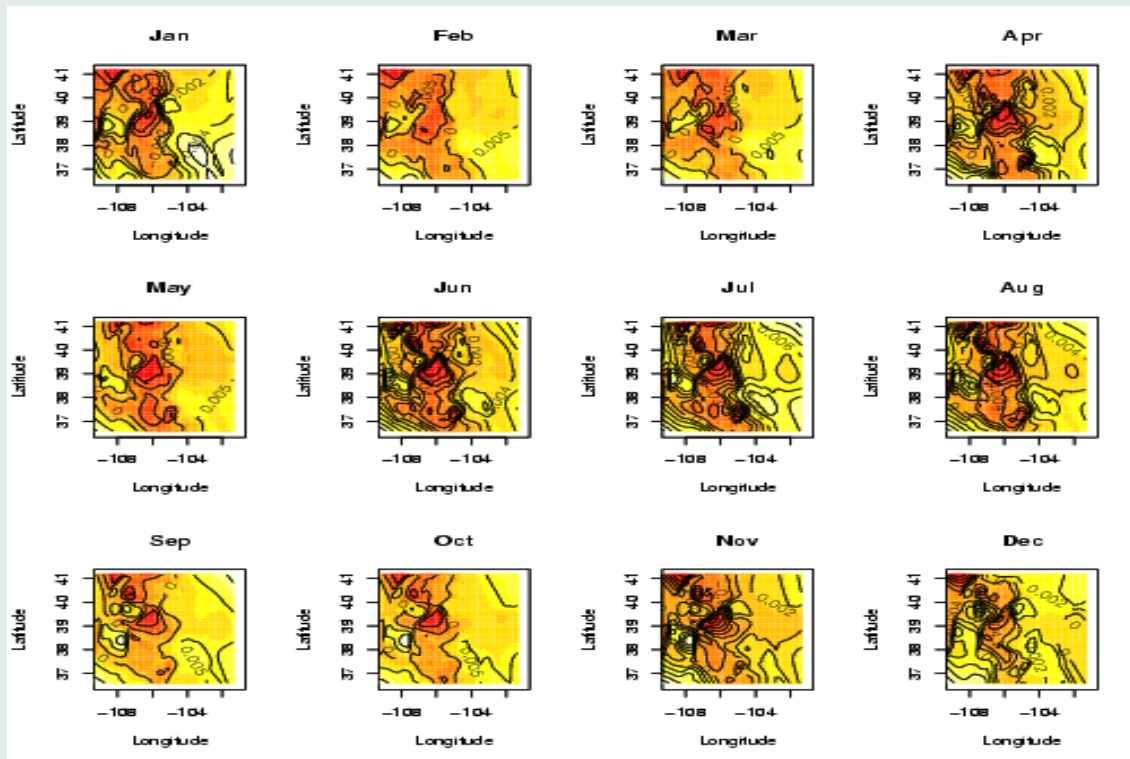
Disturbance processes  $w_0$  and  $w_1$  may be related

eg. linear transformations of independent GP's



$S$ : region in the state of Colorado, USA ( + - monitoring stations )

$T$ : { *Jan/1997*,  $\dots$ , *Dec/1997* }



Posterior mean of the spatio-temporal variation of the regression coefficient  $\beta_1$  for a region of the State of Colorado, USA (Gelfand, Banerjee and Gamerman, 2005).

## 5.2. Areal data

[based on Vivar and Ferreira (JCGS, 2012)]

Correlations are no longer based on CF's and Euclidean space

Mostly based on precision (inv. variance) matrices and neighbourhoods

Dynamic GP's replaced by dynamic Markov random fields (MRF)

Generalized spatio-temporal linear model

$$Y_{i,t} \sim \mathcal{F} \text{ with mean } \mu_{i,t} \quad [ \text{eg } N( \mu_{i,t}, \sigma^2 ) ]$$

$$g[\mu_{i,t}] = Z(s, t)^T X_{i,t}$$

$$X_t = G(t)X_{t-1} + w_t, \quad w_t \text{ iid } MRF(0, Q) \equiv N(0, Q^{-1})$$

$Q$  is sparse (full of 0's) ...  $q_{ij} \neq 0$  indicates neighborhood  $(i, j)$



Illustration:

Evolution of the homicide rate in Rio de Janeiro state municipalities

Contamination model

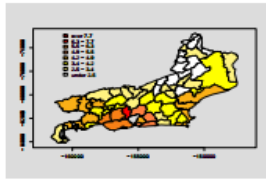
$$Y_{i,t} = \mu_{i,t} + e_{i,t}, \text{ where } e_{i,t} \sim N(0, \sigma_{i,t}^2)$$

$$\mu_t = H\mu_{t-1} + w_t, \text{ where } w_t \sim N(0, Q^{-1})$$

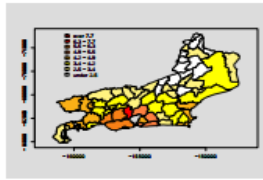
$$h_{i,j} = c \times \begin{cases} 1, & \text{if } i = j \\ \alpha, & \text{if } i \text{ and } j \text{ are neighbors} \\ 0, & \text{otherwise} \end{cases}$$

$\alpha$  is the contamination index

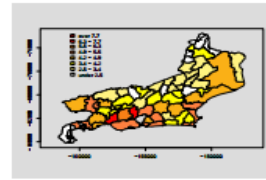
$c$  is the contamination persistence



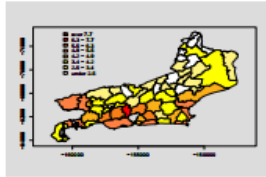
$\beta_0$



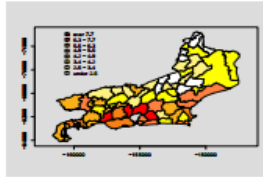
1979



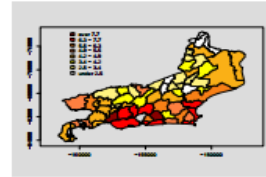
1980



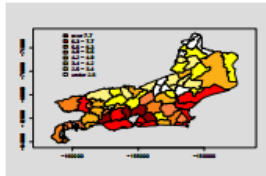
1981



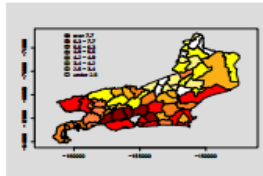
1982



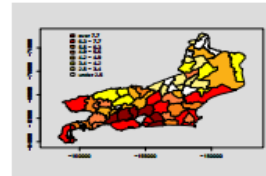
1983



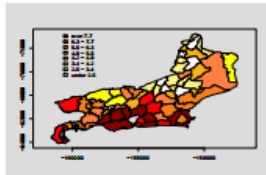
1984



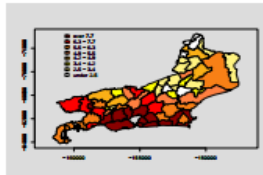
1985



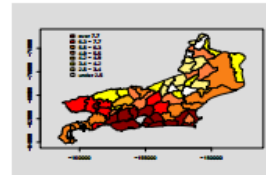
1986



1987



1988



1989

Posterior mean of the area level

### 5.3. Point pattern

based on Pinto Jr. (2014)

Data: space-time locations of occurrence of events

Usual models: PP, NHPP, Cox process (CP), log-Gaussian CP

LGCP (space only):  $Y \sim NHPP(\lambda)$  and  $\log \lambda = X \sim GP$

Liang et al (2008):  $\log \lambda(s) = Z(s)^T \eta + X(s)$ ,  $X \sim GP$

Our extension:

$Y \sim PP(\lambda)$  where  $\lambda : S \times T \rightarrow R^+$

$$g[\lambda(s, t)] = Z(s, t)^T \eta(s, t) \quad [eg. : g = \log]$$

$$\eta(s, t) = G(t)\eta(s, t - 1) + w_\eta(s, t), \quad w_\eta(\cdot, t) \text{ iid } GP$$

Likelihood:

$$l(\lambda; Y) = \prod_i \lambda(s_i, t_i) \prod_t \exp \left[ - \int_S \lambda(u, t) du \right]$$

difficult problem  $\rightarrow$  depends on an infinite-dimensional unknown function

Møller et al (1998) solution for CP: discretize  $\lambda$  over space  $\rightarrow$

problem: it is an approximation

Adams et al (2009) solution for CP: use thinning  $\rightarrow$

problems: does not extend easily for discrete time and for further model components and dimension explosion and MCMC...

We are currently working on these solutions for our DGP model

## Illustration:

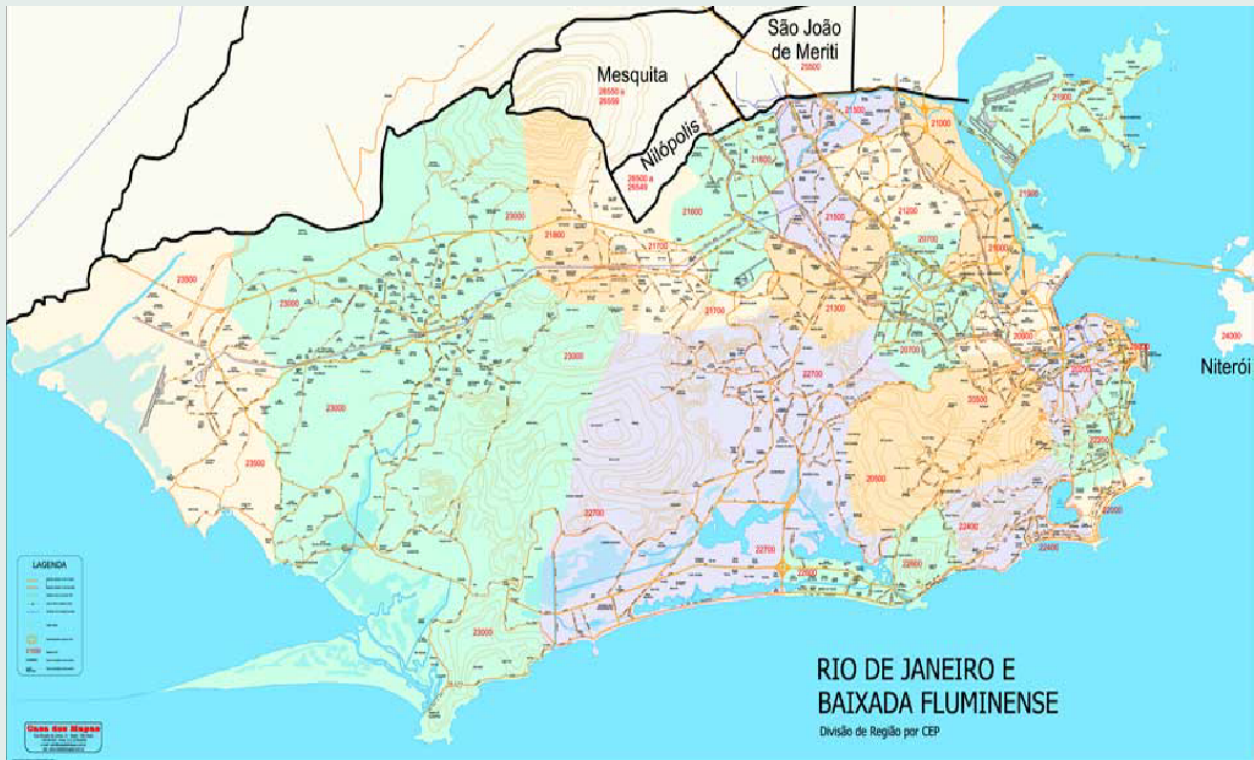
Evolution of the vehicle casualties (**theft/robbery**) in Rio de Janeiro city considering Type (private/commercial) and Age (manufacturing year)

$$Y \sim PP(r \lambda) \text{ where } r, \lambda : S \times T \rightarrow R^+$$

$$r(s, t; \text{Type}, \text{Age}) = \text{total exposure at } (s, t) \quad [\text{known offset}]$$

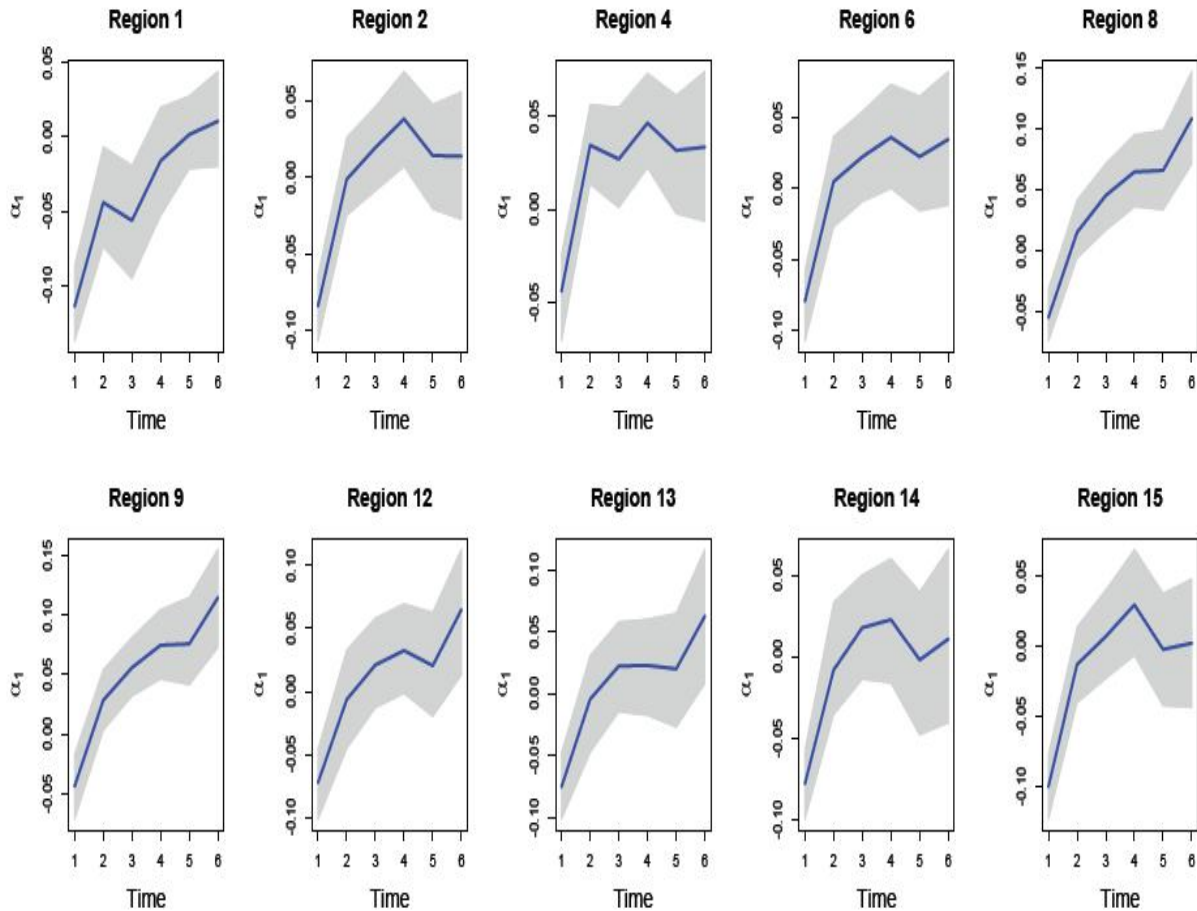
$$\lambda(s, t; \text{Type}, \text{Age}) = \exp\{X_0(s, t) + \text{Type } X_1(s, t) + \text{Age } X_2(s, t)\}$$

$$\{X_0, X_1, X_2\}(s, t) = \{X_0, X_1, X_2\}(s, t - 1) + w_X(s, t), \quad w_X(\cdot, t) \text{ iid } GP$$



$S$ : Rio de Janeiro municipality

$T$ :  $\{sem1/2009, \dots, sem2/2011\}$



Posterior mean of the spatio-temporal coefficient process of Age

## 6. Final comments

Flexible approach

Handles non-separable processes

Handles spatial non-stationary/stationary processes

Handles temporal non-stationary/stationary processes

More on the point process: [Jony Pinto Jr. \(here\)](#)

More on point process without discretization: [F. B. Gonçalves \(JSM 2014\)](#)



Thank you!

dani@im.ufrj.br